Education and Partisanship

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Abstract

Education weakens the historical link between income and partisanship across democracies, challenging classical models of political economy. Using administrative data on millions of voters exposed to discontinuities in compulsory schooling laws and college admissions in Florida and California, I show that high schools and selective colleges reduce Republican Party affiliation by 2 percentage points per year of attainment, raising independent and Democratic registration. Effects generalize across generations (1969 to present), settings, demographics, and institutions. Peer socialization and career paths emerge as key mechanisms, while instructor-driven political influence is unlikely to explain these results.

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1 Introduction

Political systems across democracies have been upended by right-wing populists that campaign against elites and institutions (Gethin et al., 2021a; Economist, 2025). A striking feature of this realignment is the displacement of income by education: many high earners with elite degrees now favor left-wing parties, while many low earners with less education now favor the right (Kitschelt and Rehm, 2019; Steel, 2024; Gennaioli and Tabellini, 2025). This inversion of party coalitions challenges classical models of political economy that predict economic self-interest and redistribution as the main axis of political conflict (Romer, 1975; Meltzer and Richard, 1981). It also carries risks that trust in research declines, ideological sorting rises, and policymakers intervene in academic affairs (Marietta and Barker, 2019; Parker, 2019; Zhang, 2023; Bender, 2025; Bender and Saul, 2025). Using several regression discontinuity designs and administrative data on millions of people treated by high schools and colleges in Florida and California, I show that education reduces Republican Party affiliation across generations, settings, demographics, and institutions.

The erosion of the class-based political cleavage predicted by economic theory has spurred a broad literature identifying three compounding reasons why education now defines party coalitions (Gethin et al., 2021b). The first is moral universalism and the rise of identity politics related to race, gender, and religion, which can determine partisanship independently of income (Kuziemko and Washington, 2018; Enke, 2020; Enke et al., 2022, 2023, 2024; Bonomi et al., 2021; Apfeld et al., 2022; Cappelen et al., 2025). The second is a shift in left-wing parties economic priorities toward green industries, immigration, and international trade, alienating blue-collar workers and trade unionists who once saw them as agents of their redistributive interests (Cavaille and Marshall, 2019; Autor et al., 2020; Kuziemko et al., 2023; Angrist et al., 2024; Choi et al., 2024). The third is education shapes the demand for parties and candidates, amplifying polarization between high-earning graduates and less-educated social outsiders (Gethin et al., 2021b; Dal Bo et al., 2023).

The political externalities of education remain an open economic question with major

theoretical and empirical stakes. Americans' expectations are evenly split: moderates believe education has null effects, while the ideological poles expect it shifts students to the left (Orth, 2022). Classical models of political economy predict education moves voters to the right by reducing demand for redistribution, but recent theories stress that identity and issue salience may complicate this (Meltzer and Richard, 1981; Bonomi et al., 2021; Gennaioli and Tabellini, 2025). To date, empirical work in economics has not estimated the partisan externalities of education, in part because educational institutions are cautious about sharing administrative data for politically sensitive questions. Research from other disciplines uses matching methods and surveys to address the challenge, reaching no consensus (Marshall, 2016; Campbell and Horowitz, 2016; Marshall, 2019; Scott, 2022; Simon, 2022).[1] While these approaches offer valuable insights, they struggle with unobserved selection into schooling and generalizability, leaving the causal question unresolved.

I estimate the impact of education on partisanship with administrative data on millions of people in California and Florida who were exposed to discontinuities in college admission rules and compulsory schooling laws. First, I link the universe of 220 million voters from all 50 American states to a sample of roughly 250,000 University of California applicants who were evaluated under the Eligibility in the Local Context (ELC) policy, which granted admission to students in the top four percent of their high school cohort. This enables estimates of the effect of attending selective colleges on party affiliation in a system that accounts for more than one eighth of elite Association of American University enrollment in the United States (Atkinson and Pelfrey, 2004; Bleemer, 2024). Second, I use compulsory schooling laws (CSLs) in Florida and California that bind on exact birthdate to identify the impact of high school on partisanship for 5 million voters near the discontinuity (Dobkin and Ferreira, 2010; McCrary and Royer, 2011) This approach generalizes treatment effects from elite colleges to high school dropouts across two states with 37 million voters and a 34-point partisan gap, ranging from a Republican +14 to a Democratic +20 margin.

[1] These papers use fixed effects, sibling comparisons, or compare birth cohorts and generally find that education shifts students to the right or has no impact.

I find that each year of schooling reduces Republican Party affiliation by roughly 2 percentage points, with impacted students becoming independents or Democrats. Despite targeting opposite ends of the schooling distribution, marginal high school dropouts and marginal elite college entrants, both research designs yield the same effects. For reference, the partisan shift from a single year of schooling is double the margin of victory in the tipping-point state in the three most recent U.S. elections. Effects emerge shortly after schooling, are consistent for birth cohorts dating back to 1969, and persist into voters' early 50s. They are larger in Republican-leaning Florida than in Democratic-leaning California and are more pronounced among men and Whites than among women and people of color.

Recent work highlights three mechanisms that may explain education's impact on partisanship. Roommates and classmates socialize students into their political views and identity.[2] Faculty, curricula, and instruction can shape these outcomes as well (Cantoni et al., 2017; Goldstein and Kolerman, 2025). The long-run effects of education on earnings, career, and residence may also matter over time (Finan et al., 2021; Cantoni and Pons, 2022).

I find evidence consistent with peer socialization as the primary mechanism, with a secondary role for later-life career paths. Students induced into more selective colleges live in peer-dense residential settings, spend less time with family, and engage with classmates from more liberal, secular, and nationally diverse backgrounds. In-sample surveys confirm that students at elite colleges perceive peers, rather than faculty or parents, as the dominant influence on their political views, and that campus social environments vary sharply by college selectivity. Parallel patterns emerge in the compulsory-schooling setting: effects are twice as large among students born after 2002, for whom CSLs increased in-person peer exposure just before COVID-era disruptions. By contrast, I find little support for instructor or curriculum-based channels. Elite college faculty are research-focused, students struggle to discern their views, and neither household partisanship nor neighborhood composition is

[2]Exogeneous variation in peer exposure changes attitudes toward diversity (Carrell et al., 2019; Alan et al., 2021; Corno et al., 2022), support for redistribution (Mendelberg et al., 2017; Londono-Velez, 2022), as well as ideology (Billings et al., 2021; Braghieri, 2021; Strother et al., 2021; Kaplan et al., 2025).

impacted by treatment. While treated students earn more and complete more schooling, career paths reinforce earlier effects by reducing party switches at the threshold.

This paper makes contributions to empirical and theoretical work in labor economics, the economics of inequality, and political economy. Empirically, it is the first to use administrative data and multiple natural experiments to show that education consistently reduces Republican Party affiliation, highlighting partisan externalities embedded in the returns to education that are largely overlooked. Theoretically, these findings challenge classical models of political economy, which predict that higher earnings from education reduces demand for redistribution and shifts voters to the right. Instead, I find that education moves students toward the political center and left while raising earnings, eroding income-based partisanship and supporting recent models of political economy that stress issue salience and identity politics (Bonomi et al., 2021; Gennaioli and Tabellini, 2025). These findings contribute to our understanding of how educational institutions produce not just human capital, but partisan externalities in the evolving landscape of economic inequality and democratic conflict.

2 Data, Methods, and Policy Context

2.1 Data

This paper uses two primary data sources: the nationwide universe of roughly 220 million registered U.S. voters from 2012 to 2024 and administrative records on over 250,000 University of California (UC) applicants between 2007 and 2011. The voter file is used independently to estimate the political externalities of compulsory schooling. These sources are then linked at the individual level to estimate the impact of elite colleges on long-run partisanship.

The voter registration data are drawn from L2 Inc.'s VM2 database, a commercially maintained, nonpartisan voter file used widely by researchers and political campaigns in the United States. The file includes each registrant's political party affiliation as well as their vote

history in all primary and general elections since 2012. These data are supplemented with commercial, demographic, and geographic information from all 50 states and Washington, D.C.

The UC applicant data were provided by an anonymous campus, referred to here as "UC San Andreas". The dataset includes all first-time applicants to that campus between 2007 and 2011. Because of the UC common application, the majority of California residents apply to UC San Andreas, making this a close proxy for the broader UC applicant pool. The dataset includes detailed applicant information not available in centralized UC records. Individual records were de-identified and matched to the voter file on name and birthdate. Postsecondary enrollment is verified using the National Student Clearinghouse and augmented with institutional characteristics from IPEDS, Opportunity Insights, and the College Scorecard.

To examine mechanisms, I merge survey data on college students and faculty from the Cooperative Institutional Research Program (CIRP) at UCLA's Higher Education Research Institute (HERI). Student-level ideology, religiosity, and issue preferences come from more than 4 million full-time freshmen surveyed at over 1,000 colleges between 2000 and 2010. Faculty characteristics are drawn from HERI's triennial surveys (1989 to 1998)[3], which cover career goals, views, and teaching practices across more than 1,000 institutions and 80,000 respondents.

Finally, I fielded a proprietary online survey in the Spring 2022 of 1,105 in-sample UC San Andreas applicants, collecting self-reported political views, civic engagement, and retrospective experiences. The survey instrument replicates wording from Pew Research surveys to facilitate comparison to national benchmarks. Appendix Table A.1 shows the survey sample is demographically similar to the full population of UC applicants. Full question text appears in Online Appendix A.

[3]Later years are not publicly available to protect the identity of faculty members responding to the survey.

2.2 University of California Admissions

California's higher education system includes one quarter of all Americans enrolled in open-access community colleges, half a million students in teaching-focused California State Universities (CSUs), and one eighth of all Americans enrolled at elite AAU research universities in the University of California (UC) campuses. While this paper links applicants to voter registration outcomes in all 50 states and the District of Columbia, more than 80 percent of in-sample registrants reside in California, which has the highest rate of college student retention in the U.S (Van Dam, 2022). Importantly, party membership is a strong predictor of actual party and ideological preferences (See Online Appendix A).

The University of California system enrolls over a quarter million students across nine undergraduate campuses and is one of the largest selective public higher education systems in the world. Mirroring university systems in other countries and American states, elite colleges in California have more left-leaning students than their less selective counterparts (Kerr et al., 2001; Gethin et al., 2021a,b; Firoozi and Geyn, 2024). Among in-sample students attending UC campuses, 60 percent are registered Democrats and 8 percent are Republicans, which is similar to peer institutions outside California such as the University of Michigan (56 percent to 10 percent), the University of Texas (56 percent to 17 percent), and NYU (70 percent to 5 percent). These patterns remain consistent in national survey data regardless of institutional control, sector, or region, supporting the generalizability of results beyond California (see Table B.1).

To estimate the causal effect of selective colleges on partisanship, I exploit a natural experiment generated by the UC's Eligibility in the Local Context (ELC) policy. Introduced in 2001 after the implementation of Proposition 209, which banned race-based affirmative action, the ELC program granted an admissions preference to California high school students who ranked in the top 4 percent of their cohort by a version of GPA that included additional weight for college-preparatory courses taken in the sophomore and junior years. Eligibility thresholds were calculated internally by the UC Office of the President based on transcript

data submitted by participating high schools, which accounted for over 90 percent of in-state UC applicants. The "reweighted GPA" rankings and cutoff values were never disclosed publicly, but students were notified of their eligibility for the policy via formal letter. Admissions offices at individual UC campuses were also informed of each applicant's eligibility and given discretion to incorporate this information into their admissions decisions.

ELC lends itself to a credible regression discontinuity design (RDD). The running variable, reweighted GPA, is continuous and centrally calculated. Students lacked access to their reweighted GPA percentile or their respective cutoff, precluding manipulation. Eligibility affected college admissions and enrollment, but was not disclosed publicly or used for other policies like financial aid. These features satisfy the key requirements for an RD design: a relevant instrument, continuity of expected outcomes around the cutoff, and imperfect control of the running variable.

The primary threat to the identification strategy's validity comes from the risk that a subset of applicants who were made aware of their eligibility for the top percentile policy selected into UC San Andreas application differentially across the GPA threshold. Notably, there were students who were contacted about their eligibility for the ELC program that selected into application for some UCs, but this does not meaningfully bias estimates of the labor market returns to UC campuses (Bleemer, 2023, 2024).

I use several approaches to validate the design and address the risks of selection into sample. McCrary density tests show no discontinuous jump in the distribution of the running variable at the cutoff (Figure D.1), and I fail to reject the null hypothesis of smooth density (McCrary, 2008; Cattaneo et al., 2018, 2019). Covariate balance tests in Tables D.1 through D.5 and Figures D.2 through D.6 show no discontinuities in 16 predicted outcomes or 18 baseline demographic and academic variables using local linear estimation with a 0.3 GPA bandwidth at a 90 percent confidence interval. Across multiple bandwidths and specifications, the rejection rate for covariate continuity is consistent with random chance (see Figures D.7 through D.11). I also re-estimate results conditional on voter registration to im-

prove balance and find substantively identical treatment effects to my main results (Tables D.3 and F.1).

To strengthen internal and external validity, I replicate my findings using two additional GPA-based admissions thresholds at highly selective UC campuses. These campus-specific policies introduced undisclosed thresholds for admission priority based on reweighted GPAs. The thresholds are located at different points in the GPA distribution than the ELC policy (around the 39th and 89th percentiles among applicants)[4], were unknown to applicants, and created sharp changes in admission probabilities. McCrary and covariate balance tests confirm the credibility of these alternative RDDs (see Figures E.1 through E.8 and Tables E.1 and E.2). These two alternative admission rules vary different margins of treatment – the extensive margin of college enrollment and the intensive margin of college selectivity – across different types of institutions and student populations to confirm the robustness of the main findings.

My RD specifications take the following form:

$$Outcome_i = \alpha + \beta \cdot Eligible_i + f(GPA_i) + \mathbf{X}_i'\Omega + \varepsilon_i \tag{1}$$

where $Outcome_i$ is an outcome for student i, GPA_i is the normalized reweighted GPA (with the cutoff set to zero), $Eligible_i$ is a binary indicator for ELC eligibility, $f(\cdot)$ is a flexible function of the running variable, $\mathbf{X}_i$ is a vector of covariates, and standard errors are clustered at the high school-cohort level. I vary the order of a polynomial control for the running variable, include an expansive set of controls, change the bandwidth used for inference, and estimate bias-aware confidence intervals to demonstrate the robustness of my estimates (Calonico et al., 2014; Kolesar and Rothe, 2018).[5]

[4]The ELC cutoff was generaly around the 79th percentile of applicant GPA for perspective.

[5]The controls I use include parental years of schooling, self-reported income, and ISIR family income, as well as indicators for female, underrepresented minority status, Cal Grant eligibility, first generation college student status, FAFSA filing, application year, county education level, high school quality, being a primary income earner as a student, and missing covariate information. I use the bounded second derivative method from Kolesar and Rothe (2018), deriving bounds based on a heuristic rule offered by the authors that makes assumptions on the maximum plausible difference between the CEF and a straight line between the CEF

2.3 Florida and California's Compulsory Schooling Laws

To generalize findings to voters with the lowest educational attainment, I exploit compulsory schooling laws (CSLs) in Florida and California that create sharp variation in years of schooling based on birthdate.[6] In both states, minimum school entry ages are enforced by date-of-birth cutoffs: children must turn five before September 1st in Florida or December 2nd in California to enroll in kindergarten that academic year.[7] Those born after the cutoff must delay entry by one year, which means that they begin schooling one year older than their counterparts born a day earlier. These entry laws, coupled with compulsory attendance requirements that prevent high school dropout until a student reaches a given age (usually 16 to 18), generate clear jumps in educational attainment on the order of 0.15 to 0.20 years of schooling (Dobkin and Ferreira, 2010; McCrary and Royer, 2011). I define a normalized running variable in days, with zero set at the cutoff, and estimate intent-to-treat (ITT) effects on long-run outcomes.

CSLs in these contexts satisfy the key assumptions for causal inference with a regression discontinuity design. Birthdate is as-good-as-random within narrow windows around the threshold and is not manipulable by parents or their children. Although induced births and cesarean deliveries respond to seasonality and holidays, there is no evidence that parents time births around the specific CSL cutoff dates in these states (LaLumia et al., 2015; Jacobson et al., 2020). Moreover, because birthdate is assigned at birth, it is determined prior to any decisions about education or politics and is not subject to measurement error in administrative records. These features support the exclusion restriction that the cutoff affects outcomes only through its impact on schooling.

Because I observe people conditional on being registered to vote, discontinuities in the

values at the endpoints of an interval of a fixed length in the support of the running variable.

[6] I choose these states because they are the fist and third most populous American states. Other large states like Texas either do not record voters' party affiliation in their state records, do not report voters' exact date of birth, or do not have a sharp birthdate cutoff paired with a school leaving age embedded in their compulsory schooling laws.

[7] Florida's September 1st cutoff binds for birth cohorts dating back to 1978. California's December 2nd cutoff is used for cohorts dating back to 1969, consistent with previous work.

number or demographics of people around the cutoff would reflect treatment effects of education on voter registration, rather than differences in birth patterns across the threshold that would threaten the identification strategy. Perhaps surprisingly given the large literature on education's civic externalities, there is no discernable discontinuity in the density of observations around the threshold, suggesting no impact of high school education on voter turnout (Figure D.12). Likewise, covariate balance tests show smooth trends in composition of the sample by race and sex at the threshold (Figure D.13).[8] These tests are consistent with previous work using birthdate discontinuities in CSLs and validate the design in this context.

The near-threshold estimation sample includes over 5 million registered voters born within 70 days (10 weeks) of the cutoff dates in Florida and California, drawn from a total of 37 million registrants in the two states. My baseline specification takes the following form:

$$Outcome_i = \alpha + \beta \cdot BornBefore_i + f(Birthdate_i) + \mathbf{X}_i'\Omega + \varepsilon_i \tag{2}$$

where $Outcome_i$ is an outcome for individual i, $BornBefore_i$ is an indicator for being born on or before the entry cutoff date, $f(\cdot)$ is a flexible control for the running variable, $\mathbf{X}_i$ includes controls for sex, race, birth year, and state of registration, and ε_i is the error term. For robustness, I vary the order of the polynomial control for the running variable, the bandwdith used for inference, and the inclusion of pre-treatment covariate controls.

[8]For each of these test I use local linear estimates with a 70 day bandwidth, uniform kernel, and 90 percent confidence interval. Evidence of balance remains the same at narrower bandwidths.

3 Results

3.1 University of California Admissions

3.1.1 First-Stage

Eligibility for the University of California's top percentile admission policy significantly increased students' access to and enrollment in highly selective UC campuses. Consistent with prior work by Bleemer (2024), I find a clear discontinuity in admissions at the eligibility threshold. Students just above the 96th percentile of GPA were admitted to approximately 0.4 more UC campuses than similar students just below it, on a mean of roughly three. This effect is robust across bandwidths, specifications, and the inclusion of covariates (see Online Appendix Tables B.2 and B.3).

I estimate that eligibility for ELC raised enrollment at highly selective four-year institutions by about 6 percentage points, with roughly half of the substitution away from CSUs, a quarter from less selective UCs, and a quarter from two-year colleges or non-enrollment. These shifts in enrollment exposed students to campuses with higher instructional expenditures, lower acceptance rates, and better graduate earnings, which are relevant for labor market outcomes and potentially for political behavior.

The reduced-form results reported throughout the paper capture the joint effect of increased admission and enrollment at more selective institutions. I use instrumental variables specifications primarily to aid interpretation of magnitudes. Full enrollment effects, covariate balance tests, and alternative bandwidth results are presented in Online Appendix B.

3.1.2 Voter Registration and Partisanship

Figure 1 presents the core reduced-form effects of UC admission eligibility on eight voter registration outcomes, plotted against students' normalized GPAs. The top six panels show the total fraction of students who register to vote, as well as unconditional registration shares

by party: Republican, non-Republicans[9], Democrat, no party preference, independent, and third party. The final two panels track changes in major party registration over time. Visual inspection reveals clear discontinuities. Eligibility for the top percentile admission policy lowers Republican registration and switching into the Republican Party, while increasing the probability of registering as a third party, a Democrat, or no party preference.

Table 1 formalizes these findings across three panels. Panel A shows a modest increase in overall registration, consistent with past research on civic returns to higher education (Firoozi and Geyn, 2024). Panel B presents the main effects on partisanship. In my preferred specification (Column 3), each 1,000 students eligible for the UC policy yields 6 fewer Republican registrants and 19 additional registrants as independents or Democrats. These results reflect substantial realignment. Given the distribution of party preferences within non-Republican registrants, I estimate that UC admission raises Democratic support by 0.4 to 0.6 percentage points (see Table A.2). The corresponding IV estimates in Table C.1 imply that each UC admission reduces Republican registration by 1.62 percentage points and increases independent or Democratic registration by 4.98 points.

The magnitude of these effects is nontrivial in aggregate. Based on the 823,961 UC admissions during the 2007 to 2011 sample window, I estimate that the policy reduced the number of registered Republicans in California by roughly 13,348 and increased Democratic and independent registration by 41,033, both roughly 0.3% of the state's totals.

Panel C of Table 1 examines partisan conversion. The L2 voter file flags individuals who switch their party registration. While I find no significant shift in conversions to the Democratic Party, eligibility for UC admission reduces conversion to the Republican Party by two per 1,000 applicants. This suggests that the effect of education on partisanship may grow well into adulthood.

I test the robustness of these results in four ways. First, I re-estimate all outcomes con-

[9]I use this as an outcome to remain consistent with national youth surveys and in-sample polling, which show that young college-educated independents overwhelmingly support Democrats and left-leaning policy positions (see Figure A.2, Tables A.2, A.3).

ditional on voter registration, motivated by the strong balance documented in Table D.3. Table F.1 shows stronger and directionally identical effects, as UC admission reduces Republican affiliation and raises Democratic and independent affiliation. Second, I demonstrate robustness to bias-aware confidence intervals (Calonico et al., 2014; Kolesar and Rothe, 2018) and high-dimensional high school-year fixed effects in Tables F.2 through F.4. Third, I vary bandwidths, covariate sets, and functional forms in Figures F.1 through F.8, showing that results are stable across a wide range of specifications.

Finally, I implement falsification tests based on placebo cutoffs. I simulate "synthetic" discontinuities across the GPA distribution, excluding a 0.05 GPA window around the true ELC threshold, and compare the resulting t-statistics to those from the actual policy cutoff (Appendix Figures G.1 through G.3).[10] The real effects on Republican registration, Democratic/independent registration, and partisan conversion all fall above the 95th percentile of this synthetic distribution, strongly suggesting the observed discontinuities represent real treatment effects.

3.1.3 Voter Turnout

In Appendix I, I evaluate the impact of UC campuses on voter turnout and political participation using eight different measures. Admission to UC campuses raises voter turnout in Democratic presidential primaries, with null effects on other types of election. Robustness checks and falsification tests confirm these results across various RD specification choices.

3.1.4 Alternative Admission Policies

To strengthen generalizability, I replicate these findings using two undisclosed GPA-based admission thresholds at highly selective UC campuses. These alternative policies, which vary both the extensive and intensive margins of enrollment and affect distinct applicant pools,

[10]Feasible points refers to each point between -1.24 and +0.27 relative to the true cutoff on the normalized reweighted GPA index, which allows the 0.3 GPA bandwidth to span the range of roughly the 1st to 99th percentiles of this normalized index. I use a 0.3 GPA bandwidth consistent with my preferred specification for a more direct comparison.

also reduce Republican registration and increase independent or Democratic affiliation. Importantly, these results hold across campuses, specifications, and student profiles, suggesting that the partisan effects of UC enrollment are not driven by a single institution or treatment margin. Full first-stage estimates, effects, and robustness checks are reported in Online Appendix Section E.

3.2 Compulsory Schooling Laws in Florida and California

3.2.1 Partisanship

To generalize the political effects of education beyond selective colleges, I exploit compulsory schooling laws (CSLs) in Florida and California that mandate minimum years of education based on exact birthdate. These laws increase educational attainment at the bottom of the distribution by enabling children born on or before the state-specific cutoff, September 1st in Florida and December 2nd in California, to enter school a year earlier, which in turn requires completion of an additional year of schooling before reaching the legal dropout age.

Figure 2 visualizes these effects. The top two panels pool both states using 70-day and 21-day bandwidths, while the bottom panels disaggregate by state at the narrower bandwidth. In every specification, students born just before the cutoff, who remain in school longer, are less likely to register as Republicans. The discontinuities are clear and consistent across settings and bandwidths, providing compelling evidence of high schools' causal effects.

Table 2 reports reduced-form estimates of party registration across six columns. Columns 1 and 2 pool both states, 3 and 4 show results for Florida, and 5 and 6 do the same for California. Odd-numbered columns show results without covariate controls whereas even-numbered columns add controls for birth year, state, sex, and race. Across all specifications and samples, Republican registration falls among those who complete more secondary schooling. Independent registration is unaffected and, in both Florida and the pooled sample, Democratic registration rises.

Comparisons of magnitudes are informative in this context. In California, where birthplace is recorded on the voter file, I restrict to in-state births to ensure CSL exposure. I do this because, for these policies to bind, a person must have been located in the state at both the relevant school entry date and their legal dropout age. In Florida, birthplace data are unavailable, but census data on interstate migration suggests roughly half of registrants were born in-state. Thus, a simple point of comparison between states would entail doubling Florida's treatment effects when benchmarking with California. Even absent that adjustment, Florida shows larger treatment effects of high school on partisanship than California, suggesting that schooling has stronger effects in more conservative political environments despite Republicans' power to shape curriculum.

To benchmark magnitudes, I note that prior work finds CSL cutoffs in these settings increase education by 0.15 to 0.2 years on average (Dobkin and Ferreira, 2010; McCrary and Royer, 2011). Given that I observe roughly 0.3 to 0.4 percentage point declines in Republican registration at the cutoff, back-of-the-envelope calculations imply that one full year of additional compulsory schooling reduces Republican registration by approximately 2 percentage points. These effects are nearly identical to those found for UC admissions,[11] suggesting that education consistently reduces Republican Party affiliation across a wide range of settings, institutions, and student populations.

I confirm the validity of the design in several ways. First, I demonstrate robustness to more flexible controls for the running variable in Table F.5. Second, I show that point estimates are similar at narrower bandwidths around the cutoff in Figure F.9. Third, in Table F.6, I re-estimate the main specifications from Table 2 on a placebo sample of foreign-born registered voters in California who were unlikely to have been subject to these schooling laws. As expected, I find no discontinuities, supporting the interpretation that the observed effects reflect true exposure to compulsory schooling laws rather than an abrupt change in

[11] A roughly 6 percentage point rise enrollment at highly selective 4-year campuses was paired with a 0.6 percentage point decline in Republican Party affiliation. This implies a roughly 2.5 percentage point decline in Republican affiliiation per year of college instruction.

birth patterns around the cutoff.

Taken together, these results affirm the external validity of the findings from UC campuses and show that education's political externalities are not confined to a particular institution, setting, or level of schooling. Both selective universities and minimum schooling laws consistently reduce Republican affiliation and increase independent or Democratic registration.

3.2.2 Heterogeneity by Generation, Sex, and Race

Heterogeneity across generations, sex, and race can offer insight into the mechanisms and scope of education's political effects. I begin by dividing the sample into three generational cohorts in Table 3: Generation X (born 1969 to 1980), Generation Y or Millennials (1981 to 1995), and Generation Z (1996 and later). Across all three cohorts, I find statistically significant reductions in Republican registration. There is no consistent effect on independent registration, but Democratic registration rises among Millennials and Generation Z. Point estimates are largest for Generation Z, a result which remains robust at narrower bandwidths. The consistency of directional effects across cohorts is notable. Even among Gen X voters, the partisan effects of additional compulsory schooling nudge students toward the left. To the extent that treatment effects are changing over time, they appear to be growing. Among Generation Z voters born after 2002, the estimated reductions in Republican registration are twice as large as for older Gen Z cohorts.

I next explore heterogeneity by sex in Table 4. Schooling induced by CSLs reduces Republican registration for both men and women, with corresponding increases in Democratic affiliation. Point estimates are larger for men, potentially reflecting their higher baseline Republican affiliation or their greater likelihood of dropping out of high school in the absence of binding school-leaving laws.

Finally, I test effects by race and ethnicity in Table 5. The largest partisan effects are among White voters, who are significantly less likely to register Republican and more likely to register as Democrats when exposed to additional schooling. Among voters of color, effects

on Republican registration are about one-half the size of those for White registrants and are not significant for Democratic registration. These patterns may reflect either the higher rates at which White voters typically register to vote as Republicans or the high share of recent immigrants among non-White Floridians, many of whom would not be subject to the state's CSLs and thus attenuate ITT estimates. Importantly, larger treatment effects among White voters also has implications for generalizability. Given that Florida and California are much more racially and ethnically diverse than the national average, the effects of education on partisanship may be even larger in Whiter states across the Midwest and Northeast.

Overall, these results reveal a consistent pattern of education reducing Republican registration and, to a lesser extent, increasing Democratic registration across generations, sexes, and racial and ethnic groups. Where heterogeneity emerges, it is directionally consistent with partisan convergence toward the left and concentrated among populations with greater baseline Republican support like men and White voters. Larger effects among younger voters suggest that the partisan externalities of education are not only durable but may be growing stronger over time.

4 Discussion

I consider three plausible mechanisms for education's partisan externalities: peer socialization, exposure to instructors or curricula, and longer-run changes in earnings, occupation, or residence. Results from both selective universities and compulsory schooling suggest that peer and long-run mechanisms are more consistent with the evidence than direct persuasion by instructors or curricula.

4.1 Peer Socialization

Peer socialization appears to be a central mechanism linking education to partisanship across settings and institutions. Importantly, both the composition of a students' peer

group and the degree of in-person exposure to peers as opposed to family are impacted by educational attainment.

The UC's ELC policy not only shifts the composition of students' peers, but also increases the intensity and duration of their exposure to those peers. Students who attend more selective UC campuses are more likely to live in student housing, spend less time with family or guardians, and engage more frequently with ideologically liberal classmates from diverse racial, religious, and socioeconomic backgrounds. These social environments differ sharply from the more moderate home settings many students would otherwise remain in as commuters, and they appear to produce lasting political effects. In-sample UC students who attend selective universities themselves cite peers, rather than professors, teachers, or parents, as the dominant influence on their political views. They also state that they discuss current events with college classmates and friends more frequently than their family and have more liberal friends compared to peers at less selective campuses (Tables A.4 through A.10). Each of these traits are associated with holding left-wing views on economic issues and progressive attitudes on social issues.

The UC's ELC policy changes the types of peers students see in the classroom alongside the amount of time they spend with them. Figure H.1 shows that UC eligibility increases exposure to peers who are affluent, Asian American, or international students, and are less likely to be White or Hispanic.[12] Table H.1 confirms these patterns, including a 3,000 dollar increase in median peer household income and a 1.2 percentage point increase in exposure to students from the richest 5 percent of families. Figures H.2 and H.3 as well as Tables H.2 and H.3 show that students at highly selective UCs are more likely to interact with liberal, secular, and left-leaning peers, and are less likely to interact with Republicans, moderates,

[12]These patterns mirror the differences in survey data between incoming UC students and their counterparts at counterfactual colleges and universities in Table J.7. I note that the racial composition numbers are likely lower bounds on the true point estimate, because the racial composition data from Opportunity insights lag behind the time period I study.

or Christians.[13][14]

Findings from compulsory schooling laws reinforce the peer socialization mechanism. The largest partisan effects are observed among students born after 2002, cohorts that had little time for labor market returns to education to materialize. This suggests that the rapid, observable partisan externalities stem from mechanisms operating during schooling itself. For these cohorts, being born before the cutoff also increased the amount of in-person peer exposure prior to the COVID-19 pandemic and the shift to remote learning, which may help explain the larger treatment effects during this period. Evidence from CSLs therefore underscores that peer environments, not just peer composition, are central to education's political externalities.

4.2 Instructors and Curricula

Instructors and curricula are frequently proposed as mechanisms underlying the political effects of higher education. Using rich data from HERI surveys and a proprietary poll of in-sample students, I find limited support for this hypothesis. While faculty at UC campuses are more left-leaning than at less selective institutions, they are less focused on shaping students values or political beliefs. In-sample students report peers and family to be far more politically influential than instructors. These patterns, coupled with my findings from

[13] In Table J.1, a large-scale survey demonstrates that entering UC students are more likely to self-identify as liberal or far-left than entering students at private colleges, CSU campuses, and community colleges. The ideological and religious gaps between UC students and their counterparts at less selective Californian colleges mirror the nationwide gap between students of more selective research universities and less selective teaching colleges (See Tables J.8 through J.9). This higher rate of left-liberal self-identification maps to both left-wing economic policy and progressive sociocultural values (See Tables J.2 through J.6).

[14] To test differences in peer ideology and religious views across the threshold, I impute these characteristics at the campus level using a mix of voter registration records and CIRP surveys from HERI. Using data available in the CIRP survey, I match summary data on entering freshmen to colleges based on their membership in one of the following groups: UCs, private Californian research universities, CSUs, private Californian teaching colleges, two year Californian colleges or no college enrollment, public out-of-state research universities, private out-of-state research universities, public out-of-state teaching colleges, private out-of-state teaching colleges, and two year out-of-state teaching colleges. Note that the method of imputation I use will likely understate the ideological gap because (1) these surveys exclude sophomores, juniors, and seniors, (2) this method treats college non-enrollees as two year college students and (3) this method homogenizes peer characteristics across broad categories of colleges and, therefore, fails to capture intra-system changes in enrollment.

compulsory schooling laws, suggest that persuasion by instructors and curricula are unlikely to be the primary drivers of education's partisan externalities.

I begin with HERI survey data to estimate faculty ideology using the same imputation strategy from Section 4.1. Table K.1 shows that UC faculty are more likely to self-identify as liberal or far-left than their counterparts at CSU campuses and private teaching-focused colleges.[15] Table K.3 shows that this ideological gap persists across both STEM and non-STEM fields. Figure H.4 and Table H.4 confirm a discontinuous increase in students exposure to liberal faculty at the ELC policy threshold.

Despite these ideological differences, UC faculty express significantly less interest in using their roles to shape student values or political behavior. As Table K.2 shows, UC faculty are less supportive of prohibiting speech they deem racist or sexist. Table K.4 shows that UC faculty rank "obtaining recognition" and "becoming an authority" in their field as primary career goals, and assign less importance to changing social values, promoting racial understanding, or influencing government and public policy. Similarly, Table K.5 indicates that UC instructors place less emphasis on developing students' morals, citizenship, or appreciation of other races and ethnicities.

Institutional priorities reinforce this pattern. UC faculty report greater commitment to promoting intellectual development and respecting opinion, but lower emphasis on multiculturalism, social change, or having students examine their values (Table K.6). These priorities are reflected in instructional practices and lower direct interpersonal engagement with students. UC faculty make greater use of teaching assistants, traditional lectures, and non-interactive methods (Tables K.7 and K.8). They assign fewer readings on race and gender, hold tenure-track positions at higher rates, and are more likely to report that research is prioritized over teaching (Tables K.9 and K.10). Each of the differences mirror those between selective universities and less selective colleges nationwide.

Students' self-reports are consistent with these institutional differences. In my proprietary

[15] This mirrors national patterns shown in Table K.11.

survey, in-sample UC students rank friends and family as significantly more influential than professors or teachers (Table A.4). They report discussing current events more frequently with peers than with family or in college (Tables A.5 through A.7) and are no more likely than CSU students to cite faculty as an important political influence (Table A.4). Notably, UC students do not perceive their faculty as significantly more liberal than do students at less selective colleges, despite large measured differences in faculty ideology self-reported by instructors (Table A.11).

While the evidence does not support intentional efforts by faculty to shape student views, unintentional instructional effects remain plausible. UC students may be more likely to enroll in academically-oriented rather than career-focused courses, or may be exposed to different emphases in skills and content. These channels are difficult to isolate empirically, but survey evidence suggests that UC students hold more left-aligned factual beliefs on politically salient topics. Even after controlling for GPA, they are significantly more likely than their CSU counterparts to agree that there is a scientific consensus on anthropogenic climate change, that violent crime has declined over time, and that COVID-19 had a far higher mortality rate than influenza or pneumonia during the pandemic.

Findings from compulsory schooling laws further challenge the instructional mechanism. The largest treatment effects I observe are among a complier population of marginal high school dropouts. Effects are also larger in Florida, a politically conservative context where Republican legislators can regulate curricula. Moreover, significant partisan shifts appear even among Generation X voters, whose education predated the current alignment of party systems. These results suggest that the political externalities of education arise from mechanisms more fundamental than pedagogy or instructional content from a specific era or educational institution.

4.3 Long-Run Mechanisms

Education may shape partisanship indirectly through longer-run mechanisms such as degree attainment, earnings, neighborhood selection, and household composition. These channels are theoretically plausible: educational credentials and career trajectories influence peer groups, and partisan cleavages persist along income, education, and residential lines. I consider the role of each in turn.

The first-stage effects of UC admission I document closely mirror those in Bleemer (2024), who finds that the policy significantly increased bachelor's degree completion, early-career earnings, and graduate school attendance. These outcomes could matter politically by shifting later-life peer environments or through the economic consequences of educational attainment itself. Yet the theoretical direction of income effects on partisanship remains ambiguous. While classical models predict rising income should reduce support for redistribution and shift voters to the right, more recent work suggests that in contexts like college campuses, where sociocultural values may be more salient than economic issues, the income effects of education may be dominated.

To assess whether these long-run changes translate into partisan sorting, I examine neighborhood characteristics for students observed in California's voter file. Figure H.5 and Table H.5 report results on median neighborhood income, education, and party affiliation. Across all measures, there is little evidence of residential sorting at the policy threshold. I likewise find no consistent discontinuity in the partisan affiliation of household members. These null effects suggest that residential context is unlikely to drive the persistent changes in partisanship observed among policy-eligible students.

Evidence from compulsory schooling laws further weakens the case for long-run mechanisms. The largest effects are found among students born after 2002. These cohorts are too young to have accumulated meaningful labor market experience or formed distinct adult residential or professional peer groups.[16] As such, the findings imply that partisan exter-

[16] Although it is worth noting that students at selective UC campuses have coworkers who they perceive

nalities of education emerge well before long-run sorting can occur, reinforcing the case for socialization during schooling as a major mechanism.

5 Conclusion

This paper provides causal evidence that education impacts partisanship. Using natural experiments from college admission rules in California and compulsory schooling laws in Florida and California, I show that educational attainment consistently reduces the probability of Republican Party affiliation and increases Democratic or independent registration. The effects appear rapidly, persist into adulthood, and hold across different institutions, generations, geographies, and demographic groups.

The identification strategies I implement take advantage of clear discontinuities in education exposure at both the top and bottom of the distribution. At the top, I exploit a GPA-based admissions preference at the University of California that increased admission and enrollment at highly selective colleges for students in the top four percent of their high school class. At the bottom, I use birthdate cutoffs embedded in compulsory schooling laws that raised high school attainment by preventing early dropout. These settings differ substantially in institutional type, geography, and baseline political context, yet produce strikingly similar results: additional schooling moves individuals away from Republican registration and toward independent or Democratic affiliation.

The scale and granularity of the data, using millions of individuals and linking administrative educational records and verified voter registration histories, enable a comprehensive test of education's partisan externalities. The rich covariates, detailed enrollment and peer data, and scope for heterogeneity analysis across states, generations, sex, and race offer an opportunity to test generalizability rarely available in the literature.

Mechanistically, the evidence is most consistent with peer socialization during schooling. UC students induced into more selective campuses spend more time in peer-dense residential

as substantially more likely to be liberal than their peers at less selective CSUs (Table A.10).

settings and are immersed in ideologically distinct student bodies. These peer environments differ systematically from those at less selective institutions. The largest partisan externalities in the compulsory schooling design are found among recent cohorts who experienced COVID-era school disruptions, where treatment increased the amount of in-person schooling prior to remote learning, further underscoring the impact of social environments. In contrast, I find little support for instructor or curriculum-based mechanisms: UC faculty are more ideologically liberal, but emphasize research over teaching and report less interest in political or civic influence. Likewise, I find no consistent evidence of long-run sorting into partisan neighborhoods or households.

These findings on mechanisms offer two core insights. First, education's effects on partisanship are durable, generalizable, and emerge well before long-run economic returns materialize. Second, the political externalities of education operate not merely through curriculum or instructors but through fundamental changes in who students live with, talk to, and indirectly learn from in early adulthood. In doing so, education reshapes the social context in which policy views and politics are formed.

While this paper focuses on the largest educational settings in most countries, public universities and state K-12 systems, future work could examine whether similar effects arise in religious schools, identity-focused institutions, or military academies. Understanding how curriculum, field of study, or pedagogy shape views is another promising avenue. More broadly, these findings contribute to our understanding of how educational institutions produce not just human capital, but partisanship in the evolving landscape of economic inequality and democratic conflict.

References

Sule Alan, Ceren Baysan, Mert Gumren, and Elif Kubilay. Building Social Cohesion in Ethnically Mixed Schools: An Intervention on Perspective Taking*. The Quarterly Journal of Economics, 136(4):2147–2194, November 2021. ISSN 0033-5533. doi: 10.1093/qje/qjab009. URL https://doi.org/10.1093/qje/qjab009.

Noam Angrist, Kevin Winseck, Harry Anthony Patrinos, and Joshua Graff Zivin. Human Capital and Climate Change. The Review of Economics and Statistics, pages 1–28, March 2024. ISSN 0034-6535. doi: 10.1162/rest_a_01444. URL https://doi.org/10.1162/rest_a_01444.

Brendan Apfeld, Emanuel Coman, John Gerring, and Stephen Jessee. Higher Education and Cultural Liberalism: Regression Discontinuity Evidence from Romania. The Journal of Politics, pages 000–000, May 2022. ISSN 0022-3816. doi: 10.1086/720644. URL https://vpn.uci.edu/+CSCO+0075676763663A2F2F6A6A6A2E77626865616E79662E68707576706E74622E727168++/doi/full/10.1086/720644. Publisher: The University of Chicago Press.

Richard C. Atkinson and Patricia A. Pelfrey. Rethinking Admissions: US Public Universities in the Post-Affirmative Action Age. October 2004. URL https://escholarship.org/uc/item/2w60b2x4.

David Autor, David Dorn, Gordon Hanson, and Kaveh Majlesi. Importing Political Polarization? The Electoral Consequences of Rising Trade Exposure. American Economic Review, 110(10):3139–3183, October 2020. ISSN 0002-8282. doi: 10.1257/aer.20170011. URL https://www.aeaweb.org/articles?id=10.1257/aer.20170011.

Michael C. Bender. Trump Administration Threatens to Withhold Funds From Public Schools. The New York Times, April 2025. ISSN 0362-4331. URL https://www.nytimes.com/2025/04/03/us/politics/public-school-funding-trump-dei.html.

Michael C. Bender and Stephanie Saul. Trump Administration Sends Harvard a List of Demands to Protect Federal Funds. The New York Times, April 2025. ISSN 0362-4331. URL https://www.nytimes.com/2025/04/03/us/politics/trump-harvard-funding-demands.html.

Stephen B. Billings, Eric Chyn, and Kareem Haggag. The Long-Run Effects of School Racial Diversity on Political Identity. American Economic Review: Insights, 3(3):267–284, September 2021. doi: 10.1257/aeri.20200336. URL https://www.aeaweb.org/articles?id=10.1257/aeri.20200336.

Zachary Bleemer. Affirmative action and its race-neutral alternatives. Journal of Public Economics, 220:104839, April 2023. ISSN 0047-2727. doi: 10.1016/j.jpubeco.2023.104839. URL https://www.sciencedirect.com/science/article/pii/S004727272300021X.

Zachary Bleemer. Top Percent Policies and the Return to Postsecondary Selectivity. 2024.

Giampaolo Bonomi, Nicola Gennaioli, and Guido Tabellini. Identity, Beliefs, and Political Conflict*. The Quarterly Journal of Economics, 136(4):2371–2411, November 2021. ISSN 0033-5533. doi: 10.1093/qje/qjab034. URL https://doi.org/10.1093/qje/qjab034.

Luca Braghieri. Political Correctness, Social Image, and Information Transmission, 2021. URL https://www.aeaweb.org/doi/10.1257/rct.5063-1.1.

Sebastian Calonico, Matias D. Cattaneo, and Rocio Titiunik. Robust Nonparametric Confidence Intervals for Regression-Discontinuity Designs. Econometrica, 82(6):2295–2326, 2014. ISSN 1468-0262. doi: 10.3982/ECTA11757. URL https://onlinelibrary.wiley.com/doi/abs/10.3982/ECTA11757. _eprint: https://onlinelibrary.wiley.com/doi/pdf/10.3982/ECTA11757.

Sebastian Calonico, Matias D Cattaneo, and Max H Farrell. Optimal bandwidth choice for robust bias-corrected inference in regression discontinuity designs. The Econometrics Journal, 23(2):192–210, May 2020. ISSN 1368-4221, 1368-423X. doi: 10.1093/ectj/utz022. URL https://academic.oup.com/ectj/article/23/2/192/5625071.

Colin Campbell and Jonathan Horowitz. Does College Influence Sociopolitical Attitudes? Sociology of Education, 89(1):40–58, January 2016. ISSN 0038-0407. doi: 10.1177/0038040715617224. URL https://doi.org/10.1177/0038040715617224. Publisher: SAGE Publications Inc.

Davide Cantoni, Yuyu Chen, David Y. Yang, Noam Yuchtman, and Y. Jane Zhang. Curriculum and Ideology. Journal of Political Economy, 125(2):338–392, April 2017. ISSN 0022-3808. doi: 10.1086/690951. URL https://www.journals.uchicago.edu/doi/abs/10.1086/690951. Publisher: The University of Chicago Press.

Enrico Cantoni and Vincent Pons. Does Context Outweigh Individual Characteristics in Driving Voting Behavior? Evidence from Relocations within the United States. American Economic Review, 112(4):1226–1272, April 2022. ISSN 0002-8282. doi: 10.1257/aer.20201660. URL https://www.aeaweb.org/articles?id=10.1257/aer.20201660.

Alexander W. Cappelen, Benjamin Enke, and Bertil Tungodden. Universalism: Global Evidence. American Economic Review, 115(1):43–76, January 2025. ISSN 0002-8282. doi: 10.1257/aer.20230038. URL https://www.aeaweb.org/articles?id=10.1257/aer.20230038.

Scott E. Carrell, Mark Hoekstra, and James E. West. The Impact of College Diversity on Behavior toward Minorities. American Economic Journal: Economic Policy, 11(4):159–182, November 2019. ISSN 1945-7731. doi: 10.1257/pol.20170069. URL https://www.aeaweb.org/articles?id=10.1257/pol.20170069.

Matias D. Cattaneo, Michael Jansson, and Xinwei Ma. Manipulation Testing Based on Density Discontinuity. The Stata Journal: Promoting communications on statistics and Stata, 18(1):234–261, March 2018. ISSN 1536-867X, 1536-8734. doi: 10.1177/1536867X1801800115. URL http://journals.sagepub.com/doi/10.1177/1536867X1801800115.

Matias D. Cattaneo, Michael Jansson, and Xinwei Ma. Simple Local Polynomial Density Estimators. Journal of the American Statistical Association, pages 1–7, July 2019. ISSN 0162-1459, 1537-274X. doi: 10.1080/01621459.2019.1635480. URL https://www.tandfonline.com/doi/full/10.1080/01621459.2019.1635480.

Charlotte Cavaille and John Marshall. Education and Anti-Immigration Attitudes: Evidence from Compulsory Schooling Reforms across Western Europe. American Political Science Review, 113(1):254–263, February 2019. ISSN 0003-0554, 1537-5943. doi: 10.1017/S0003055418000588. Publisher: Cambridge University Press.

Jiwon Choi, Ilyana Kuziemko, Ebonya Washington, and Gavin Wright. Local Economic and Political Effects of Trade Deals: Evidence from NAFTA. American Economic Review, 114(6):1540–1575, June 2024. ISSN 0002-8282. doi: 10.1257/aer.20220425. URL https://www.aeaweb.org/articles?id=10.1257/aer.20220425.

Lucia Corno, Eliana La Ferrara, and Justine Burns. Interaction, Stereotypes, and Performance: Evidence from South Africa. American Economic Review, 112(12):3848–3875, December 2022. ISSN 0002-8282. doi: 10.1257/aer.20181805. URL https://www.aeaweb.org/articles?id=10.1257/aer.20181805.

Ernesto Dal Bo, Frederico Finan, Olle Folke, Torsten Persson, and Johanna Rickne. Economic and Social Outsiders but Political Insiders: Swedens Populist Radical Right. The Review of Economic Studies, 90(2):675–706, March 2023. ISSN 0034-6527. doi: 10.1093/restud/rdac037. URL https://doi.org/10.1093/restud/rdac037.

Carlos Dobkin and Fernando Ferreira. Do school entry laws affect educational attainment and labor market outcomes? Economics of Education Review, 29(1):40–54, February 2010. ISSN 0272-7757. doi: 10.1016/j.econedurev.2009.04.003. URL https://www.sciencedirect.com/science/article/pii/S0272775709000685.

The Economist. Hard-right parties are now Europes most popular. The Economist, 2025. ISSN 0013-0613. URL https://www.economist.com/graphic-detail/2025/02/28/hard-right-parties-are-now-europes-most-popular.

Benjamin Enke. Moral Values and Voting. Journal of Political Economy, 128(10):3679–3729, October 2020. ISSN 0022-3808. doi: 10.1086/708857. URL https://www.journals.uchicago.edu/doi/abs/10.1086/708857. Publisher: The University of Chicago Press.

Benjamin Enke, Mattias Polborn, and Alex Wu. Morals as Luxury Goods and Political Polarization. Working Paper 30001, National Bureau of Economic Research, April 2022. URL https://www.nber.org/papers/w30001. Series: Working Paper Series.

Benjamin Enke, Ricardo Rodrguez-Padilla, and Florian Zimmermann. Moral Universalism and the Structure of Ideology. The Review of Economic Studies, 90(4):1934–1962, July 2023. ISSN 0034-6527, 1467-937X. doi: 10.1093/restud/rdac066. URL https://academic.oup.com/restud/article/90/4/1934/6726197.

Benjamin Enke, Raymond Fisman, Luis Mota Freitas, and Steven Sun. Universalism and Political Representation: Evidence from the Field. American Economic Review: Insights, 6(2):214–229, June 2024. ISSN 2640-205X. doi: 10.1257/aeri.20230222. URL https://www.aeaweb.org/articles?id=10.1257/aeri.20230222.

Frederico Finan, Enrique Seira, and Alberto Simpser. Voting with ones neighbors: Evidence from migration within Mexico. Journal of Public Economics, 202:104495, October 2021. ISSN 0047-2727. doi: 10.1016/j.jpubeco.2021.104495. URL https://www.sciencedirect.com/science/article/pii/S0047272721001316.

Daniel Firoozi and Igor Geyn. Do Tuition Subsidies Raise Political Participation? American Economic Journal: Economic Policy, Forthcoming, 2024.

Nicola Gennaioli and Guido Tabellini. Identity Politics. In Presidential Address. Econometric Society, 2025.

Amory Gethin, Clara Martinez-Toledano, and Thomas Piketty. Brahmin Left Versus Merchant Right: Changing Political Cleavages in 21 Western Democracies, 1948-2020. The Quarterly Journal of Economics, (qjab036), October 2021a. ISSN 0033-5533. doi: 10.1093/qje/qjab036. URL https://doi.org/10.1093/qje/qjab036.

Amory Gethin, Clara Martnez-Toledano, and Thomas Piketty, editors. Political cleavages and social inequalities: a study of fifty democracies, 1948-2020. Harvard University Press, Cambridge, Massachusetts London, England, 2021b. ISBN 978-0-674-24842-7.

Yoav Goldstein and Matan Kolerman. Changing Minds: How Academic Fields Shape Political Attitudes, March 2025. URL https://papers.ssrn.com/abstract=5196889.

Mireille Jacobson, Maria Kogelnik, and Heather Royer. Holiday, Just One Day Out of Life: Birth Timing and Post-natal Outcomes, June 2020. URL https://www.nber.org/papers/w27326.

Ethan Kaplan, Jorg L. Spenkuch, and Cody Tuttle. A Different World: Enduring Effects of School Desegregation on Ideology and Attitudes, January 2025. URL https://www.nber.org/papers/w33365.

Clark Kerr, Marian L. Gade, and Maureen Kawaoka. The gold and the blue: a personal memoir of the University of California, 1949-1967, volume 2. University of California Press, Berkeley, 2001. ISBN 978-0-520-22367-7 978-0-520-23641-7.

Herbert P. Kitschelt and Philipp Rehm. Secular Partisan Realignment in the United States: The Socioeconomic Reconfiguration of White Partisan Support since the New Deal Era. Politics & Society, August 2019. doi: 10.1177/0032329219861215. URL https://journals.sagepub.com/doi/abs/10.1177/0032329219861215. Publisher: SAGE PublicationsSage CA: Los Angeles, CA.

Michal Kolesar and Christoph Rothe. Inference in Regression Discontinuity Designs with a Discrete Running Variable. American Economic Review, 108(8):2277–2304, August 2018.

ISSN 0002-8282. doi: 10.1257/aer.20160945. URL https://pubs.aeaweb.org/doi/10.1257/aer.20160945.

Ilyana Kuziemko and Ebonya Washington. Why Did the Democrats Lose the South? Bringing New Data to an Old Debate. American Economic Review, 108(10):2830–2867, October 2018. ISSN 0002-8282. doi: 10.1257/aer.20161413. URL https://www.aeaweb.org/articles?id=10.1257/aer.20161413.

Ilyana Kuziemko, Nicolas Longuet-Marx, and Suresh Naidu. Compensate the Losers? Economic Policy and Partisan Realignment in the US, October 2023. URL https://www.nber.org/papers/w31794.

Sara LaLumia, James M. Sallee, and Nicholas Turner. New Evidence on Taxes and the Timing of Birth. American Economic Journal: Economic Policy, 7(2):258–293, May 2015. ISSN 1945-7731. doi: 10.1257/pol.20130243. URL https://www.aeaweb.org/articles?id=10.1257/pol.20130243.

Juliana Londono-Velez. The impact of diversity on perceptions of income distribution and preferences for redistribution. Journal of Public Economics, 214:104732, October 2022. ISSN 0047-2727. doi: 10.1016/j.jpubeco.2022.104732. URL https://www.sciencedirect.com/science/article/pii/S0047272722001347.

Morgan Marietta and David C. Barker. One Nation, Two Realities: Dueling Facts in American Democracy. Oxford University PressNew York, 1 edition, April 2019. ISBN 978-0-19-067717-6 978-0-19-067720-6. doi: 10.1093/oso/9780190677176.001.0001. URL https://academic.oup.com/book/38739.

John Marshall. Education and Voting Conservative: Evidence from a Major Schooling Reform in Great Britain. The Journal of Politics, 78(2):382–395, April 2016. ISSN 0022-3816. doi: 10.1086/683848. URL https://www.journals.uchicago.edu/doi/10.1086/683848. Publisher: The University of Chicago Press.

John Marshall. The Anti-Democrat Diploma: How High School Education Decreases Support for the Democratic Party. American Journal of Political Science, 63(1):67–83, 2019. ISSN 1540-5907. doi: 10.1111/ajps.12409. URL https://onlinelibrary.wiley.com/doi/abs/10.1111/ajps.12409. _eprint: https://onlinelibrary.wiley.com/doi/pdf/10.1111/ajps.12409.

Justin McCrary. Manipulation of the running variable in the regression discontinuity design: A density test. Journal of Econometrics, 142(2):698–714, February 2008. ISSN 0304-4076. doi: 10.1016/j.jeconom.2007.05.005. URL http://www.sciencedirect.com/science/article/pii/S0304407607001133.

Justin McCrary and Heather Royer. The Effect of Female Education on Fertility and Infant Health: Evidence from School Entry Policies Using Exact Date of Birth. American Economic Review, 101(1):158–195, February 2011. ISSN 0002-8282. doi: 10.1257/aer.101.1.158. URL https://pubs.aeaweb.org/doi/10.1257/aer.101.1.158.

Allan H. Meltzer and Scott F. Richard. A Rational Theory of the Size of Government. Journal of Political Economy, 89(5):914–927, 1981. ISSN 0022-3808. URL https://www.jstor.org/stable/1830813. Publisher: University of Chicago Press.

Tali Mendelberg, Katherine T. McCabe, and Adam Thal. College Socialization and the Economic Views of Affluent Americans. American Journal of Political Science, 61(3):606–623, 2017. ISSN 1540-5907. doi: 10.1111/ajps.12265. URL https://onlinelibrary.wiley.com/doi/abs/10.1111/ajps.12265. _eprint: https://onlinelibrary.wiley.com/doi/pdf/10.1111/ajps.12265.

Taylor Orth. Most college graduates say college makes people more liberal. Technical report, YouGov, 2022. URL https://today.yougov.com/politics/articles/42503-most-say-college-makes-people-more-liberal.

Kim Parker. Views of Higher Education Divided by Party, August 2019. URL https://www.pewresearch.org/social-trends/2019/08/19/the-growing-partisan-divide-in-views-of-higher-education-2/.

Thomas Romer. Individual welfare, majority voting, and the properties of a linear income tax. Journal of Public Economics, 4(2):163–185, February 1975. ISSN 0047-2727. doi: 10.1016/0047-2727(75)90016-X. URL http://www.sciencedirect.com/science/article/pii/004727277590016X.

Ralph Scott. Does university make you more liberal? Estimating the within-individual effects of higher education on political values. Electoral Studies, 77:102471, June 2022. ISSN 0261-3794. doi: 10.1016/j.electstud.2022.102471. URL https://www.sciencedirect.com/science/article/pii/S0261379422000312.

Elizabeth Simon. Demystifying the link between higher education and liberal values: A within-sibship analysis of British individuals attitudes from 19942020. The British Journal of Sociology, n/a(n/a), 2022. ISSN 1468-4446. doi: 10.1111/1468-4446.12972. URL https://onlinelibrary.wiley.com/doi/abs/10.1111/1468-4446.12972. _eprint: https://onlinelibrary.wiley.com/doi/pdf/10.1111/1468-4446.12972.

Reilly Steel. The Political Transformation of Corporate America, 2001-2022, October 2024. URL https://papers.ssrn.com/abstract=4974868.

Logan Strother, Spencer Piston, Ezra Golberstein, Sarah E. Gollust, and Daniel Eisenberg. College roommates have a modest but significant influence on each others political ideology. Proceedings of the National Academy of Sciences, 118(2), January 2021. ISSN 0027-8424, 1091-6490. doi: 10.1073/pnas.2015514117. URL https://www.pnas.org/content/118/2/e2015514117. Publisher: National Academy of Sciences Section: Social Sciences.

Andrew Van Dam. States with the worst brain drain - and more! The Washington Post, September 2022. URL https://www.washingtonpost.com/business/2022/09/09/films-assigned-college/.

Floyd Jiuyun Zhang. Political endorsement by Nature and trust in scientific expertise during COVID-19. Nature Human Behaviour, 7(5):696–706, May 2023. ISSN 2397-3374. doi: 10.1038/s41562-023-01537-5. URL https://www.nature.com/articles/s41562-023-01537-5. Publisher: Nature Publishing Group.

Tables

Table 1: Effects of UC Admission Policy on Registration and Partisanship

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Total Voter Registration Rate						
Registered to Vote	0.0118	0.0110	0.0127^{+}	0.0119^{+}	0.0158	0.0148
	(0.0078)	(0.0078)	(0.0069)	(0.0069)	(0.0102)	(0.0101)
B. Political Party Membership						
Republican Party	-0.0060^{+}	-0.0061^{+}	-0.0061*	-0.0063*	-0.0089*	-0.0091*
	(0.0032)	(0.0032)	(0.0029)	(0.0028)	(0.0043)	(0.0043)
Democrat/Independent	0.0202*	0.0197*	0.0188**	0.0182**	0.0247*	0.0239*
	(0.0080)	(0.0079)	(0.0069)	(0.0069)	(0.0103)	(0.0102)
Democratic Party	0.0107	0.0103	0.0099	0.0097	0.0113	0.0110
	(0.0069)	(0.0069)	(0.0064)	(0.0063)	(0.0093)	(0.0093)
No Party Preference	0.0097^{+}	0.0094^{+}	0.0113*	0.0109*	0.0146^{+}	0.0142^{+}
	(0.0056)	(0.0056)	(0.0049)	(0.0049)	(0.0076)	(0.0076)
Third Party	-0.0025	-0.0025	-0.0024	-0.0024	-0.0013	-0.0013
	(0.0016)	(0.0016)	(0.0016)	(0.0016)	(0.0024)	(0.0024)
C. Early Life Conversion between Major Parties						
Republican Convert	-0.0026**	-0.0025**	-0.0015^{+}	-0.0014^{+}	-0.0027*	-0.0026*
	(0.0010)	(0.0010)	(0.0008)	(0.0008)	(0.0012)	(0.0012)
Democratic Convert	-0.0013	-0.0014	-0.0013	-0.0013	-0.0013	-0.0014
	(0.0014)	(0.0014)	(0.0014)	(0.0014)	(0.0020)	(0.0020)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Democrat/Independent" refers to the fraction of students who are registered as Democrat, as a no party preference voter, or as a member of a third party. Democratic and Republican converts are voters who are currently registered with the Democratic and Republican Party in California, but at any time in the past were a registered member of the other major party.

Table 2: Effects of Compulsory Schooling on Partisanship

	(1)	(2)	(3)	(4)	(5)	(6)
Location	All	All	FL	FL	CA	CA
A. Political Party Membership						
Republican	-0.0042**	-0.0037**	-0.0041**	-0.0050**	-0.0032**	-0.0028**
	(0.0008)	(0.0007)	(0.0012)	(0.0011)	(0.0010)	(0.0010)
Democrat/Independent	0.0042**	0.0037**	0.0041**	0.0050**	0.0032**	0.0028**
	(0.0008)	(0.0007)	(0.0012)	(0.0011)	(0.0010)	(0.0010)
Democrat	0.0034**	0.0024**	0.0030*	0.0036**	0.0016	0.0015
	(0.0009)	(0.0008)	(0.0012)	(0.0012)	(0.0012)	(0.0012)
No Party	0.0003	0.0012	0.0014	0.0017	0.0012	0.0009
	(0.0008)	(0.0008)	(0.0012)	(0.0012)	(0.0010)	(0.0010)
Third Party	0.0005	0.0000	-0.0003	-0.0003	0.0004	0.0004
	(0.0004)	(0.0004)	(0.0005)	(0.0005)	(0.0007)	(0.0007)
Bandwidth	70	70	70	70	70	70
Polynomial	1	1	1	1	1	1
Controls	No	Yes	No	Yes	No	Yes
Sample Size	5,110,316	5,110,316	2,380,489	2,380,489	2,729,827	2,729,827

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors in parentheses. Bandwidths are measured in days relative to the compulsory schooling law cutoff birthdate. Controls include sex, race, state of voter registration, and year of birth fixed effects.

Table 3: Effects of Compulsory Schooling on Partisanship by Generation

	(1)	(2)	(3)	(4)	(5)	(6)
Generation	Gen X	Gen X	Gen Y	Gen Y	Gen Z	Gen Z
A. Political Party Membership						
Republican	-0.0045*	-0.0047*	-0.0029**	-0.0025*	-0.0069**	-0.0056**
	(0.0019)	(0.0018)	(0.0010)	(0.0010)	(0.0014)	(0.0013)
Democrat/Independent	0.0045*	0.0047*	0.0029**	0.0025*	0.0069**	0.0056**
	(0.0019)	(0.0018)	(0.0010)	(0.0010)	(0.0014)	(0.0013)
Democrat	0.0024	0.0023	0.0031**	0.0023*	0.0047**	0.0029+
	(0.0020)	(0.0020)	(0.0011)	(0.0011)	(0.0016)	(0.0016)
No Party	0.0010	0.0014	-0.0003	0.0006	0.0014	0.0022
	(0.0017)	(0.0017)	(0.0011)	(0.0011)	(0.0016)	(0.0016)
Third Party	0.0011	0.0010	0.0001	-0.0005	0.0009	0.0004
	(0.0010)	(0.0010)	(0.0005)	(0.0005)	(0.0008)	(0.0008)
Bandwidth	70	70	70	70	70	70
Polynomial	1	1	1	1	1	1
Controls	No	Yes	No	Yes	No	Yes
Sample Size	951,135	951,135	2,891,260	2,891,260	1,427,996	1,427,996

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors in parentheses. Bandwidths are measured in days relative to the compulsory schooling law cutoff birthdate. Generation X includes all people born between 1969 and 1980, Generation Y includes all people born between 1980 and 995, and Generation Z includes all people born after 1995. Controls include sex, race, state of voter registration, and year of birth fixed effects.

Table 4: Effects of Compulsory Schooling on Partisanship by Sex

	(1)	(2)	(3)	(4)
Sex	Men	Men	Women	Women
A. Political Party Membership				
Republican	-0.0050**	-0.0046**	-0.0035**	-0.0028**
	(0.0011)	(0.0011)	(0.0010)	(0.0010)
Democrat/Independent	0.0050**	0.0046**	0.0035**	0.0028**
	(0.0011)	(0.0011)	(0.0010)	(0.0010)
Democrat	0.0037**	0.0029*	0.0032**	0.0020+
	(0.0012)	(0.0012)	(0.0012)	(0.0012)
No Party	0.0009	0.0018	-0.0003	0.0007
	(0.0012)	(0.0012)	(0.0011)	(0.0011)
Third Party	0.0004	-0.0001	0.0006	0.0002
	(0.0006)	(0.0006)	(0.0006)	(0.0006)
Bandwidth	70	70	70	70
Polynomial	1	1	1	1
Controls	No	Yes	No	Yes
Sample Size	2,454,819	2,454,819	2,655,497	2,655,497

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors in parentheses. Bandwidths are measured in days relative to the compulsory schooling law cutoff birthdate. Controls include sex, race, state of voter registration, and year of birth fixed effects.

Table 5: Effects of Compulsory Schooling on Partisanship by Race and Ethnicity

	(1)	(2)	(3)	(4)
Race	White	White	Minority	Minority
A. Political Party Membership				
Republican	-0.0068**	-0.0062**	-0.0026**	-0.0022*
	(0.0013)	(0.0013)	(0.0009)	(0.0009)
Democrat/Independent	0.0068**	0.0062**	0.0026**	0.0022*
	(0.0013)	(0.0013)	(0.0009)	(0.0009)
Democrat	0.0055**	0.0047**	0.0021+	0.0010
	(0.0013)	(0.0012)	(0.0011)	(0.0011)
No Party	0.0014	0.0019	-0.0005	0.0008
	(0.0012)	(0.0012)	(0.0011)	(0.0010)
Third Party	-0.0002	-0.0005	0.0010+	0.0004
	(0.0007)	(0.0007)	(0.0005)	(0.0005)
Bandwidth	70	70	70	70
Polynomial	1	1	1	1
Controls	No	Yes	No	Yes
Sample Size	2,112,126	2,112,126	2,998,190	2,998,190

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors in parentheses. Bandwidths are measured in days relative to the compulsory schooling law cutoff birthdate. In this table "White" is defined as all registrants self-identifying as non-hispanic white, whereas all other registrants are categorized as "Minority". Controls include sex, race, state of voter registration, and year of birth fixed effects.

Figures

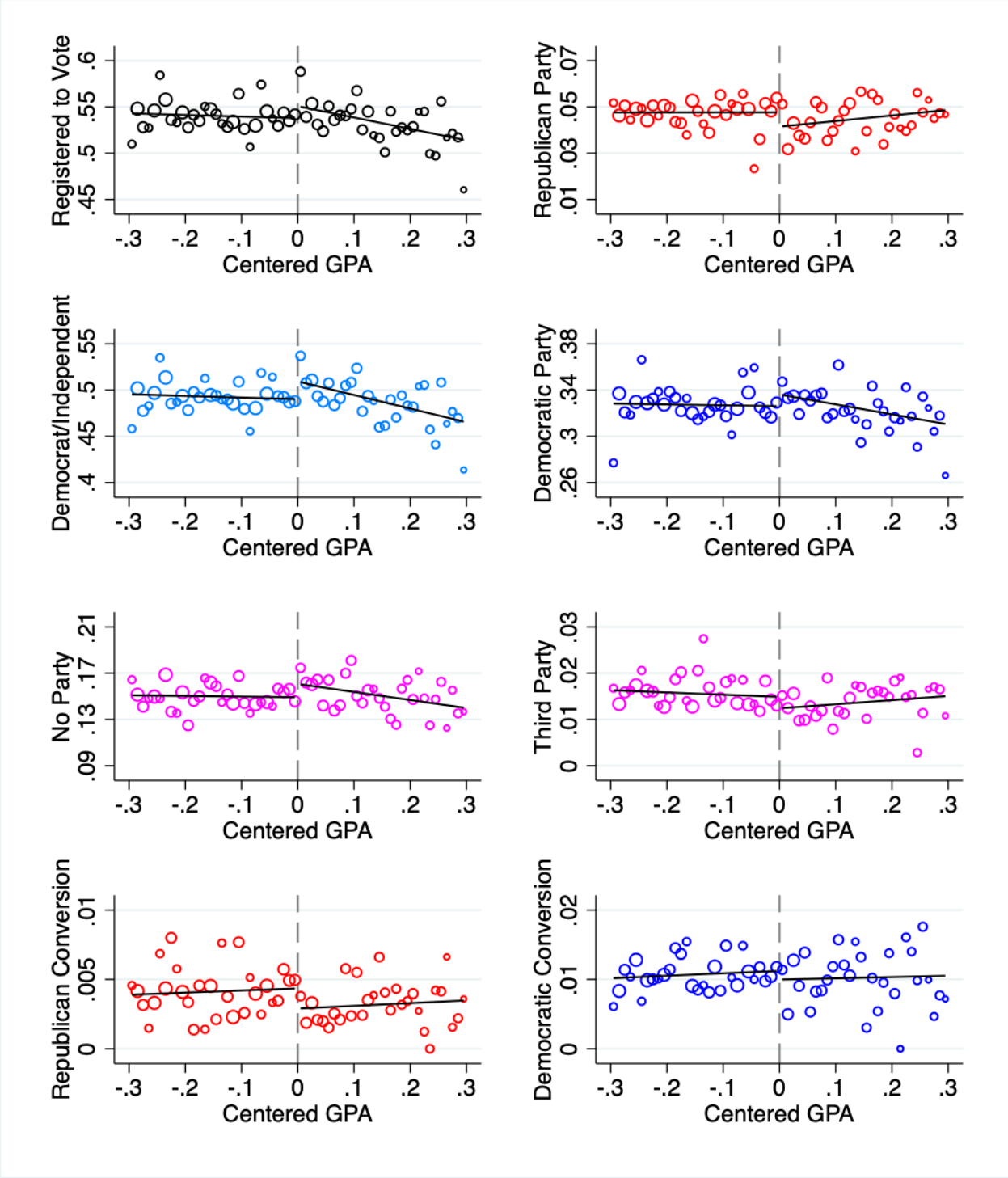

Figure 1: RD Graphs of Voter Registration Outcomes

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table 1.

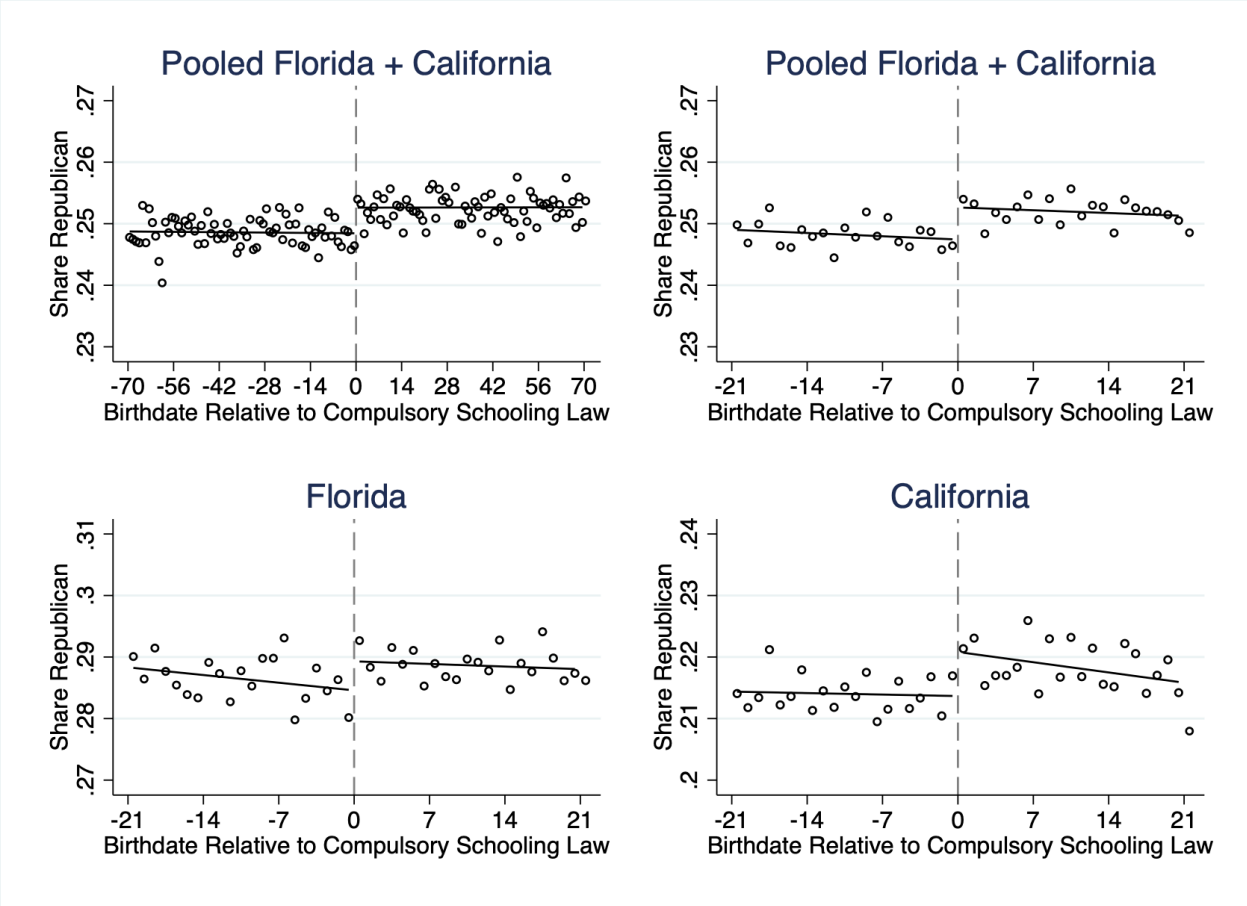

Figure 2: RD Graphs of Compulsory Schooling Laws

Note: Birthdates are normalized relative to the compulsory schooling law cutoff date for school entry.

Online Appendices

A In-Sample Survey Appendix

A.1 Survey Descriptive Statistics

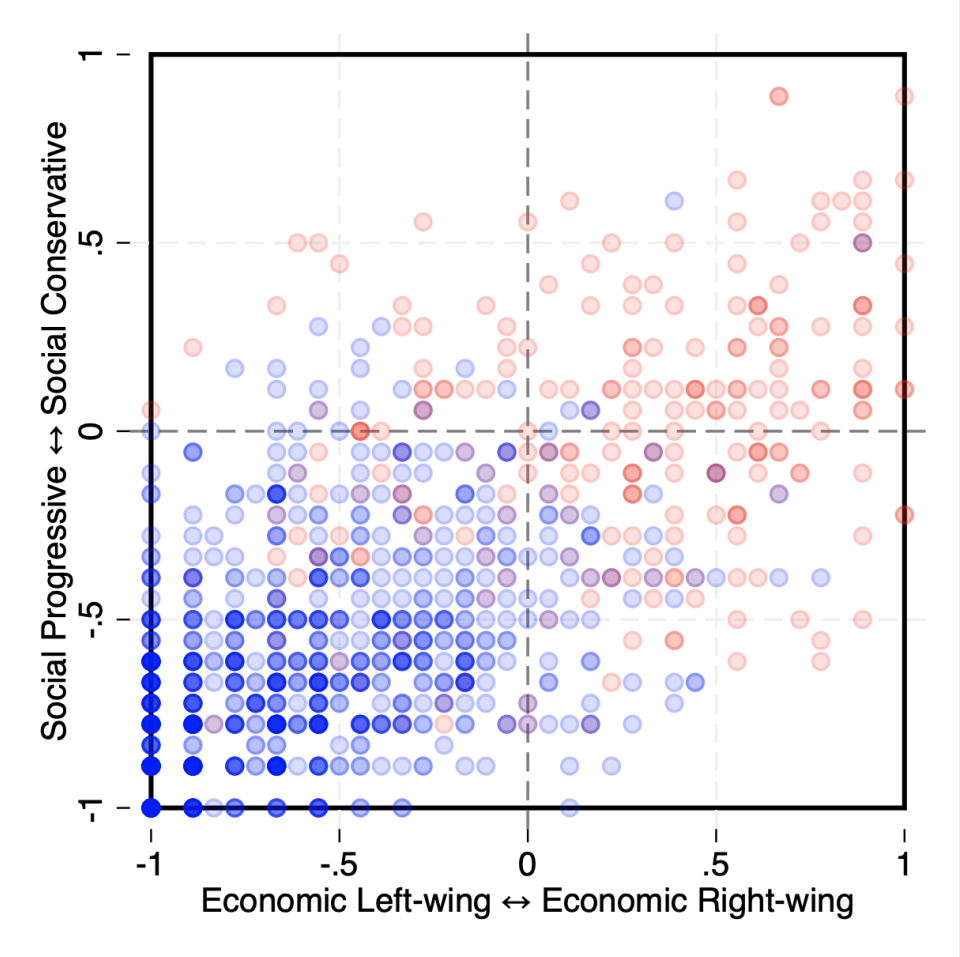

Figure A.1: Poll Respondent Ideologies and Two-Party Preference

Note: The two ideological indexes in this figure are calculated using the questions in Block 3 of the Survey in Section A.2. Index values are calculated as the average policy view on a particular set of questions with the most liberal response assigned -1, the most conservative response assigned +1, and all other responses interpolated at equidistant points. Each dot reflects a point in the two-dimensional ideology space. The darker the color of a dot, the more individuals are located at that particular point. The color gradient from blue to red reflects the proportion of individuals at a given point who say they favor the Democratic Party over the Republican Party on policy issues, with blue dots corresponding to the Democratic Party and red dots corresponding to the Republican Party. Plugging in the median response for each question from Pew Research samples of American voters yields a score of (0,0).

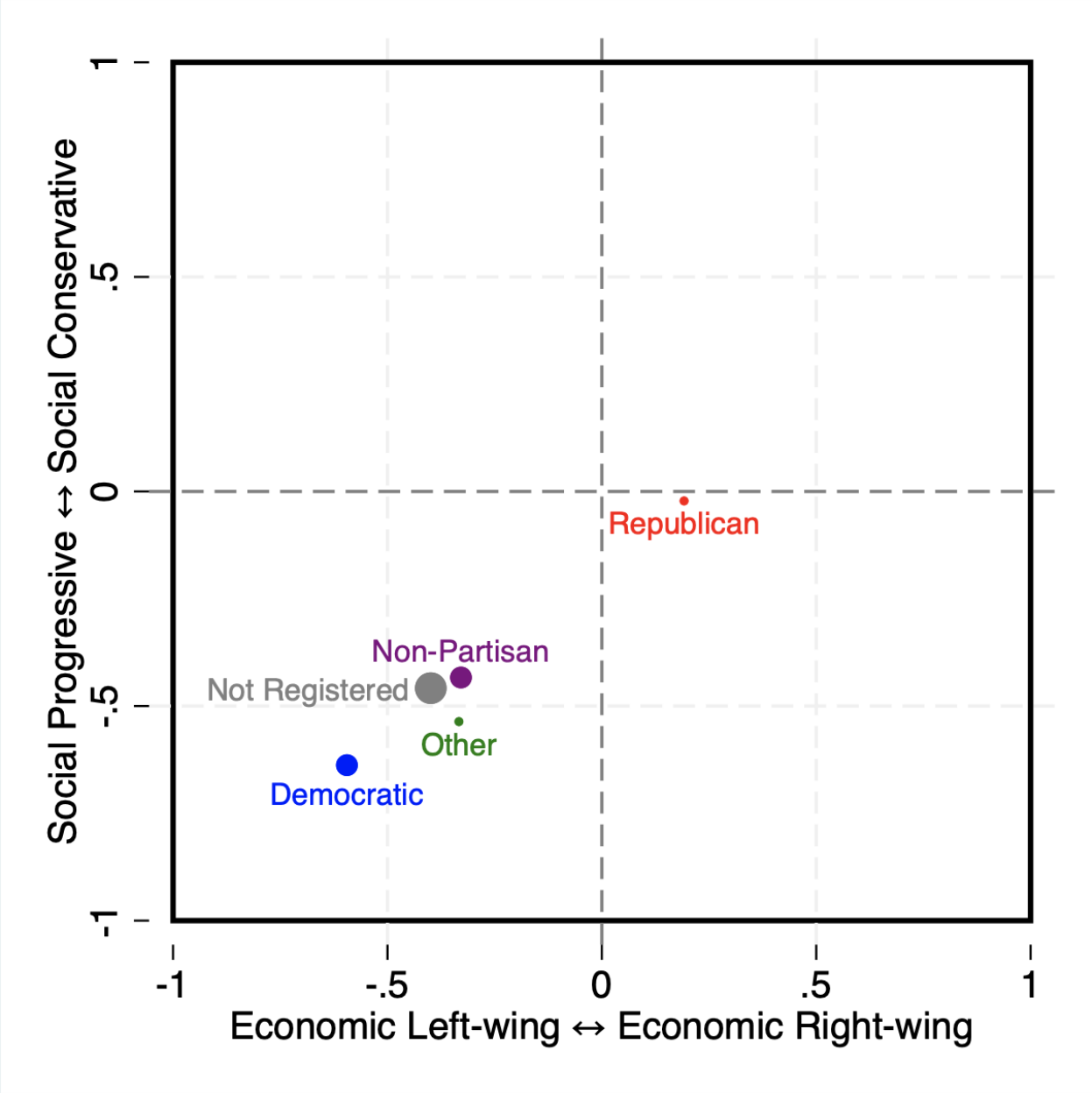

Figure A.2: Mean Respondent Ideology by Voter Registration Status

Note: The two ideological indexes in this figure are calculated using the questions in Block 3 of the Survey in Section A.2. Index values are calculated as the average policy view on a particular set of questions with the most liberal response assigned -1, the most conservative response assigned +1, and all other responses interpolated at equidistant points. Each dot reflects the average ideological scores of a particular voter registration group from in-sample respondents in the two-dimensional ideology space. The dots are sized roughly based on the number of respondents within the particular voter registration group. Plugging in the median response for each question from Pew Research samples of American voters yields a score of (0,0).

Table A.1: Comparison of Survey Takers and Full Sample

Demographics	Respondents	Full Sample
Female	52.2%	52.8%
Underrepresented Minority	25.2%	24.4%
Likely Cal Grant Eligible	38.1%	34.0%
First Generation Student	45.2%	45.1%
FAFSA Filer	68.9%	63.7%
Student Works Pre-College	5.1%	4.7%
Low Enrollment County	5.3%	5.6%
Low Quality High School	14.2%	14.1%
Raised by Single Parent	17.1%	15.6%
Dad's Years of Schooling	14.2 years	14.3 years
Mom's Years of Schooling	14.0 years	14.0 years
ISIR Family Income	$77,594	$81,476
Reported Family Income	$88,413	$93,056
High School GPA	3.61	3.59
Household Size	3.95	4.04
Party Registration	Respondents	Full Sample
No Registration	45.3%	48.9%
Democratic	32.1%	29.5%
No Party	15.6%	14.8%
Republican	4.8%	4.8%
Third Party	2.0%	1.7%
College Sector	Respondents	Full Sample
University of California	56.7%	52.8%
California Sate University	15.2%	15.1%
California Private	7.1%	7.2%
Out-of-State	8.0%	9.2%
2-yr or No College	12.7%	15.3%

Note: The column titled "Respondents" reflects the mean value or percentage among people who participated in my proprietary survey, which was sent to all in-sample UC applicants. The column titled "Full Sample" shows the corresponding value for all UC applicants within my sample, regardless of whether or not they participated in the survey.

Table A.2: Two-Party Policy Preference by Party Registration

	Two-Party Preference		
Registration Status	Republican	Democratic	Total %
Democratic	6.2	93.8	100.0
Non-Partisan	26.6	73.4	100.0
Not Registered	25.7	74.3	100.0
Other	21.7	78.3	100.0
Republican	77.4	22.6	100.0
N	243	862	1,105

Note: The Democratic and Republican columns reflect the proportion of survey respondents with a given voter registration status who say they favor a given major political party on policy issues. "Non-partisan" refers to individuals who are registered to vote, but are unaffiliated with a political party. "Other" refers to individuals who are registered members of third parties.

Table A.3: Ideology Scores by Party Registration

	Mean Ideology	
Registration Status	Economic	Social
Democratic	-0.595	-0.638
Non-Partisan	-0.329	-0.434
Not Registered	-0.399	-0.458
Other	-0.333	-0.536
Republican	0.192	-0.022

Note: The economic and social columns reflect the mean ideological score or survey respondents with a given voter registration status. The two ideological indexes are calculated using the questions in Block 3 of the Survey in Section A.2. Index values are calculated as the average policy view on a particular set of questions with the most liberal response assigned -1, the most conservative response assigned +1, and all other responses interpolated at equidistant points. "Non-partisan" refers to individuals who are registered to vote, but are unaffiliated with a political party. "Other" refers to individuals who are registered members of third parties.

Table A.4: Self-Reported Political Influence Ratings

	Mean Influence Score			
College Enrollment	Family	Friends	Coworkers	Educators
UC	1.91	1.86	3.46	2.77
CSU	1.66	1.96	3.57	2.80
Other CA	1.80	2.03	3.41	2.77
Other OOS	1.72	1.98	3.56	2.74
No 4 Yr	1.86	2.01	3.49	2.64
Total	1.84	1.92	3.48	2.75

Note: Each column reflects the mean self-reported influence rank respondents assign to a particular group. The most influential group is assigned the value 1, the second most influential is assigned the value 2, the third is assigned 3, and the least influential is assigned the value 4. Responses are sort into rows by the college enrollment category of an individual in the fall term following their application to the UC system. "Other OOS" refers to out-of-state four year colleges.

Table A.5: Self-Reported Current Events Discussions with Family

	Discusses Current Events with Family				
College Enrollment	Rarely	Yearly	Monthly	Weekly	Total
UC	11.2	9.1	32.4	47.4	100.0
CSU	8.9	7.7	34.3	49.1	100.0
Other CA	11.4	5.1	32.9	50.6	100.0
Other OOS	9.0	5.6	32.6	52.8	100.0
No 4 Yr	9.2	2.8	29.8	58.2	100.0
Total	10.4	7.5	32.4	49.7	100.0

Note: Each column reflects the percent of respondents who say they discussed current events with the stated frequency. Responses are sort into rows by the college enrollment category of an individual in the fall term following their application to the UC system. "Other OOS" refers to out-of-state four year colleges.

Table A.6: Self-Reported Current Events Discussions in College

	Discussed Current Events in College				
College Enrollment	Rarely	Yearly	Monthly	Weekly	Total
UC	16.4	4.8	26.5	52.3	100.0
CSU	15.4	7.7	26.6	50.3	100.0
Other CA	13.9	2.5	25.3	58.2	100.0
Other OOS	13.5	4.5	14.6	67.4	100.0
No 4 Yr	15.6	5.0	19.9	59.6	100.0
Total	15.7	5.1	24.6	54.6	100.0

Note: Each column reflects the percent of respondents who say they discussed current events with the stated frequency. Responses are sort into rows by the college enrollment category of an individual in the fall term following their application to the UC system. "Other OOS" refers to out-of-state four year colleges.

Table A.7: Self-Reported Current Events Discussions with Friends

	Discusses Current Events with Friends				
College Enrollment	Rarely	Yearly	Monthly	Weekly	Total
UC	4.3	4.6	29.2	61.9	100.0
CSU	7.1	2.4	34.9	55.6	100.0
Other CA	1.3	10.1	29.1	59.5	100.0
Other OOS	6.7	0.0	29.2	64.0	100.0
No 4 Yr	7.1	3.5	30.5	58.9	100.0
Total	5.1	4.2	30.2	60.5	100.0

Note: Each column reflects the percent of respondents who say they discussed current events with the stated frequency. Responses are sort into rows by the college enrollment category of an individual in the fall term following their application to the UC system. "Other OOS" refers to out-of-state four year colleges.

Table A.8: Self-Reported College Student Housing

	Ever Lived with Students		
College Enrollment	Yes	No	Total
UC	82.8	17.2	100.0
CSU	57.4	42.6	100.0
Other CA	84.8	15.2	100.0
Other OOS	88.8	11.2	100.0
No 4 Yr	58.2	41.8	100.0
Total	76.4	23.6	100.0

Note: Each column reflects the percent of respondents who say they have or have not ever lived in on-campus student housing or in a housing complex mostly composed of college students. Responses are sort into rows by the college enrollment category of an individual in the fall term following their application to the UC system. "Other OOS" refers to out-of-state four year colleges.

Table A.9: Self-Reported Pereptions of Friend Ideology

	Perceived Friend Ideology			
College Enrollment	Liberal	Moderate	Conservative	Total
UC	63.2	31.1	5.7	100.0
CSU	54.4	36.7	8.9	100.0
Other CA	64.6	30.4	5.1	100.0
Other OOS	64.0	31.5	4.5	100.0
No 4 Yr	56.0	36.9	7.1	100.0
Total	61.1	32.7	6.2	100.0

Note: Each column reflects the percent of respondents who would use the respective ideological label to characterize their friends. Responses are sort into rows by the college enrollment category of an individual in the fall term following their application to the UC system. "Other OOS" refers to out-of-state four year colleges.

Table A.10: Self-Reported Pereptions of Coworker Ideology

	Perceived Coworker Ideology			
College Enrollment	Liberal	Moderate	Conservative	Total
UC	38.3	47.7	14.0	100.0
CSU	33.1	47.9	18.9	100.0
Other CA	40.5	39.2	20.3	100.0
Other OOS	41.6	46.1	12.4	100.0
No 4 Yr	42.6	48.2	9.2	100.0
Total	38.5	47.1	14.5	100.0

Note: Each column reflects the percent of respondents who would use the respective ideological label to characterize their coworkers. Responses are sort into rows by the college enrollment category of an individual in the fall term following their application to the UC system. "Other OOS" refers to out-of-state four year colleges.

Table A.11: Self-Reported Pereptions of Educator Ideology

	Perceived Educator Ideology			
College Enrollment	Liberal	Moderate	Conservative	Total
UC	57.6	38.9	3.5	100.0
CSU	55.0	37.3	7.7	100.0
Other CA	50.6	44.3	5.1	100.0
Other OOS	51.7	42.7	5.6	100.0
No 4 Yr	56.0	38.3	5.7	100.0
Total	56.0	39.3	4.7	100.0

Note: Each column reflects the percent of respondents who would use the respective ideological label to characterize their professors or teachers. Responses are sort into rows by the college enrollment category of an individual in the fall term following their application to the UC system. "Other OOS" refers to out-of-state four year colleges.

A.2 Survey Questions

Survey Block 1

Question 1. Indicate how often you have: (Select one option in each row)

	Rarely	**Yearly**	**Monthly**	**Weekly**
Discussed current events with friends				
Discussed current events with family				
Discussed current events during college				
Demonstrated or volunteered for a cause				
Attended religious services				

Question 2. Rank the following groups of people based on how big of an impact you feel they had on your political views. (Drag and drop to move them. 1 means largest impact, 4 means smallest impact.)
Your Professors or Teachers
Your Friends
Your Family
Your Coworkers

Question 3. At roughly what age would you say that you developed most of your social and economic views?
Before age 18
Ages 18 to 21
Ages 21 to 24
Ages 24 to 30
After age 30

Question 4. Have you ever lived in an on-campus college dormitory or in a housing complex mostly composed of college students?
Yes
No

Question 5. If you had to choose, which party is more closely aligned with your policy views?
the Republican Party
the Democratic Party

Survey Block 2

Question 6. Compared to other Americans, would you say that members of {Unselected choice from Question 5} are more, about the same, or less... (Select one option in each row)

	More	**About the Same**	**Less**
Moral			
Open-minded			
Intelligent			

Question 7. To the best of your knowledge, which the following claims are true and which are false? (Select one option in each row)

	True	**False**
COVID killed over 5 times as many Americans as the flu and pneumonia last year.		
Over 95% of climate scientists agree that humans are causing global warming and climate change.		
The violent crime and murder rates were lower last year than 30 years ago.		
More than 75% of immigrants currently in the US are living in the country legally.		
Over 90% of expert economists believe gas price changes are predominantly due to market forces, not government policy.		

Question 8. Which of the following best describes the beliefs of... (Select one option in each row)

	Liberal	**Moderate**	**Conservative**
Your Family			
Your Friends			
Your Coworkers			
Your Professors or Teachers			
Yourself			

Survey Block 3

Question 9. Which of the following statements comes closest to your overall view of gun laws in the United States?

Gun laws should be MORE strict than they are today

Gun laws are about right

Gun laws should be LESS strict than they are today

Question 10. Do you think abortion should be...?

Legal in all cases, no exceptions

Legal in most cases, some exceptions

Illegal in most cases, some exceptions

Illegal in all cases, no exceptions

Question 11. When it comes to transgender people which statement comes closest to your views, even if neither is exactly right?

Someone's gender can be different from the sex they were assigned at birth

Someone's gender is determined by the sex they were assigned at birth

Question 12. Which comes closest to your views about what needs to be done to ensure equal rights for all Americans regardless of their racial or ethnic backgrounds, even if none are exactly right?

Most U.S. laws and major institutions need to be completely rebuilt because they are fundamentally biased against some racial and ethnic groups

While there are many inequities in U.S. laws and institutions, necessary changes can be made by working within the current systems

Little needs to be done

Nothing at all needs to be done

Question 13. Should LEGAL immigration into the United States be...?

Increased

Kept at present level

Decreased

Question 14. Do you favor or oppose the death penalty for people convicted of murder?

Strongly Favor

Somewhat Favor

Somewhat Oppose

Strongly Oppose

Question 15. Thinking about the assistance government provides to people in need, do you think the government...?
Should provide MORE assistance
Is providing about the right amount of assistance
Should provide LESS assistance

Question 16. Thinking about the country's energy supply, do you think the US should...?
Phase out the use of fossil fuels completely, relying instead on renewable sources only
Use a mix of energy sources including fossil fuels along with renewable energy sources

Question 17. Would you favor or oppose making tuition at public colleges and universities free for all American students?
Strongly Favor
Somewhat Favor
Somewhat Oppose
Strongly Oppose

Question 18. Do you think it is the responsibility of the federal government to make sure all Americans have health care coverage?
Yes, it should be provided through a single national health insurance system run by the government
Yes, it should be provided through a mix of private insurance companies and government programs
No, but government should continue programs like Medicare and Medicaid for seniors and the very poor
No, government should not be involved in providing health insurance at all

Question 19. Would you favor or oppose raising the federal minimum wage to $15.00 an hour?
Strongly Favor
Somewhat Favor
Somewhat Oppose
Strongly Oppose

Question 20. If you had to choose, would you rather have a smaller government providing fewer services, or a bigger government providing more services?
Bigger government, more services
Smaller government, fewer services

B Descriptives and First-Stage Appendix

I focus on reduced-form effects, because scoring above the 96th percentile threshold impacts multiple dimensions of admission and enrollment, violating the exclusion restriction.[17] I also present IV estimates using total UC admissions as the treatment to help interpret magnitudes. I view total UC admissions as a better measure of treatment than a binary variable for admission to any UC or enrollment outcomes, because of substitution between UC campuses and less severe exclusion restriction violations. Substitution between UC campuses matters because there are meaningful within-system differences in campus characteristics that may act as causal mechanisms. The exclusion restriction is violated for many enrollment measures because enrollment changes along multiple dimensions.[18]

I begin by illustrating the impact of the UC's top percentile policy on UC applications and admissions. The UC conferred a significant advantage to college applicants who ranked marginally above the 96th percentile of reweighted GPA. I illustrate this visually in Figure B.1 by plotting against students' centered GPA values (1) the number of UC campuses to which they applied in gray and (2) the number of UC campuses to which students were admitted in black. Just below the threshold for eligibility, the typical student applied to roughly 4.5 and was admitted to just under 3 UC campuses. While there is a discrete jump in the number of admissions per student, there is no comparable change in the total number of UC applications.

I show the estimates for these outcomes explicitly in Table B.2, varying the inclusion of covariate controls, the order of a polynomial control for the running variable, and the bandwidth used between 0.3 GPA points and the MSE-optimal bandwidth (Calonico et al., 2020). I find consistently across specifications that there are no meaningful or statistically significant changes in UC application rates at a 90 percent confidence interval. However, there is a sizable discontinuity in UC admission rates on the order of roughly 0.4 campuses at the threshold. Although I prefer the reduced-form estimates throughout this paper, I use this admission effect as a first-stage to understand the scale of the top percentile policy's impact.

Turning to enrollment, I demonstrate that the UC's top percentile admission policy

[17] It is worth noting that reduced-form effects are also preferable because the composition of a prospective student's college applications, which includes campuses outside of the UC system, is unobserved.

[18] This leads the net changes I observe for any single measure of enrollment to understate the gross proportion of applicants who change their enrollment decision. The result would be both overstated IV estimates for enrollment and the potential for misattribution of the effect to one particular enrollment characteristic, when another is more consequential. Aggregate UC admissions faces an similar but less severe problem if eligibility for top percentile admission increases college application rates outside of the UC or changes the composition of UCs to which eligible students apply.

changes the enrollment patterns of policy-eligible students along multiple dimensions. Conferring an admission advantage at the UC increased enrollment at both the extensive margin of four-year college attendance and the intensive margin of selectivity, in part by attracting students to highly selective UC campuses from CSUs and, to a smaller degree, from less selective UCs, two-year colleges, or non-enrollment in college.

In Figure B.2, I illustrate the effect of the UC's top percentile policy on UC application success rates, and enrollment in UCs, CSUs, private Californian colleges, out-of-state colleges, and two-year colleges or no college enrollment.[19] The final two panels in the figure decompose four year colleges by a collapsed version of Opportunity Insights' selectivity ratings.[20] I find that student enrollment rises at highly selective colleges and UCs, primarily at the expense of CSUs, less selective colleges, and non-enrollment in college. Specifically, I find a 6 percentage point increase in highly selective colleges, driven by UCs in this category, with one quarter of counterfactual enrollment coming from two-year colleges or non-enrollment, half coming from CSUs, and the final quarter coming from UCs below the "Highly Selective" category. Figure B.3 illustrates that students flow to UC campuses with higher instructional expenditures, applicant rejection rates, timely graduation rates, and median graduate earnings than countercactual institutions. I demonstrate robustness by testing each of these intermediate outcomes across six different specifications in Tables B.3 and B.4, finding similar results across each.

[19] UC application success rates refer to the ratio between the number of UC campuses an applicant was admitted to and the number of UC campuses to which they applied.

[20] Four-year colleges rated highly selective or better are categorized as "Highly Selective", four year colleges rated selective or below are labeled "Selective", and all other enrollment categories are grouped into "2 Year/No College".

Table B.1: Student Ideology and Partisanship by College Selectivity

A. Ideology of American College Seniors by Selectivity				
Mean SAT Percentile	Left	Middle	Right	Total
>90th Percentile	40.90	36.77	22.33	100.00
75th to 90th	37.51	39.09	23.40	100.00
50th to 75th	29.44	42.11	28.44	100.00
<50th Percentile	27.03	45.73	27.25	100.00
B. Ideology of Californian College Seniors by Selectivity				
Mean SAT Percentile	Left	Middle	Right	Total
>90th Percentile	44.10	38.46	17.45	100.00
75th to 90th	39.41	39.31	21.28	100.00
50th to 75th	27.85	36.11	36.04	100.00
<50th Percentile	25.40	40.55	34.06	100.00
C. Partisanship of In-Sample UC Applicants by Selectivity				
Mean SAT Percentile	Democratic	Neither	Republican	Total
>90th Percentile	60.51	32.27	7.22	100.00
75th to 90th	58.84	31.60	9.55	100.00
50th to 75th	56.97	32.89	10.14	100.00
<50th Percentile	55.56	33.20	11.24	100.00

Note: Panels A and B use data on self-reported student ideology and selectivity data from the CIRP College Senior Survey maintained by UCLA's Higher Education Research Institute. "Left" denotes liberal or far-left, "Middle" denotes middle-of-the-road, and "Right" denotes conservative or far-right. Panel C uses L2 party registration data assessed 10 to 14 years later among my in-sample UC applicants combined with college characteristics from Opportunity Insights. "Neither" denotes students who are registered voters, but do not affiliate with a major party. In all panels, "Mean SAT Percentile" reflects the percentile rank of a campus's average SAT score relative to all campuses within the national UCLA HERI sample, weighted by student population.

Table B.2: Effects of the UC Top Percent Policy on First Stage Outcomes

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
UC Applications	-0.0088	-0.0138	-0.0267	-0.0308	0.0202	0.0087
	(0.0210)	(0.0205)	(0.0198)	(0.0192)	(0.0284)	(0.0277)
UC Admissions	0.4153**	0.4043**	0.3784**	0.3749**	0.4542**	0.4425**
	(0.0277)	(0.0268)	(0.0216)	(0.0207)	(0.0309)	(0.0298)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "UC Applications" refers to the aggregate number of UC campuses to which an applicant applied. "UC Admissions" refers to the aggregate number of UC campuses to which an applicant was admitted.

Table B.3: Effects of the UC Top Percent Policy on Admission and Enrollment

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Admission Outcomes						
UC Success Rate	0.0947**	0.0940**	0.0900**	0.0899**	0.0999**	0.0991**
	(0.0050)	(0.0049)	(0.0038)	(0.0037)	(0.0054)	(0.0053)
B. Enrollment Decomposed by Sector						
UC	0.0339**	0.0332**	0.0318**	0.0309**	0.0382**	0.0391**
	(0.0075)	(0.0073)	(0.0068)	(0.0065)	(0.0098)	(0.0094)
CSU	-0.0399**	-0.0391**	-0.0289**	-0.0282**	-0.0433**	-0.0422**
	(0.0052)	(0.0051)	(0.0041)	(0.0041)	(0.0059)	(0.0058)
Other CA	0.0042	0.0038	0.0026	0.0023	0.0040	0.0032
	(0.0040)	(0.0040)	(0.0040)	(0.0039)	(0.0057)	(0.0057)
Other OOS	0.0162**	0.0148**	0.0086*	0.0082*	0.0185**	0.0169**
	(0.0049)	(0.0048)	(0.0039)	(0.0038)	(0.0055)	(0.0054)
2 Year/No College	-0.0147**	-0.0140**	-0.0141**	-0.0132**	-0.0174**	-0.0171**
	(0.0041)	(0.0040)	(0.0036)	(0.0036)	(0.0053)	(0.0052)
C. Four Year Enrollment Decomposed by Selectivity						
Highly Selective	0.0743**	0.0739**	0.0604**	0.0588**	0.0834**	0.0816**
	(0.0085)	(0.0082)	(0.0064)	(0.0062)	(0.0093)	(0.0089)
Selective	-0.0587**	-0.0572**	-0.0463**	-0.0456**	-0.0660**	-0.0645**
	(0.0073)	(0.0071)	(0.0059)	(0.0058)	(0.0085)	(0.0083)
2 Year/No College	-0.0147**	-0.0140**	-0.0141**	-0.0132**	-0.0174**	-0.0171**
	(0.0041)	(0.0040)	(0.0036)	(0.0036)	(0.0053)	(0.0052)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). The "UC Success Rate" refers to the ratio between the number of UC campuses an individual applied to and the number of UC campuses to which they were actually admitted. "Other OOS" refers to out-of-state four year colleges. "Highly Selective" refers to four year colleges classified by Opportunity Insights ratings as Highly Selective, Elite, or Ivy Plus. "Selective" refers to four year colleges classified by Opportunity Insights ratings as Selective or a lower catgeory of selectivity.

Table B.4: Effects of the UC Top Percent Policy on Enrollment by Quality and Selectivity

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
Instr. Spending	2646.47**	2594.58**	2496.89**	2456.31**	2962.02**	2898.69**
	(235.59)	(228.96)	(184.78)	(178.38)	(263.99)	(257.02)
Rejection Rate	0.0395**	0.0388**	0.0378**	0.0371**	0.0445**	0.0435**
	(0.0041)	(0.0040)	(0.0032)	(0.0031)	(0.0046)	(0.0044)
Graduation Rate	0.0311**	0.0305**	0.0290**	0.0281**	0.0367**	0.0359**
	(0.0040)	(0.0039)	(0.0032)	(0.0030)	(0.0046)	(0.0044)
Median Income	1997.58**	1958.35**	1765.27**	1743.78**	2242.71**	2187.53**
	(176.18)	(169.12)	(136.62)	(130.32)	(197.44)	(189.78)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Instr. Spending" refers to average per student instructional expenditures. "Rejection Rate" refers to the fraction of applicants to a particular campus who were rejected. "Graduation Rate" refers to the proportion of first time full-time freshmen who enter a given campus who complete their intended degree within 150 percent of normative time to degree. "Median Income" in this context refers to the median post-enrollment earnings for students who attended a given campus. Data are from Opportunity Insights.

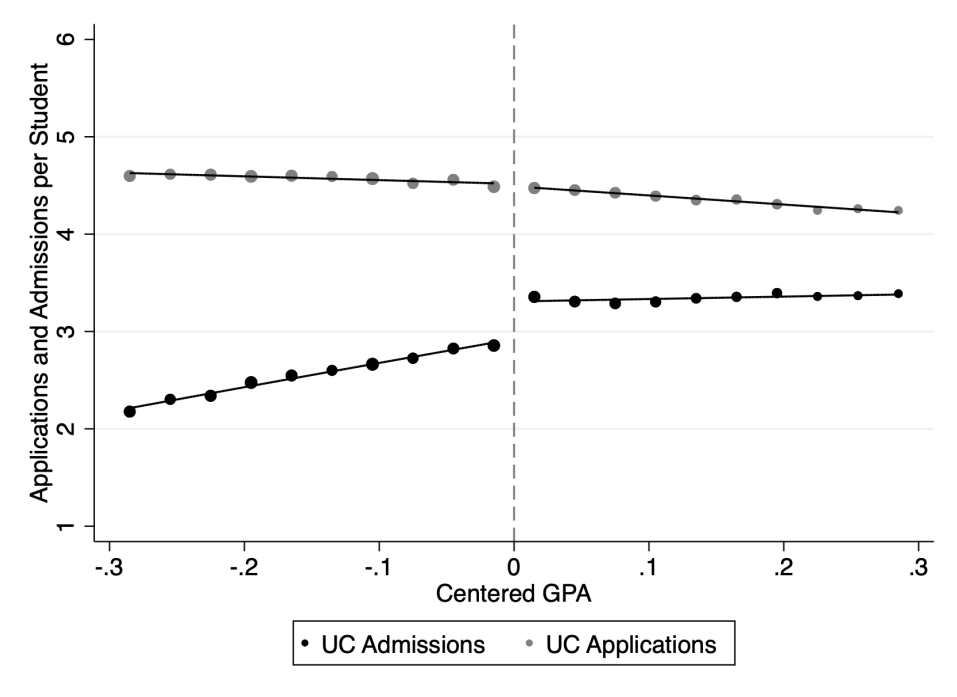

Figure B.1: RD Graph of UC Applications and Admissions

Note: Gray dots reflect the number of UC applications per student. Black dots reflect the number of UC admissions per student. Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort.

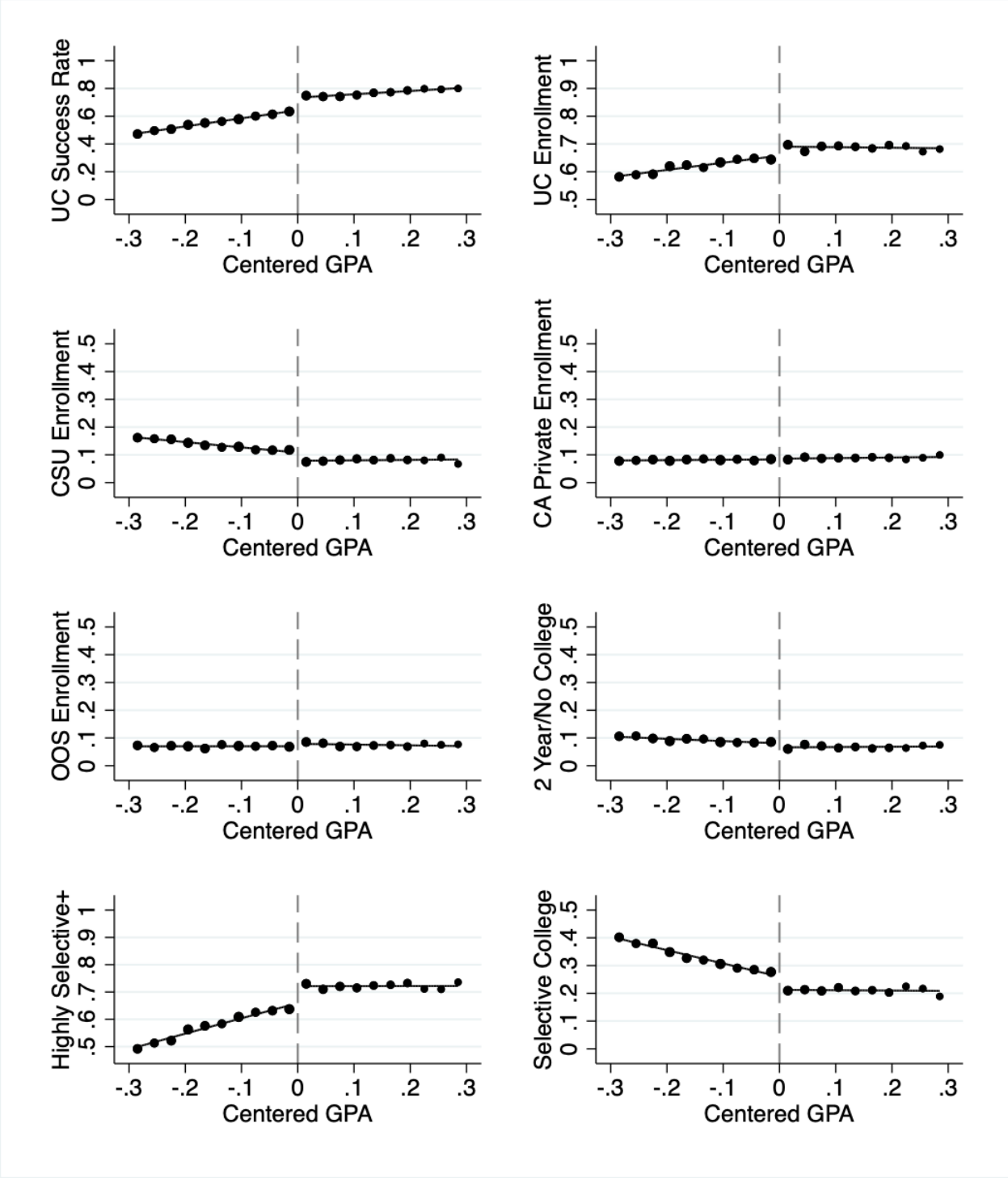

Figure B.2: RD Graphs of College Enrollment

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table B.3.

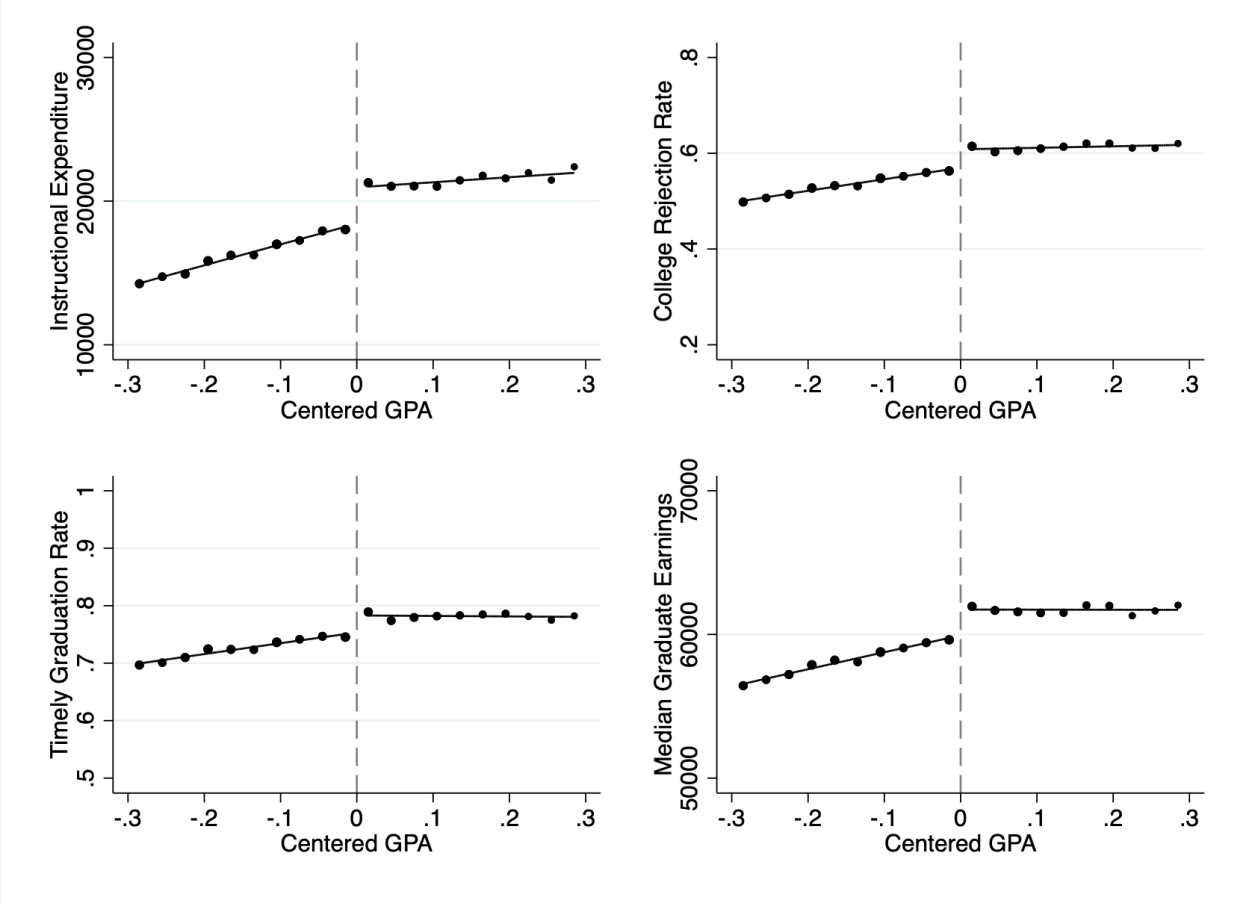

Figure B.3: RD Graphs of College Quality

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table B.4.

C IV Estimates Appendix

Table C.1: IV Estimates of Effects on Voter Registration Outcomes

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Total Voter Registration Rate						
Registered to Vote	0.0113	0.0092	0.0336^{+}	0.0317^{+}	0.0347	0.0335
	(0.0166)	(0.0170)	(0.0184)	(0.0185)	(0.0225)	(0.0229)
B. Political Party Membership						
Republican Party	-0.0129^{+}	-0.0139*	-0.0162*	-0.0169*	-0.0196*	-0.0206*
	(0.0069)	(0.0070)	(0.0075)	(0.0076)	(0.0095)	(0.0097)
Democrat/Independent	0.0349*	0.0334^{+}	0.0498**	0.0486**	0.0543*	0.0541*
	(0.0173)	(0.0175)	(0.0184)	(0.0185)	(0.0228)	(0.0233)
Democratic Party	0.0117	0.0113	0.0263	0.0259	0.0249	0.0249
	(0.0154)	(0.0156)	(0.0167)	(0.0168)	(0.0206)	(0.0210)
No Party Preference	0.0218^{+}	0.0210^{+}	0.0300*	0.0292*	0.0322^{+}	0.0321^{+}
	(0.0120)	(0.0123)	(0.0130)	(0.0132)	(0.0168)	(0.0172)
Third Party	-0.0067	-0.0067	-0.0065	-0.0065	-0.0029	-0.0029
	(0.0040)	(0.0041)	(0.0042)	(0.0043)	(0.0053)	(0.0054)
C. Early Life Conversion between Major Parties						
Republican Convert	-0.0018	-0.0016	-0.0038^{+}	-0.0038^{+}	-0.0059*	-0.0060*
	(0.0018)	(0.0018)	(0.0021)	(0.0021)	(0.0028)	(0.0028)
Democratic Convert	-0.0025	-0.0028	-0.0033	-0.0036	-0.0029	-0.0032
	(0.0033)	(0.0033)	(0.0037)	(0.0037)	(0.0044)	(0.0045)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Democrat/Independent" refers to the fraction of students who are registered as Democrat, as a no party preference voter, or as a member of a third party. Democratic and Republican converts are voters who are currently registered with the Democratic and Republican Party in California, but at any time in the past were a registered member of the other major party. Crossing the 96th percentile threshold is used as the excluded instrument for the number of UC campuses to which an individual was admitted.

Table C.2: IV Estimates of Effects on Voter Turnout Outcomes

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Total Voter Turnout Rates						
Ever Voted	0.0056	0.0036	0.0202	0.0184	0.0307	0.0294
	(0.0162)	(0.0165)	(0.0182)	(0.0183)	(0.0222)	(0.0227)
Total Votes Cast	0.0995	0.0951	0.1406	0.1353	0.1512	0.1446
	(0.0764)	(0.0782)	(0.0896)	(0.0903)	(0.1104)	(0.1128)
B. Presidential and Midterm Election Votes						
Presidential Votes	0.0642	0.0603	0.0985	0.0947	0.1031	0.0994
	(0.0541)	(0.0554)	(0.0626)	(0.0632)	(0.0767)	(0.0783)
Midterm Votes	0.0358	0.0328	0.0420	0.0406	0.0480	0.0452
	(0.0283)	(0.0289)	(0.0320)	(0.0322)	(0.0401)	(0.0410)
C. General and Primary Election Votes						
General Votes	0.0324	0.0280	0.0564	0.0525	0.0475	0.0427
	(0.0519)	(0.0530)	(0.0584)	(0.0589)	(0.0715)	(0.0729)
Primary Votes	0.0807*	0.0790*	0.0842*	0.0828*	0.1036*	0.1018*
	(0.0338)	(0.0345)	(0.0378)	(0.0381)	(0.0468)	(0.0479)
D. Partisan Primary Turnout Rates						
Republican Primaries	-0.0054	-0.0064	-0.0057	-0.0064	-0.0048	-0.0057
	(0.0084)	(0.0085)	(0.0088)	(0.0089)	(0.0107)	(0.0110)
Democratic Primaries	0.0434*	0.0436*	0.0488*	0.0489*	0.0580*	0.0584*
	(0.0213)	(0.0217)	(0.0231)	(0.0232)	(0.0284)	(0.0290)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Voted" refers to the extensive margin of ever having cast a ballot in a regularly scheduled federal election and "votes" refers to the aggregate number of ballots cast by an individual in a regularly scheduled federal election. Republican and Democratic primaries refer to the total ballots cast in partisan presidential primary elections. Crossing the 96th percentile threshold is used as the excluded instrument for the number of UC campuses to which an individual was admitted.

D RD Validation Appendix

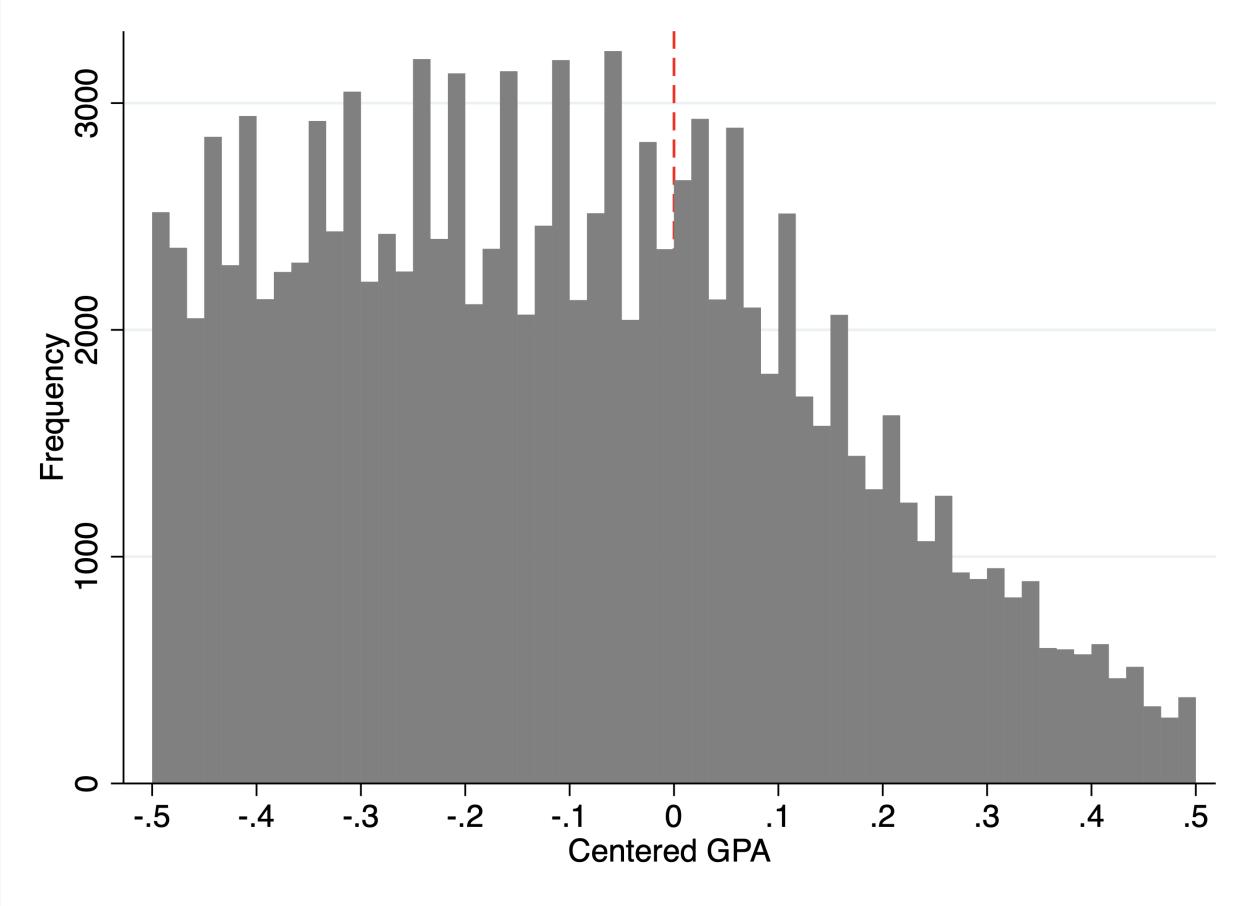

Figure D.1: McCrary Test

Note: This figure displays density of observations across the reweighted GPA normalized to the 96th percentile cutoff within a high school cohort.

Table D.1: Balance Checks for Predicted Voter Registration Outcomes

Outcome	(1)	(2)	(3)
A. Total Voter Registration Rate			
Predicted Voter Registration	0.0008	0.0008	0.0006
	(0.0008)	(0.0008)	(0.0012)
B. Political Party Membership			
Predicted Republican	0.0002	0.0001	0.0001
	(0.0002)	(0.0002)	(0.0003)
Predicted Dem/Ind	0.0008	0.0006	0.0004
	(0.0007)	(0.0007)	(0.0010)
Predicted Democrat	0.0003	0.0003	0.0002
	(0.0006)	(0.0006)	(0.0009)
Predicted No Party	0.0004	0.0003	0.0003
	(0.0004)	(0.0005)	(0.0007)
Predicted Third Party	-0.0000	-0.0000	-0.0000
	(0.0001)	(0.0001)	(0.0001)
C. Midlife Conversion Between Major Parties			
Predicted Republican Conversion	-0.0000	-0.0000	-0.0001^{+}
	(0.0000)	(0.0000)	(0.0000)
Predicted Democrat Conversion	0.0000	0.0000	0.0001
	(0.0000)	(0.0000)	(0.0001)
Bandwidth	Optimal	0.3	0.3
Polynomial	1	1	2
Sample Size	Varies	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

Table D.2: Balance Checks for Predicted Voter Turnout Outcomes

Outcome	(1)	(2)	(3)
A. Total Voter Turnout Rates			
Predicted Voter	0.0010	0.0007	0.0006
	(0.0008)	(0.0008)	(0.0012)
Predicted Votes Cast	0.0027	0.0019	0.0023
	(0.0041)	(0.0043)	(0.0062)
B. Presidential and Midterm Election Votes			
Predicted Regular Votes	0.0017	0.0013	0.0012
	(0.0028)	(0.0030)	(0.0043)
Predicted Midterm Votes	0.0007	0.0005	0.0011
	(0.0013)	(0.0013)	(0.0019)
C. General and Primary Election Votes			
Predicted General Votes	0.0016	0.0011	0.0012
	(0.0027)	(0.0028)	(0.0040)
Predicted Primary Votes	0.0011	0.0007	0.0012
	(0.0014)	(0.0015)	(0.0022)
D. Partisan Primary Turnout Rates			
Predicted Republican Primary Votes	0.0002	0.0002	0.0003
	(0.0002)	(0.0002)	(0.0003)
Predicted Democratic Primary Votes	0.0002	0.0002	0.0002
	(0.0008)	(0.0008)	(0.0012)
Bandwidth	Optimal	0.3	0.3
Polynomial	1	1	2
Sample Size	Varies	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

Table D.3: Balance Checks for Predicted Partisanship (Conditional on Registration)

Outcome	(1)	(2)	(3)
Predicted Republican	0.0001	0.0001	0.0002
	(0.0004)	(0.0004)	(0.0005)
Predicted Dem/Ind	-0.0001	-0.0001	-0.0002
	(0.0004)	(0.0004)	(0.0005)
Predicted Democrat	-0.0001	-0.0003	-0.0004
	(0.0010)	(0.0011)	(0.0015)
Predicted No Party	0.0002	0.0002	0.0003
	(0.0007)	(0.0007)	(0.0011)
Predicted Third Party	-0.0001	-0.0001	-0.0001
	(0.0001)	(0.0001)	(0.0001)
Bandwidth	Optimal	0.3	0.3
Polynomial	1	1	2
Sample Size	Varies	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Results correspond to those in Table D.1, but with outcomes that are predicted using partisanship conditional on registration rather than unconditional partisanship. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

Table D.4: Covariate Balance Checks

Outcome	(1)	(2)	(3)
Female	0.0024	-0.0012	0.0028
	(0.0070)	(0.0067)	(0.0099)
URM	-0.0066	-0.0045	-0.0103
	(0.0053)	(0.0056)	(0.0083)
Cal Grant	-0.0041	-0.0030	-0.0093
	(0.0066)	(0.0066)	(0.0095)
First Generation	-0.0039	-0.0022	-0.0101
	(0.0070)	(0.0067)	(0.0095)
Dad's Schooling	0.0760^{+}	0.0541	0.1105^{+}
	(0.0403)	(0.0418)	(0.0594)
Mom's Schooling	0.0202	0.0202	0.0722
	(0.0408)	(0.0408)	(0.0574)
Dad's Info Missing	0.0024	0.0038	-0.0016
	(0.0037)	(0.0039)	(0.0058)
Mom's Info Missing	-0.0022	-0.0020	-0.0042
	(0.0030)	(0.0033)	(0.0049)
FAFSA Filed	0.0016	0.0021	-0.0053
	(0.0057)	(0.0061)	(0.0089)
Application Year	0.0127	0.0190	0.0300
	(0.0177)	(0.0185)	(0.0240)
ISIR Income	1128.7592	826.7595	2124.8770
	(1065.6962)	(1090.0484)	(1597.6353)
ISIR Missing	-0.0035	-0.0043	0.0037
	(0.0058)	(0.0062)	(0.0090)
Self-Reported Income	986.2342	666.3863	809.9436
	(1219.9420)	(1159.4477)	(1651.7142)
No Income Self-Report	-0.0002	-0.0011	0.0073
	(0.0047)	(0.0051)	(0.0076)
Bandwidth	Optimal	0.3	0.3
Polynomial	1	1	2
Sample Size	Varies	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020).

Table D.5: Covariate Balance Checks

Outcome	(1)	(2)	(3)
Household Size	-0.0110	-0.0063	-0.0145
	(0.0121)	(0.0130)	(0.0193)
Low Quality HS	-0.0011	-0.0011	0.0025
	(0.0050)	(0.0048)	(0.0063)
Low Enrollment County	-0.0000	0.0004	-0.0027
	(0.0030)	(0.0031)	(0.0040)
Student Worker	0.0003	0.0001	-0.0012
	(0.0013)	(0.0013)	(0.0018)
Bandwidth	Optimal	0.3	0.3
Polynomial	1	1	2
Sample Size	Varies	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020).

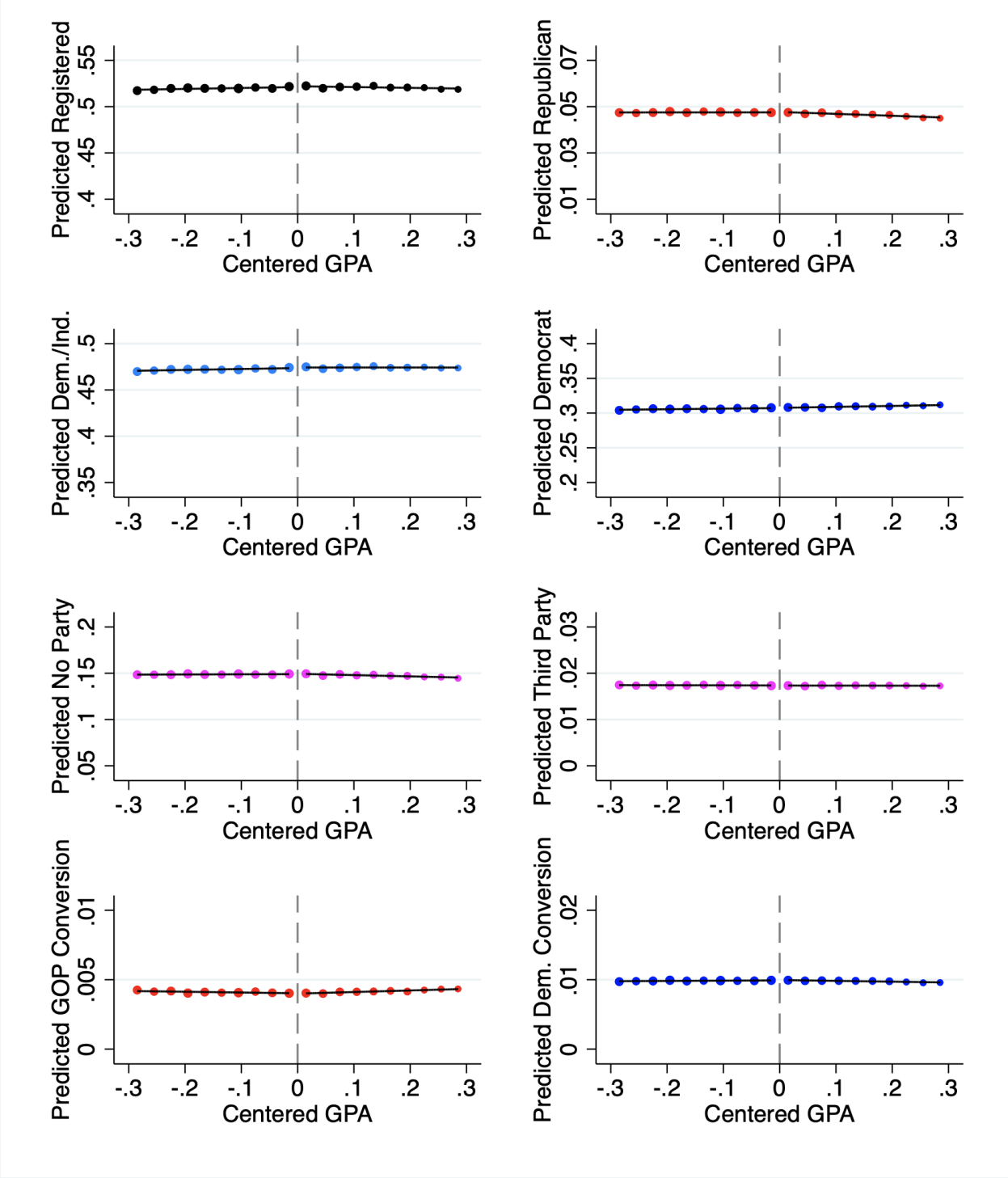

Figure D.2: RD Graph of Predicted Voter Registration Outcomes

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

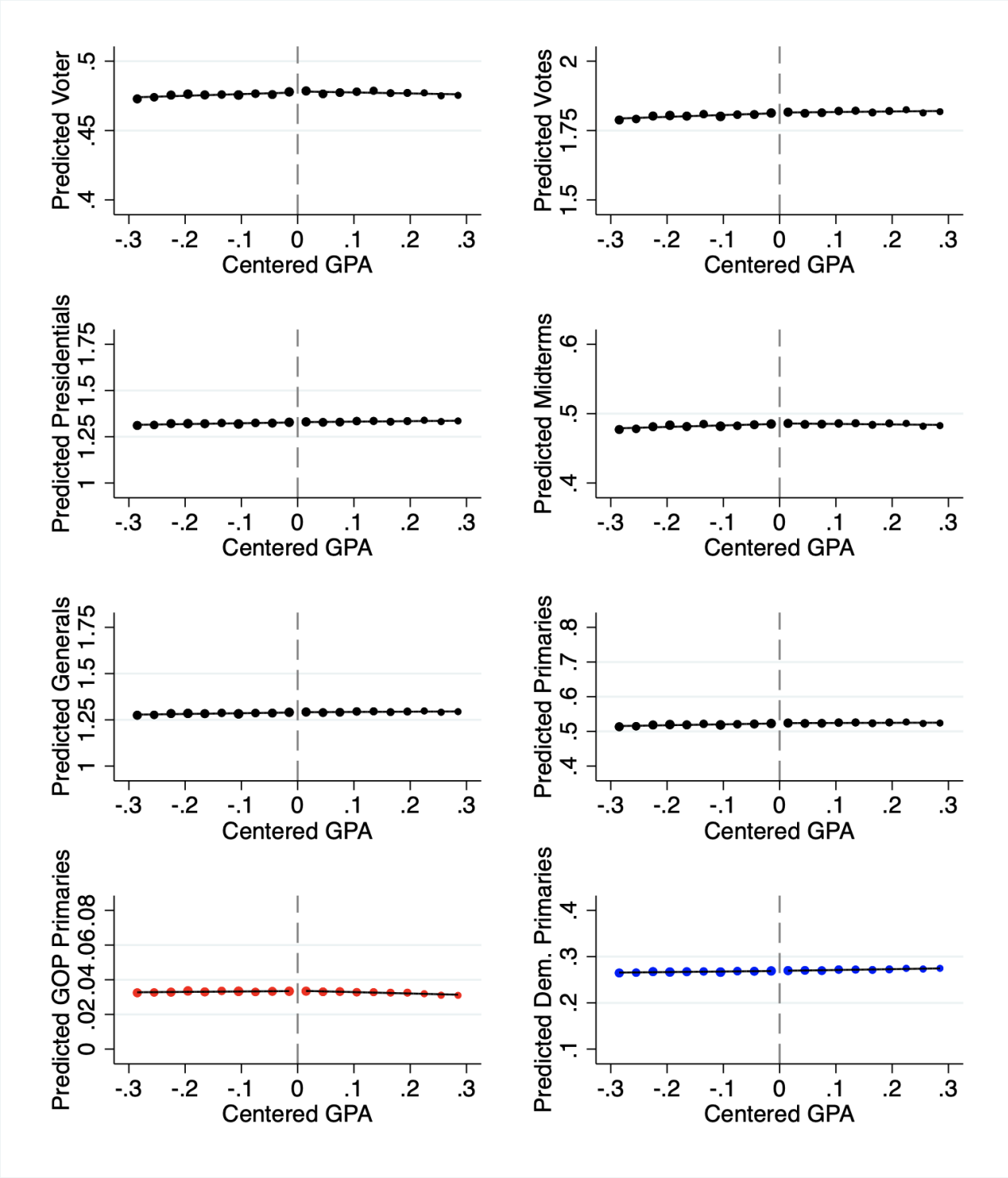

Figure D.3: RD Graph of Predicted Voter Turnout Outcomes

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

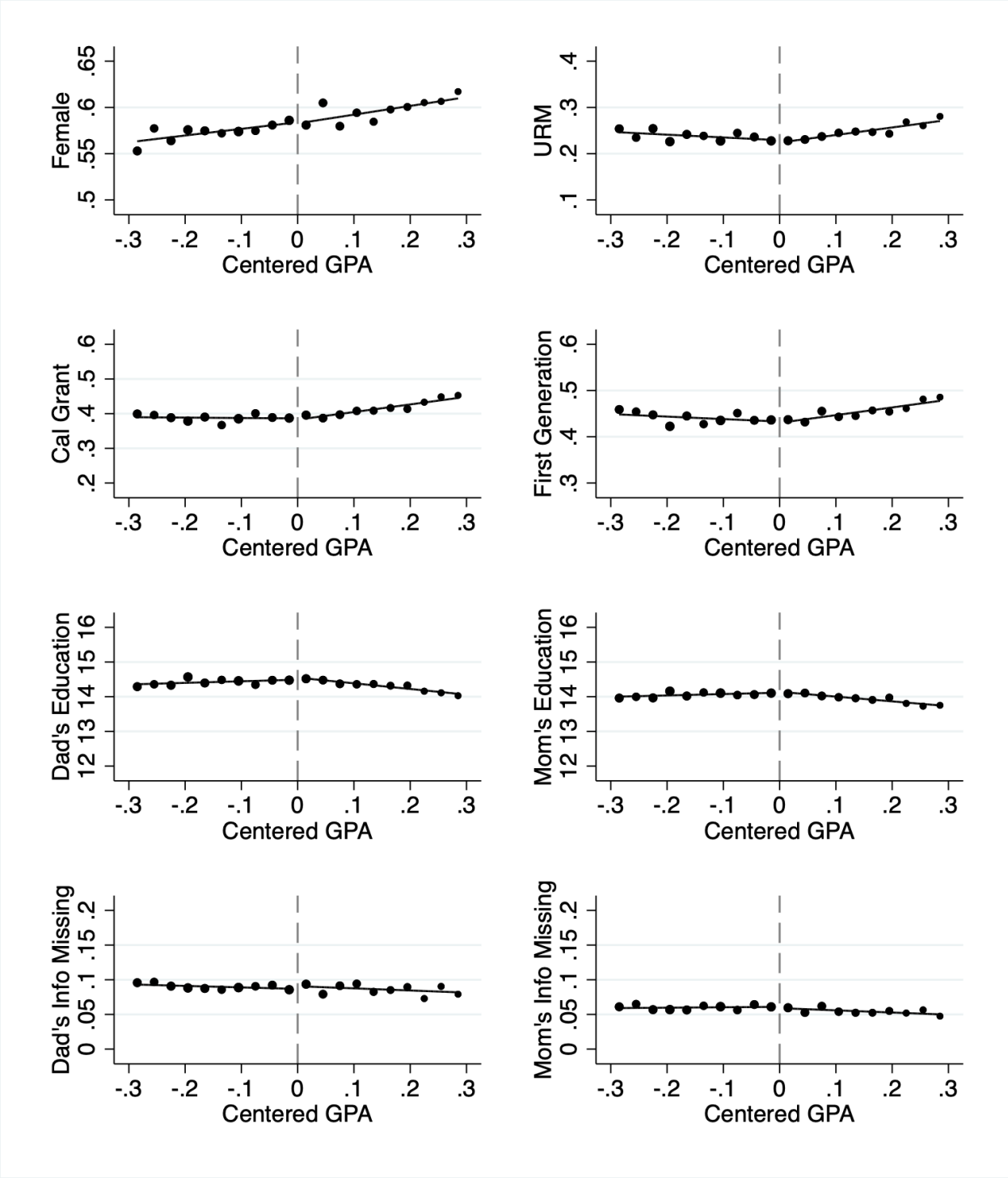

Figure D.4: Covariate RD Graphs

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort.

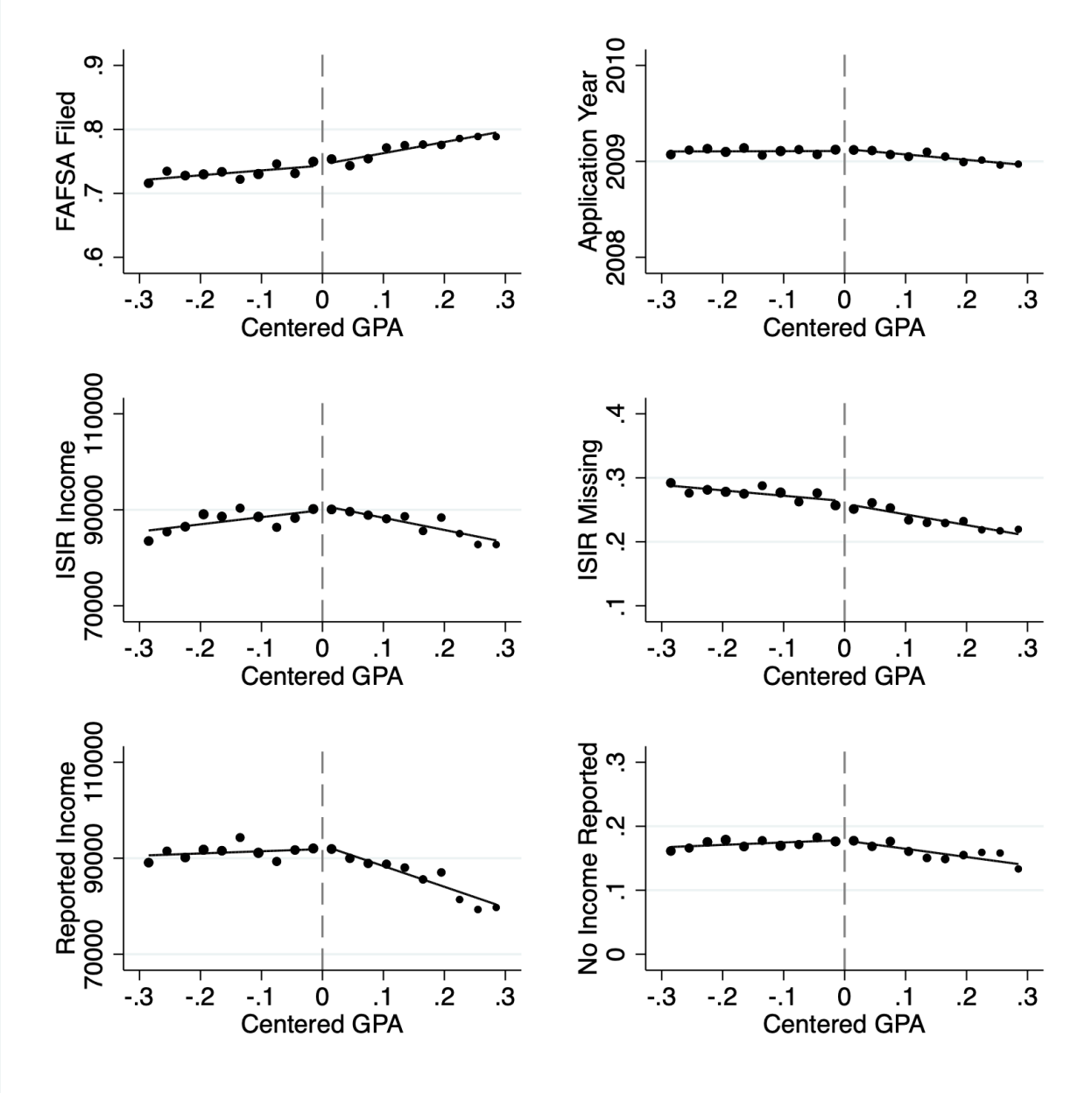

Figure D.5: Covariate RD Graphs

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort.

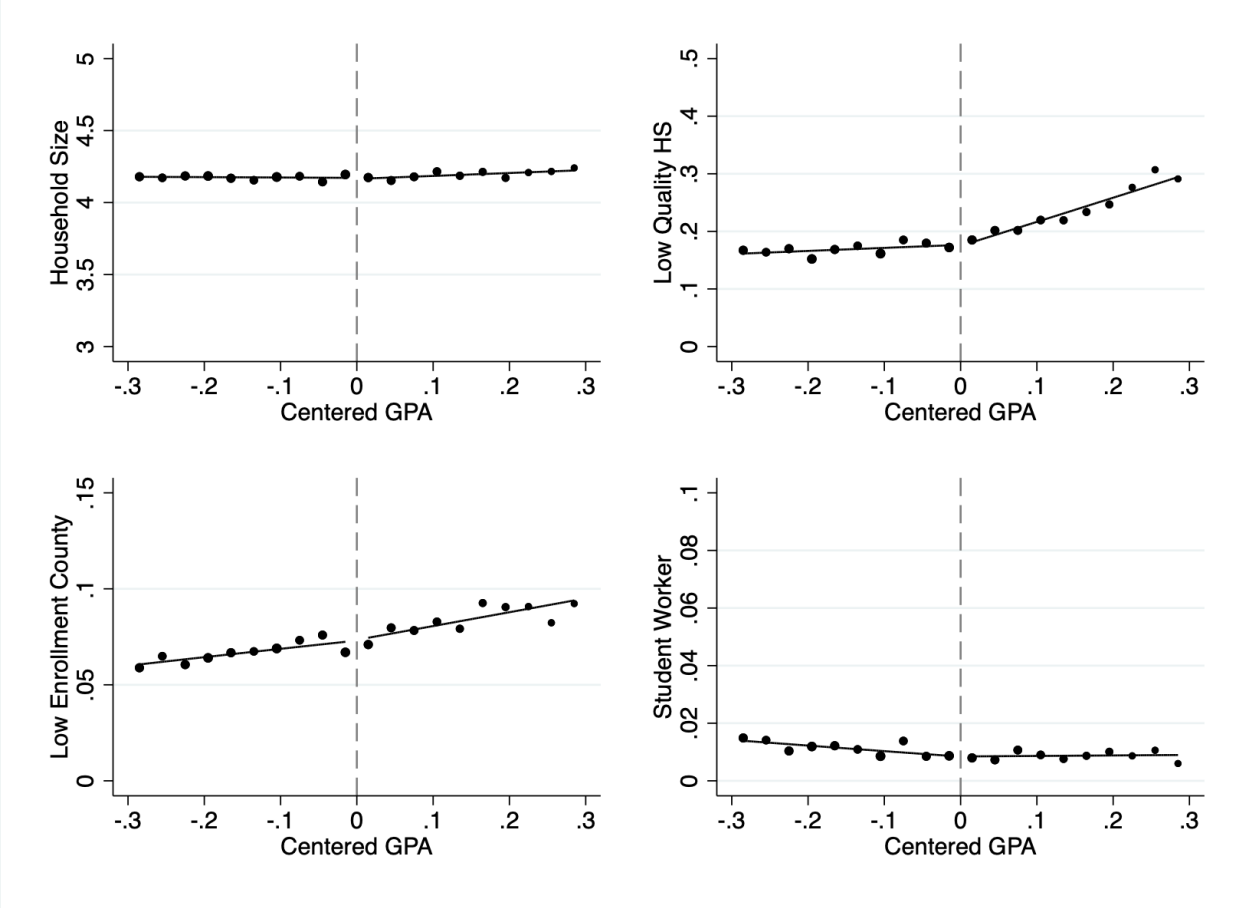

Figure D.6: Covariate RD Graphs

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort..

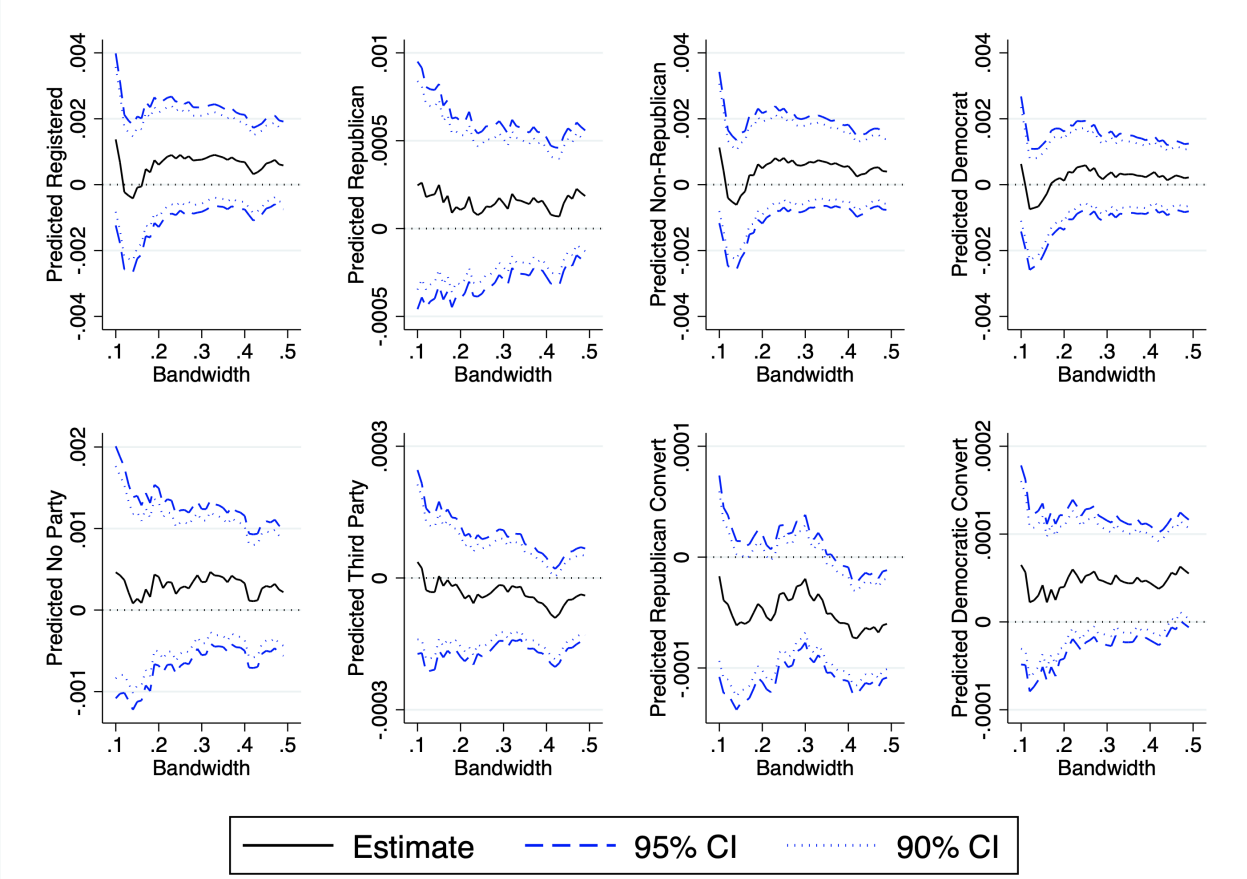

Figure D.7: Predicted Outcome Bandwidth Graphs

Note: Each graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the discontinuity at the threshold in a given predicted outcome using a local linear specification at a respective bandwidth. Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

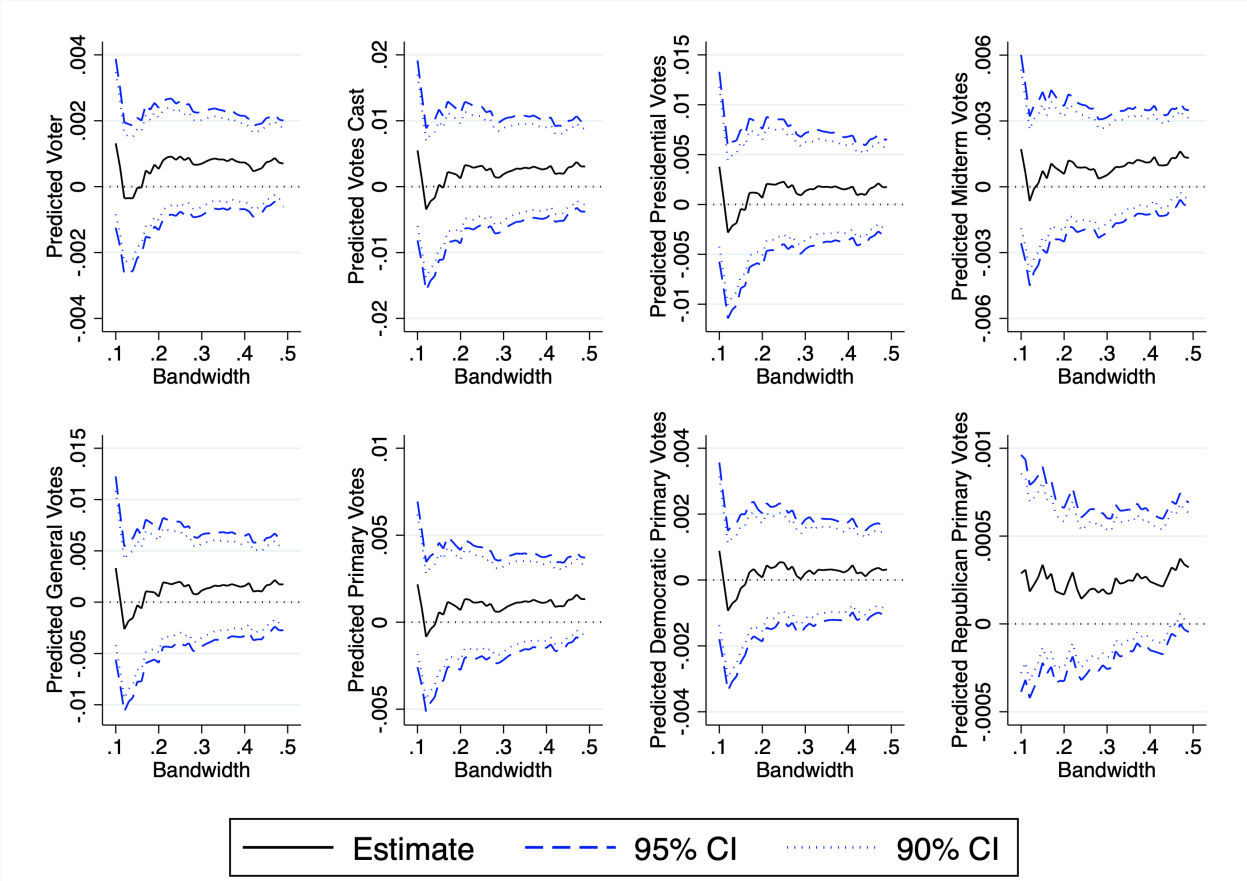

Figure D.8: Predicted Outcome Bandwidth Graphs

Note: Each graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the discontinuity at the threshold in a given predicted outcome using a local linear specification at a respective bandwidth. Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

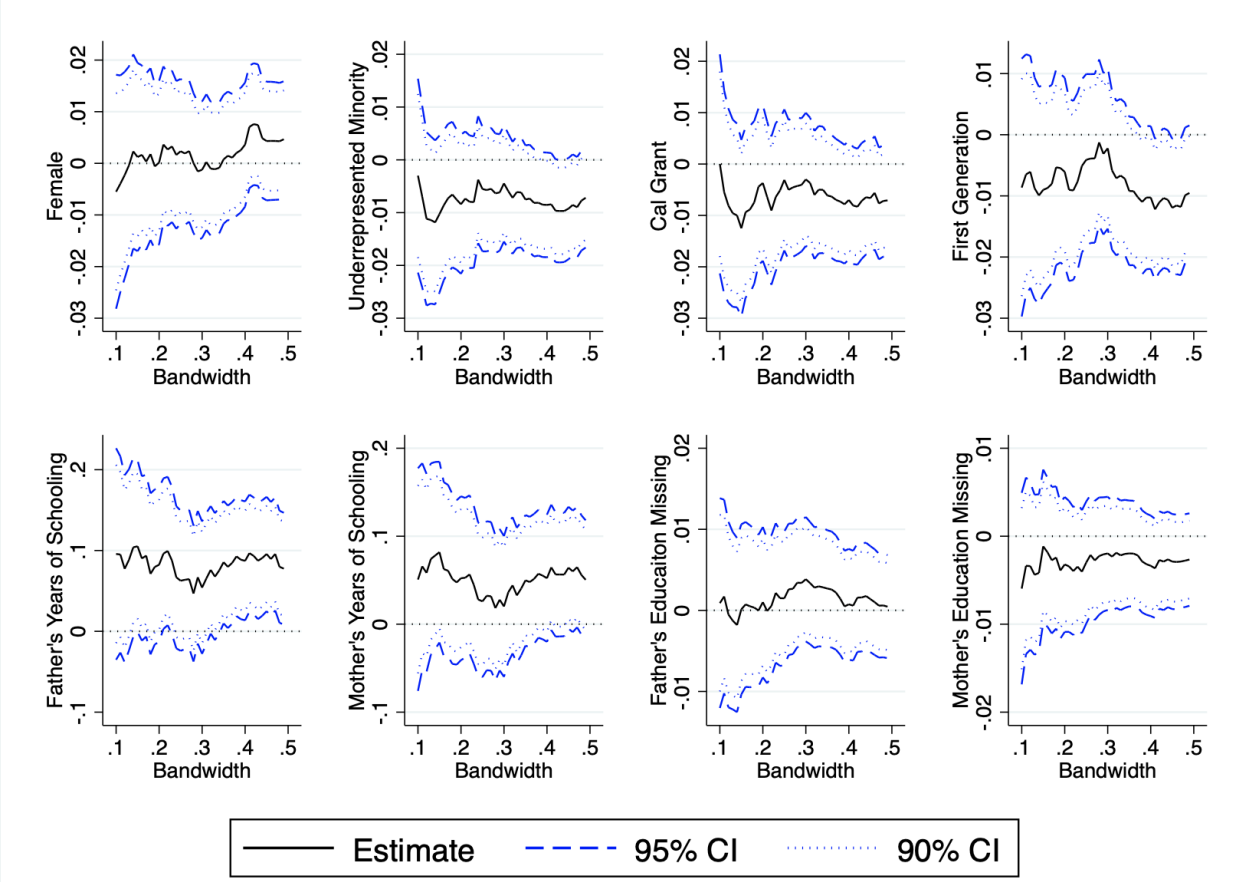

Figure D.9: Covariate Bandwidth Graphs

Note: Each graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the discontinuity at the threshold in a given covariate using a local linear specification at a respective bandwidth.

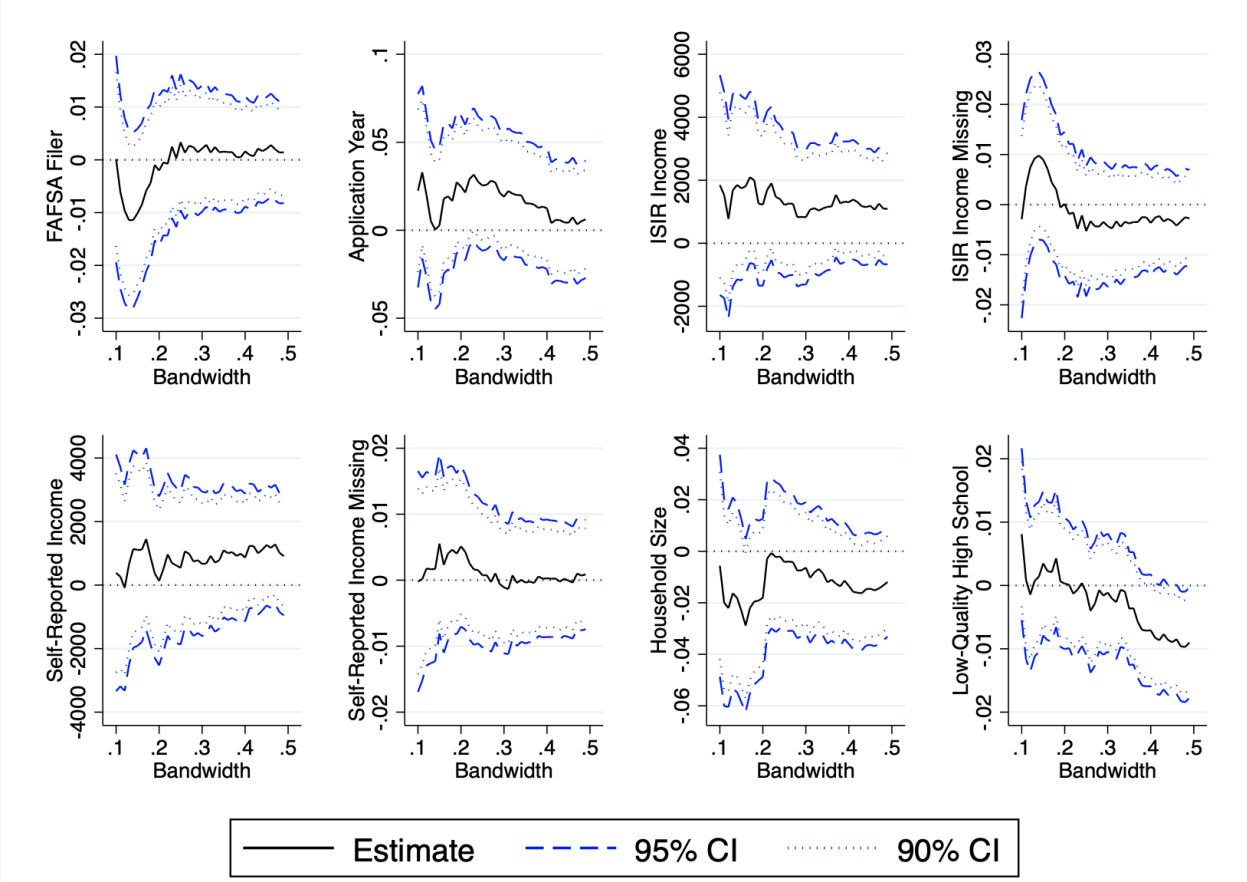

Figure D.10: Covariate Bandwidth Graphs

Note: Each graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the discontinuity at the threshold in a given covariate using a local linear specification at a respective bandwidth.

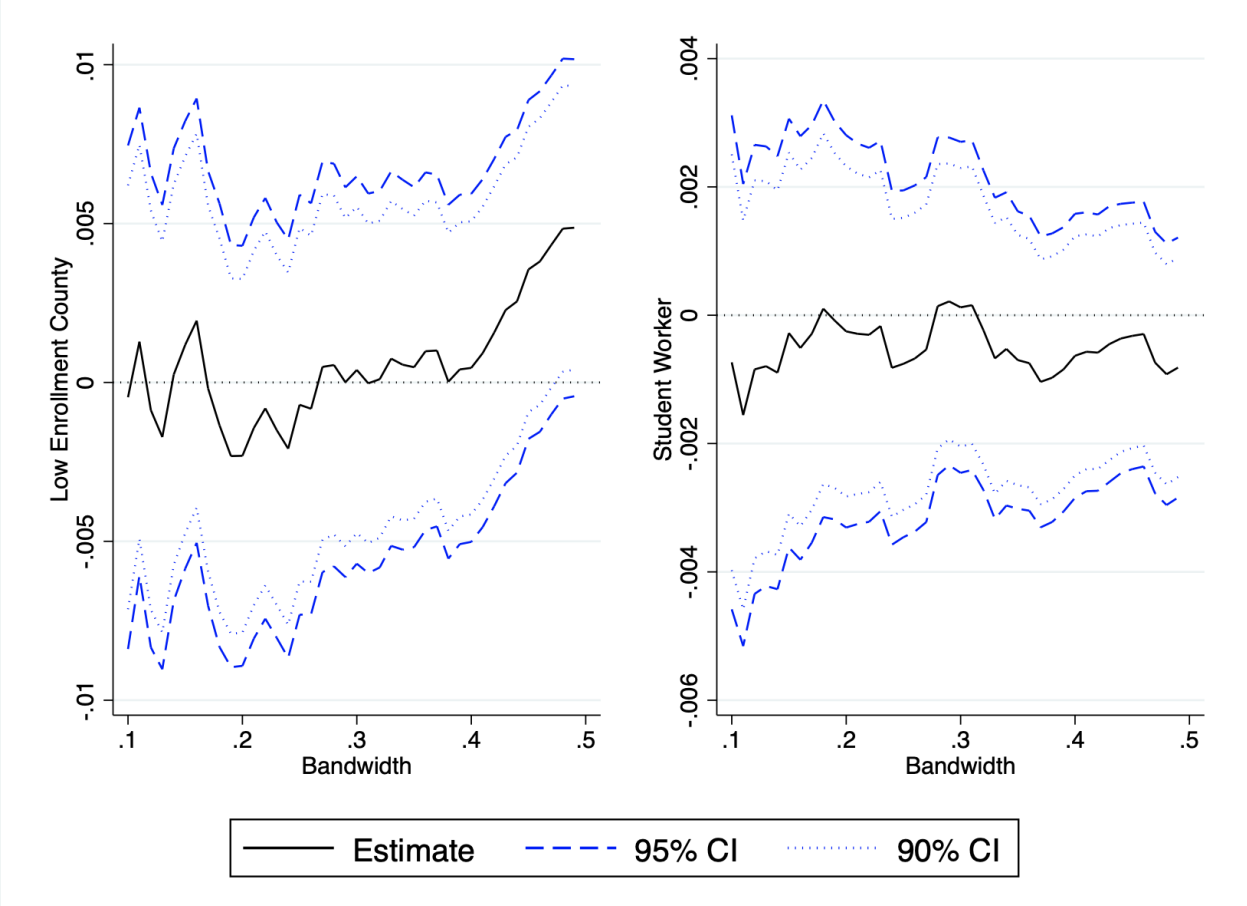

Figure D.11: Covariate Bandwidth Graphs

Note: Each graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the discontinuity at the threshold in a given covariate using a local linear specification at a respective bandwidth.

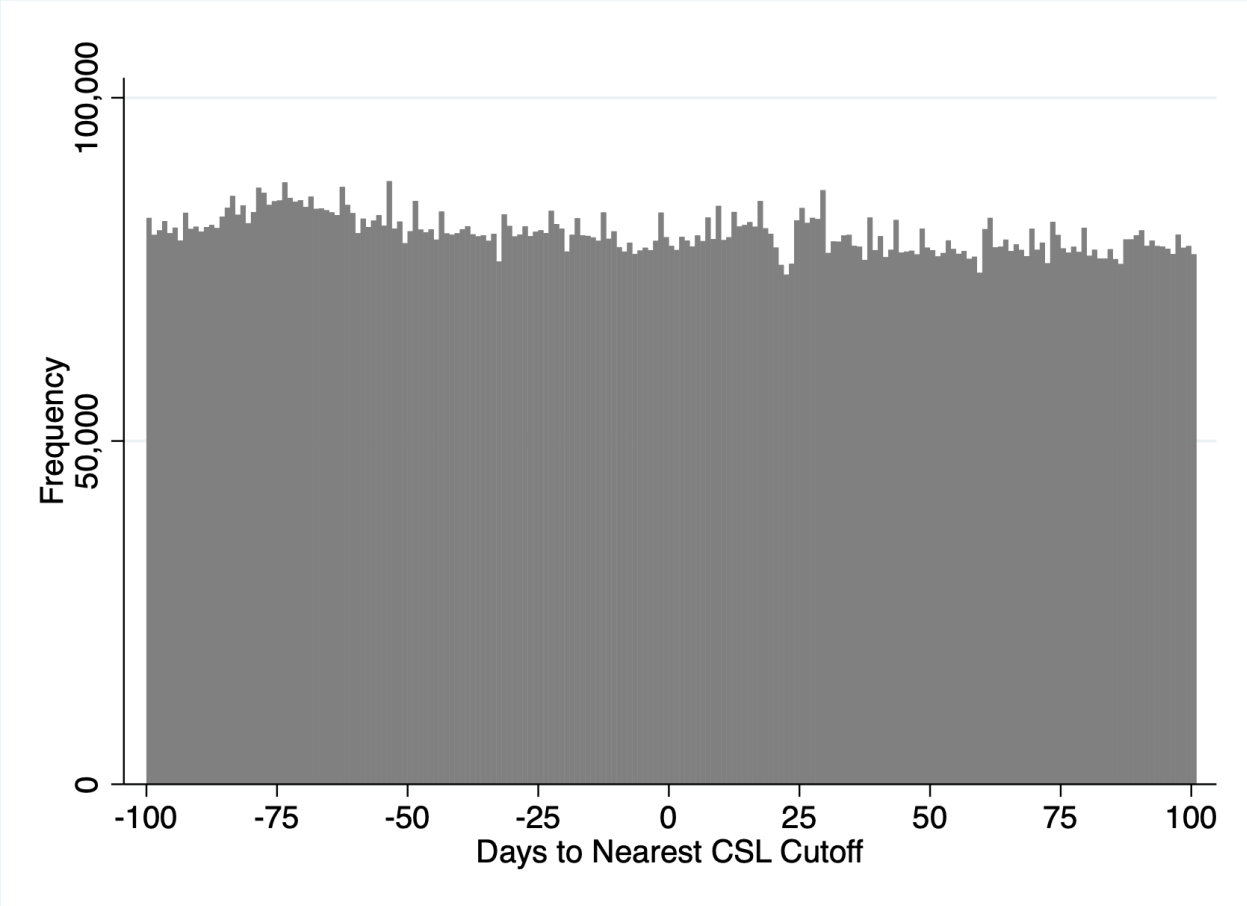

Figure D.12: McCrary Test for Compulsory Schooling Laws

Note: Birthdates are normalized relative to the compulsory schooling law cutoff date for school entry.

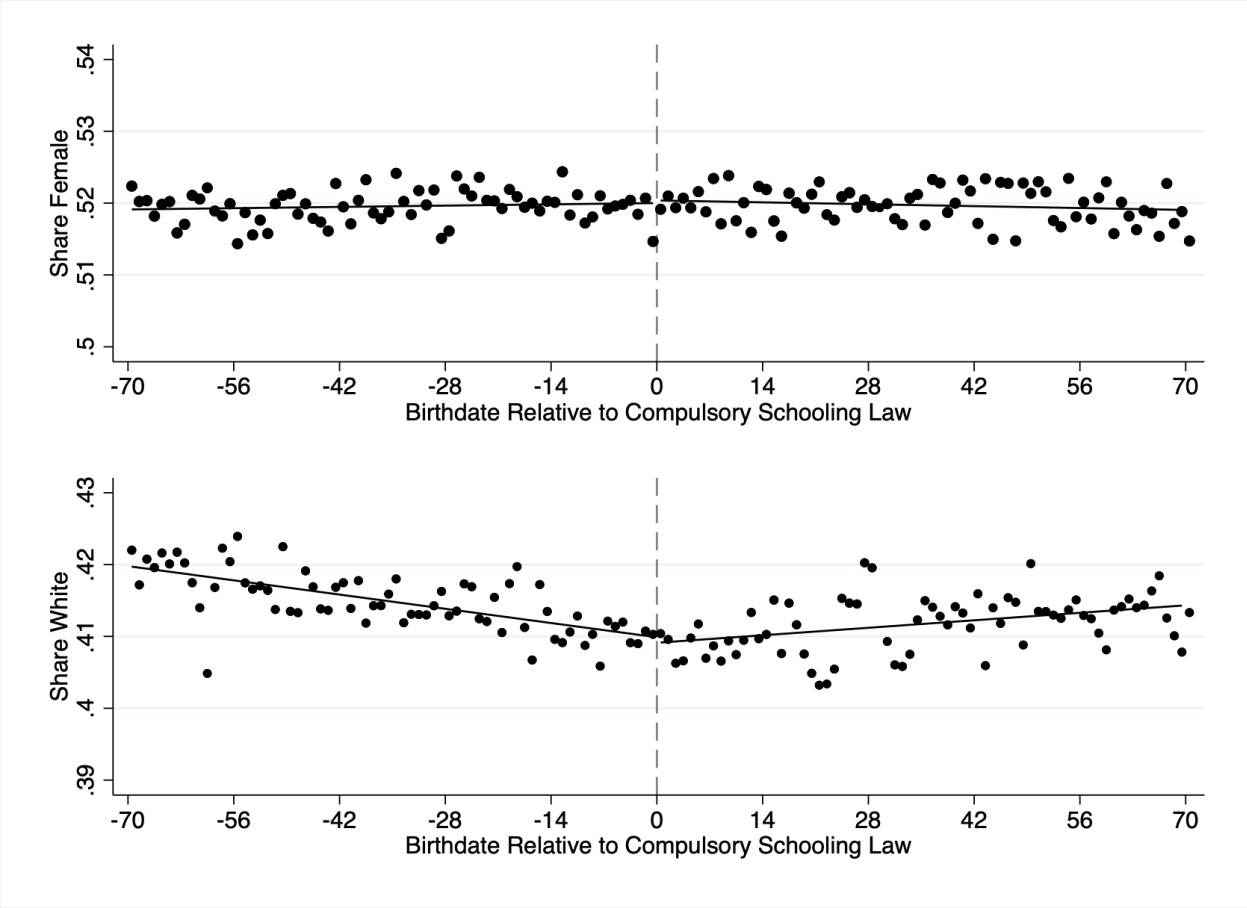

Figure D.13: Demographic Balance Test for Compulsory Schooling Laws

Note: Birthdates are normalized relative to the compulsory schooling law cutoff date for school entry.

E Alternative Admission Policies Appendix

To complement findings from the UC's top percentile policy, I evaluate the impact of enrollment at individual UC campuses using two more admission rules that sharply increase the selectivity of institutions that students attend. The two admission rules vary different margins of treatment – the extensive margin of college enrollment and the intensive margin of college selectivity – and are used to confirm the generalizability of my findings to different campuses and different near-threshold sample populations. The relevant colleges are renamed UC San Andreas and the Elite UC to preserve their anonymity per my data agreement.

Starting with the first-stage for each policy, I plot admission outcomes and enrollment choices against students normalized, reweighted GPAs in Figures E.9 for UC San Andreas and Figure E.10 for the Elite UC. In each case, being just above the GPA threshold for a campus increases the probability a student will be admitted to and enroll at the relevant campus. For UC San Andreas's admission policy, this raises both the extensive margin of 4-year college enrollment and the average selectivity of the college a student attends proxied by the college's applicant rejection rate and the selectivity category from Opportunity Insights. For the Elite UC's admission policy, this raises the total number of UC campuses to which a student was admitted and the intensive margin of campus selectivity, but not the extensive margin of 4-year college enrollment. Students conferred an admission advantage to the Elite UC are more likely to enroll at UC campuses that are at least as selective as the Elite UC, at colleges with higher applicant rejection rates, and at colleges labeled highly selective or elite by Opportunity Insights.

Figures E.11 and E.12 show measures of partisanship plotted against normalized GPA for each admission policy. Enrolling at UC San Andreas reduces the share of students who are registered Republicans and increases the share who are independents with no party affiliation. Enrollment at the Elite UC also reduces the share of students who eventually register as Republicans and increases the share who are independents with no party preference. Notably, attending the Elite UC also reduces the share of students who switch between major parties in either direction, with larger reductions in Democratic to Republican switches. Tables E.3 and E.4 estimate results formally using the same set of outcomes as Table 1, testing for robustness to different specifications that vary the bandwidth used for inference, the inclusion of pre-treatment covariate controls, and the order of a polynomial control for the running variable.

Overall, these two RD designs provide evidence that corroborate my main findings for UC admission from the ELC policy. In each case, UC campuses raise the selectivity of the college at which a student ultimately enrolls, shifting those same students away from

Republican Party registration and toward registration as either independents or Democrats. Effects appear to grow rather than shrink over a longer time window. The fact that these policies generate variation in admission to different campuses illustrates that these effects are not an artifact of a single campus within the UC system. Moreover, the use of near-threshold sample populations from different points of the GPA distribution with alternative less-selective counterfactual enrollment choices underscores the generalizability of my findings to a broader set of college applicants on different treatment margins.

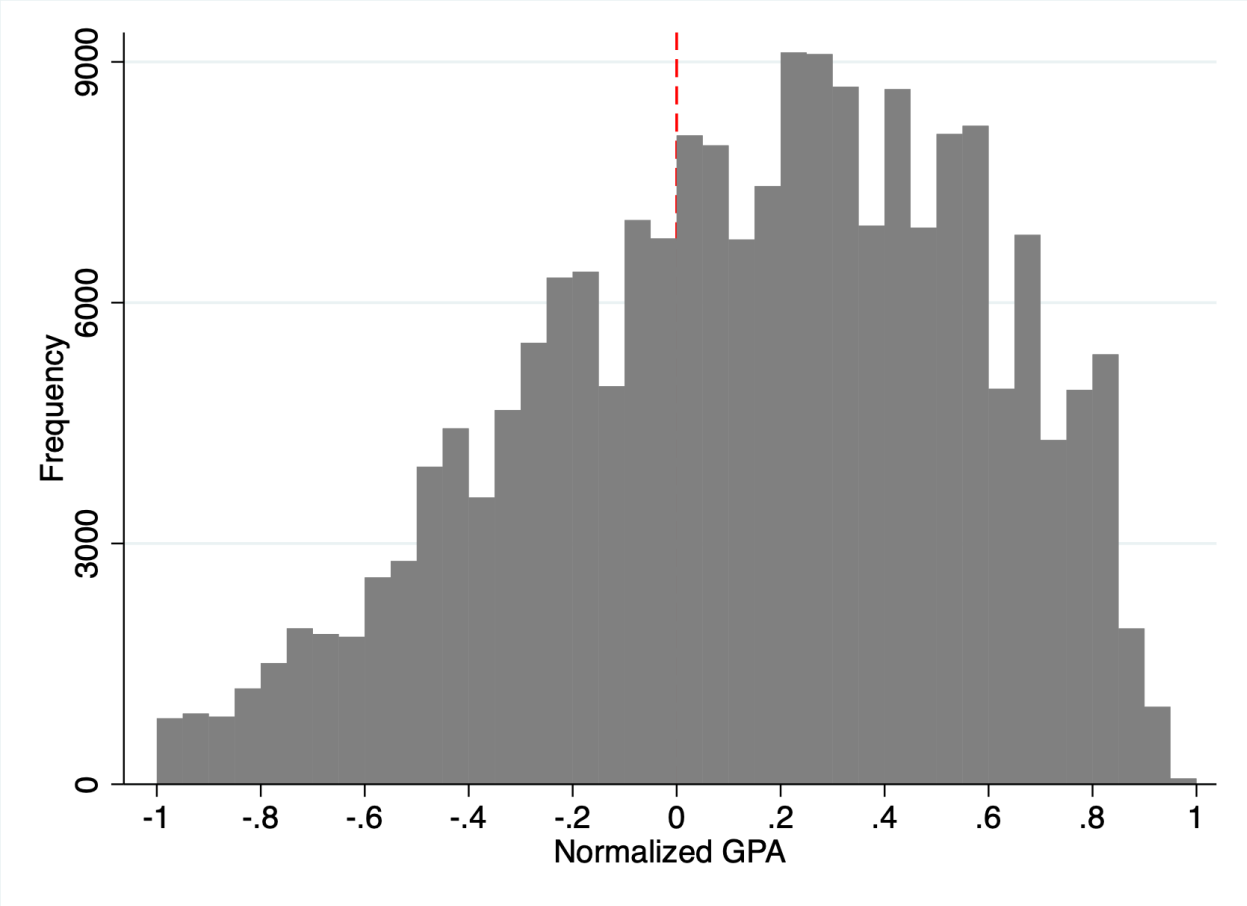

Figure E.1: McCrary Test for UC San Andreas Policy

Note: Reweighted GPA values are normalized to the admission cutoff for UC San Andreas.

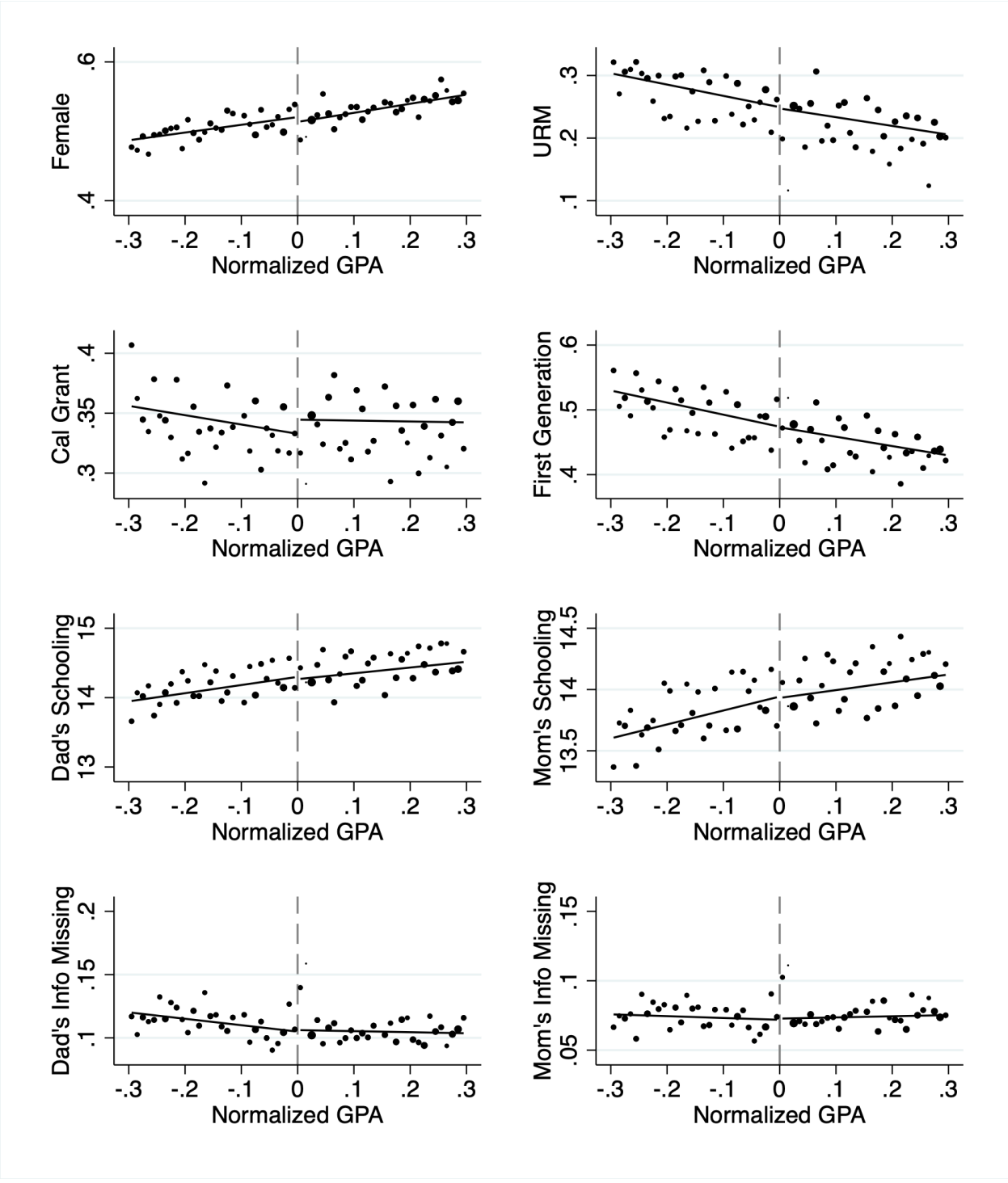

Figure E.2: Covariate RD Graphs for UC San Andreas Policy

Note: Reweighted GPA values are normalized to the admission cutoff for UC San Andreas.

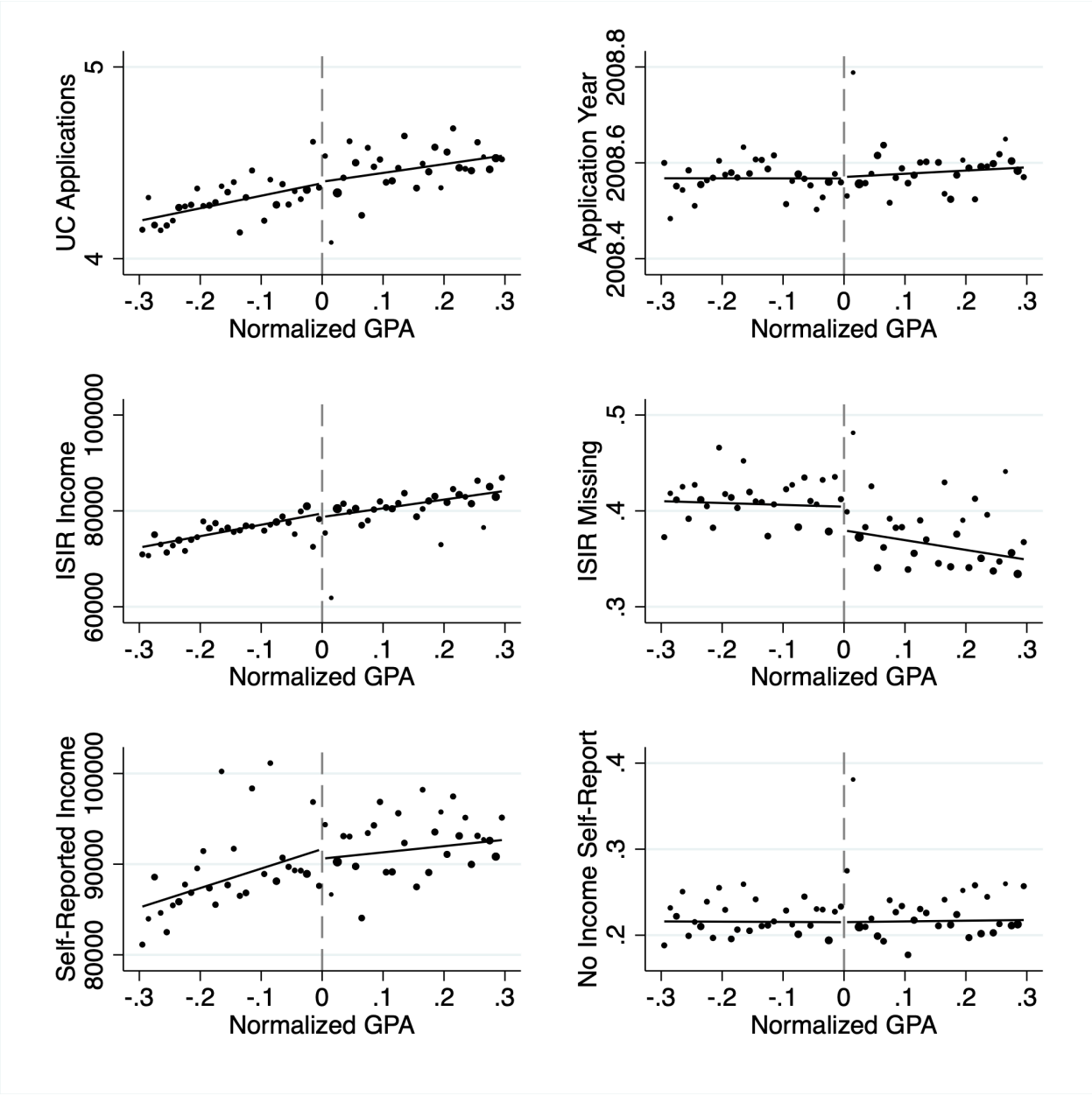

Figure E.3: Covariate RD Graphs for UC San Andreas Policy

Note: Reweighted GPA values are normalized to the admission cutoff for UC San Andreas.

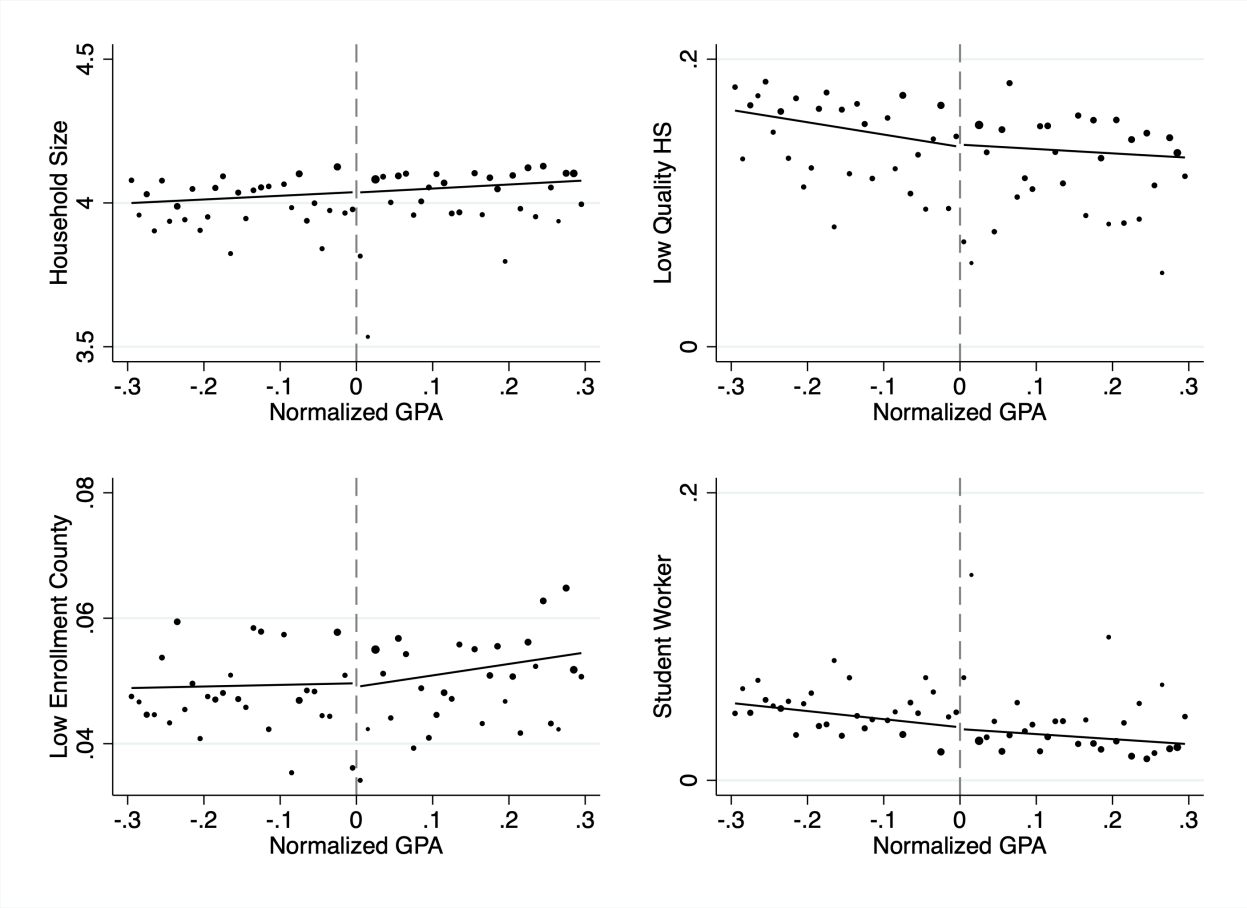

Figure E.4: Covariate RD Graphs for UC San Andreas Policy

Note: Reweighted GPA values are normalized to the admission cutoff for UC San Andreas.

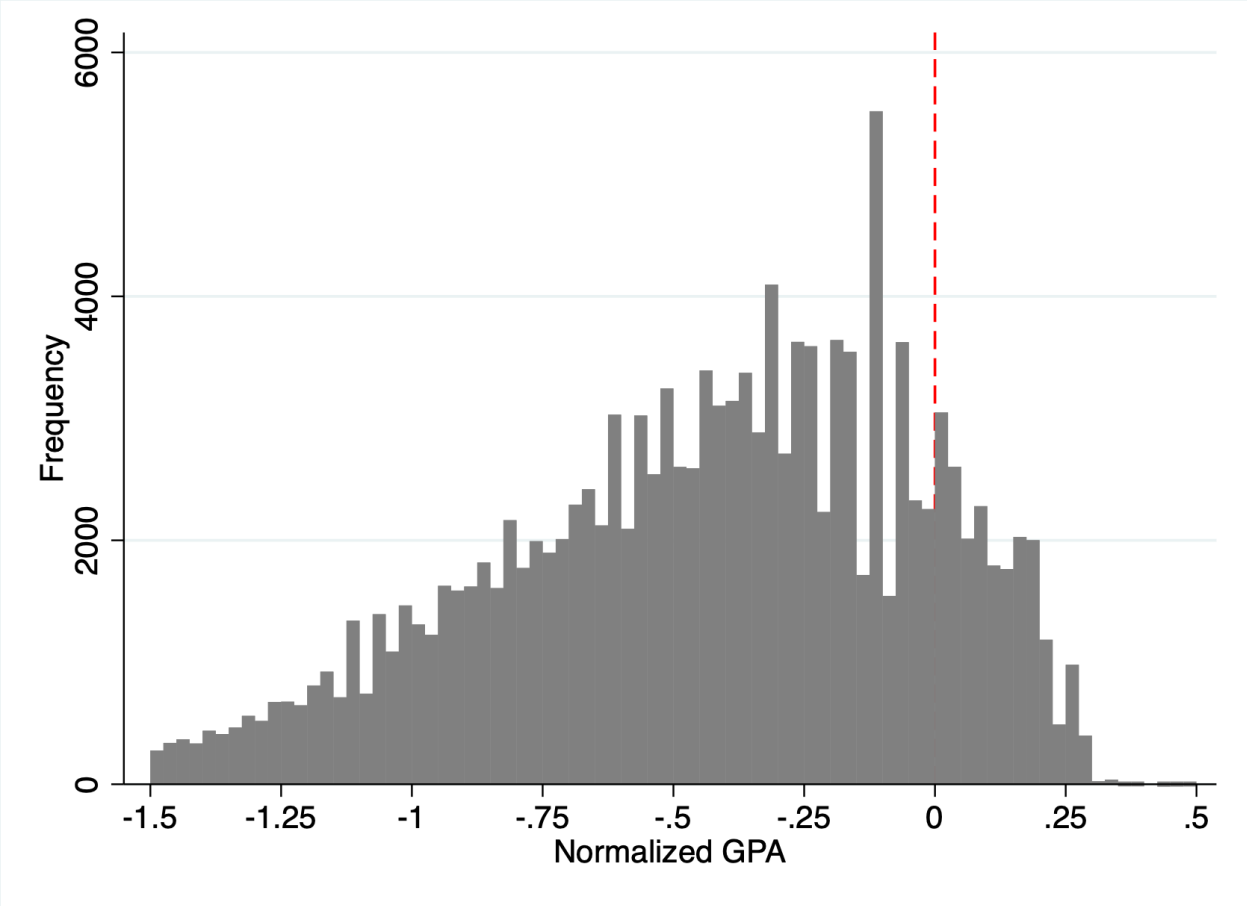

Figure E.5: McCrary Test for Elite UC Policy

Note: Reweighted GPA values are normalized to the admission cutoff for the Elite UC.

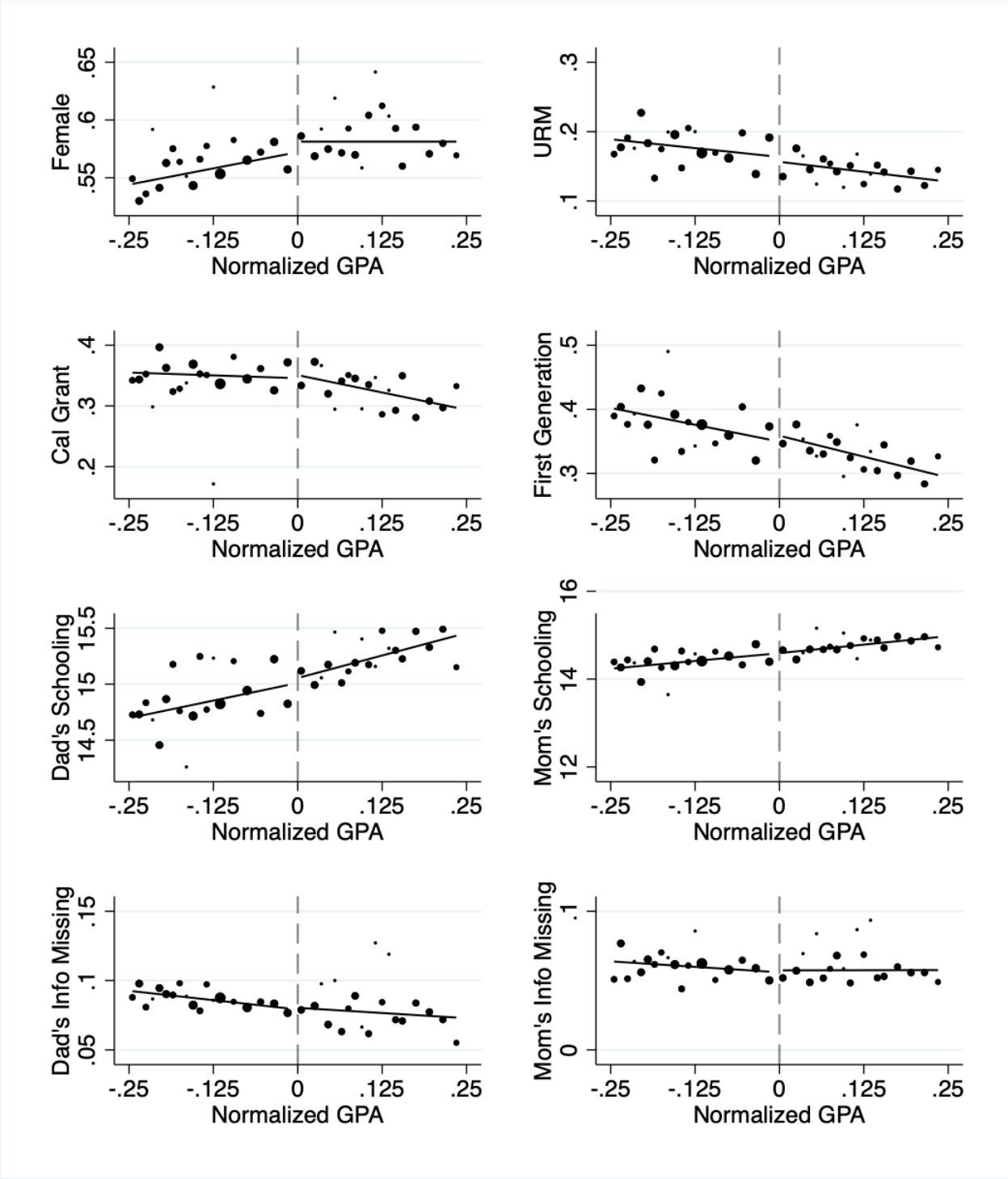

Figure E.6: Covariate RD Graphs for Elite UC Policy

Note: Reweighted GPA values are normalized to the admission cutoff for Elite UC.

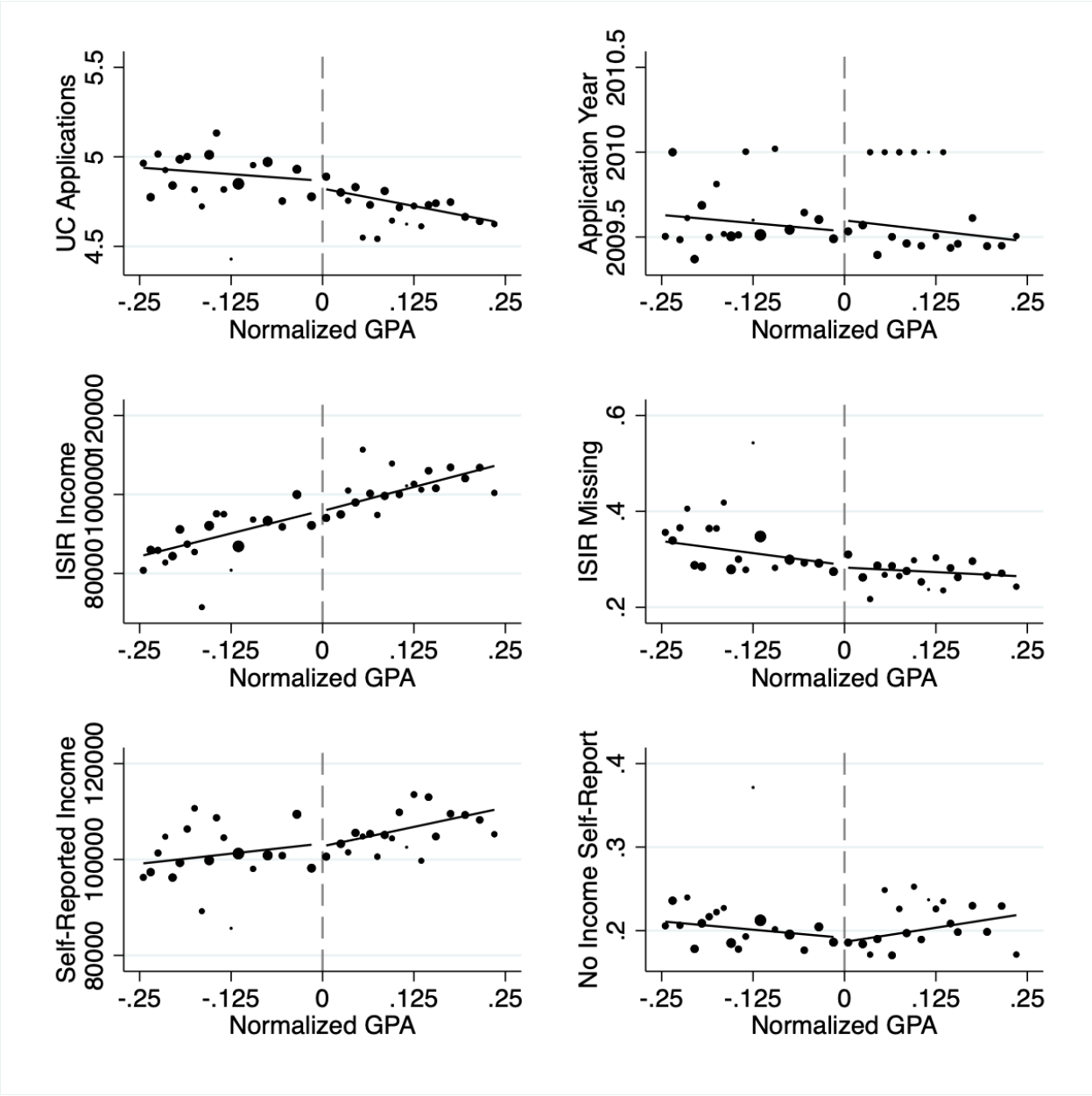

Figure E.7: Covariate RD Graphs for Elite UC Policy

Note: Reweighted GPA values are normalized to the admission cutoff for Elite UC.

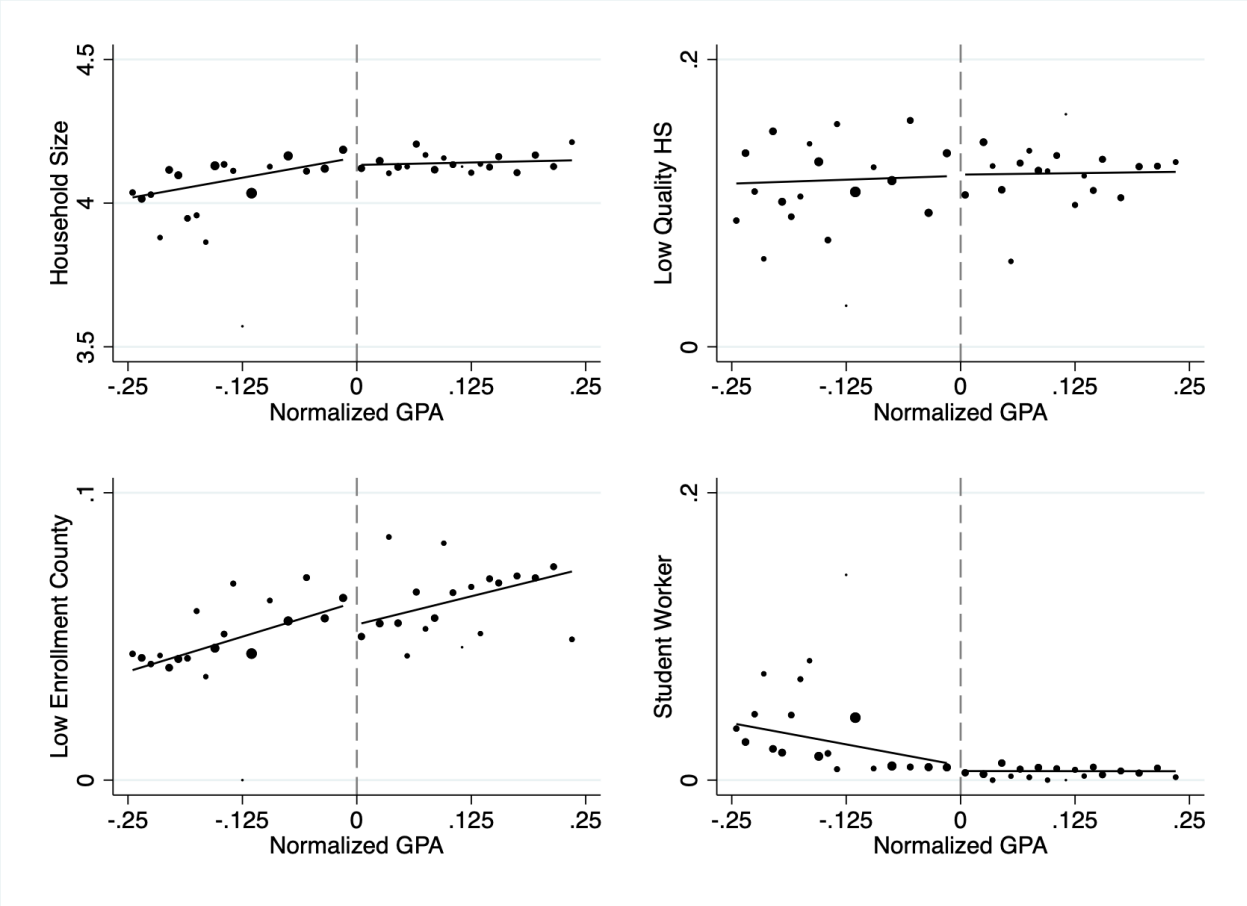

Figure E.8: Covariate RD Graphs for Elite UC Policy

Note: Reweighted GPA values are normalized to the admission cutoff for Elite UC.

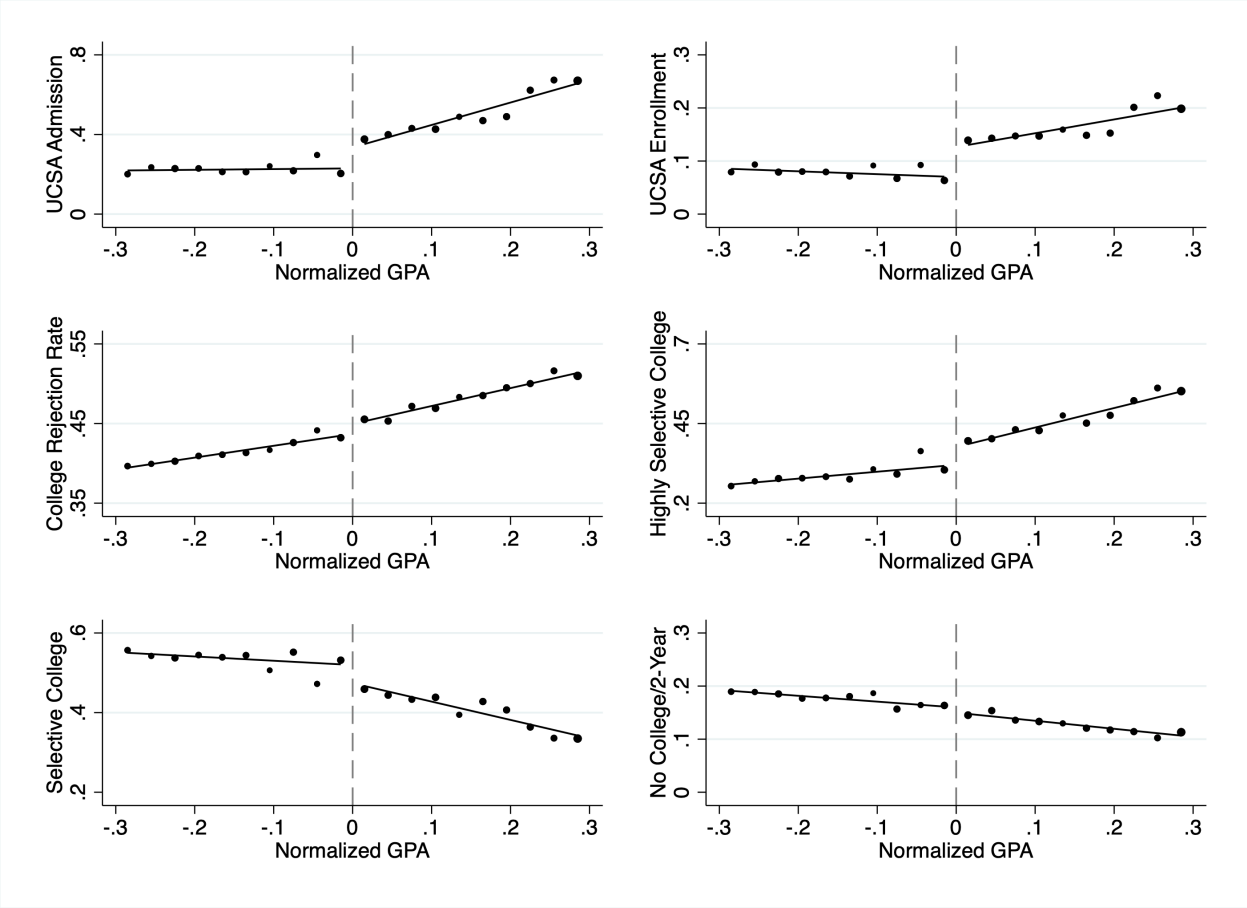

Figure E.9: RD Graphs of College Enrollment for UC San Andreas Policy

Note: Reweighted GPA values are normalized to the admission cutoff for UC San Andreas. The outcome UCSA Admission refers to a binary variable for admission to UC San Andreas. UCSA enrollment likewise refers to enrollment at UC San Andreas. College Rejection Rate refers to the proportion of college applicants that were rejected by the college at which a student enrolled, with this outcome set to zero of open access institutions and non-enrollment in college. Highly Selective College refers to a binary indicator for enrollment at any 4-year college rated as Highly Selective, Elite or Ivy-Plus by Opportunity Insights. Selective College refers to a binary indicator for enrollment at any 4-year college rated as Selective or Non-selective by Opportunity Insights. No College/2-Year refers to enrollment at 2-year colleges, community colleges, online colleges, or not having any college enrollment record.

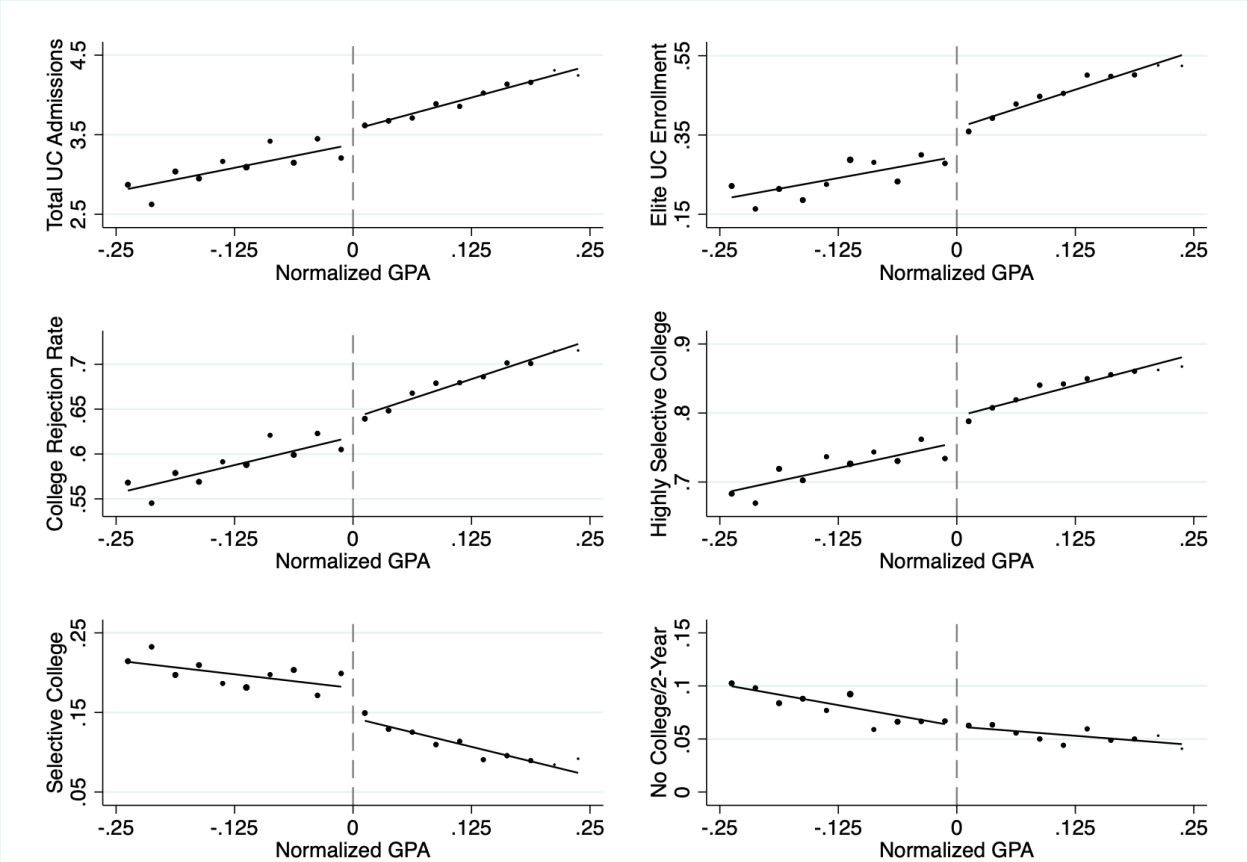

Figure E.10: RD Graphs of College Enrollment for Elite UC Policy

Note: Reweighted GPA values are normalized to the admission cutoff for Elite UC. The outcome Total UC Admissions refers to the total number of UC campuses to which a student was admitted. Elite UC enrollment refers to a binary indicator for enrollment at the Elite UC campus or any campus that is more selective than the Elite UC. College Rejection Rate refers to the proportion of college applicants that were rejected by the college at which a student enrolled, with this outcome set to zero of open access institutions and non-enrollment in college. Highly Selective College refers to a binary indicator for enrollment at any 4-year college rated as Highly Selective, Elite or Ivy-Plus by Opportunity Insights. Selective College refers to a binary indicator for enrollment at any 4-year college rated as Selective or Non-selective by Opportunity Insights. No College/2-Year refers to enrollment at 2-year colleges, community colleges, online colleges, or not having any college enrollment record.

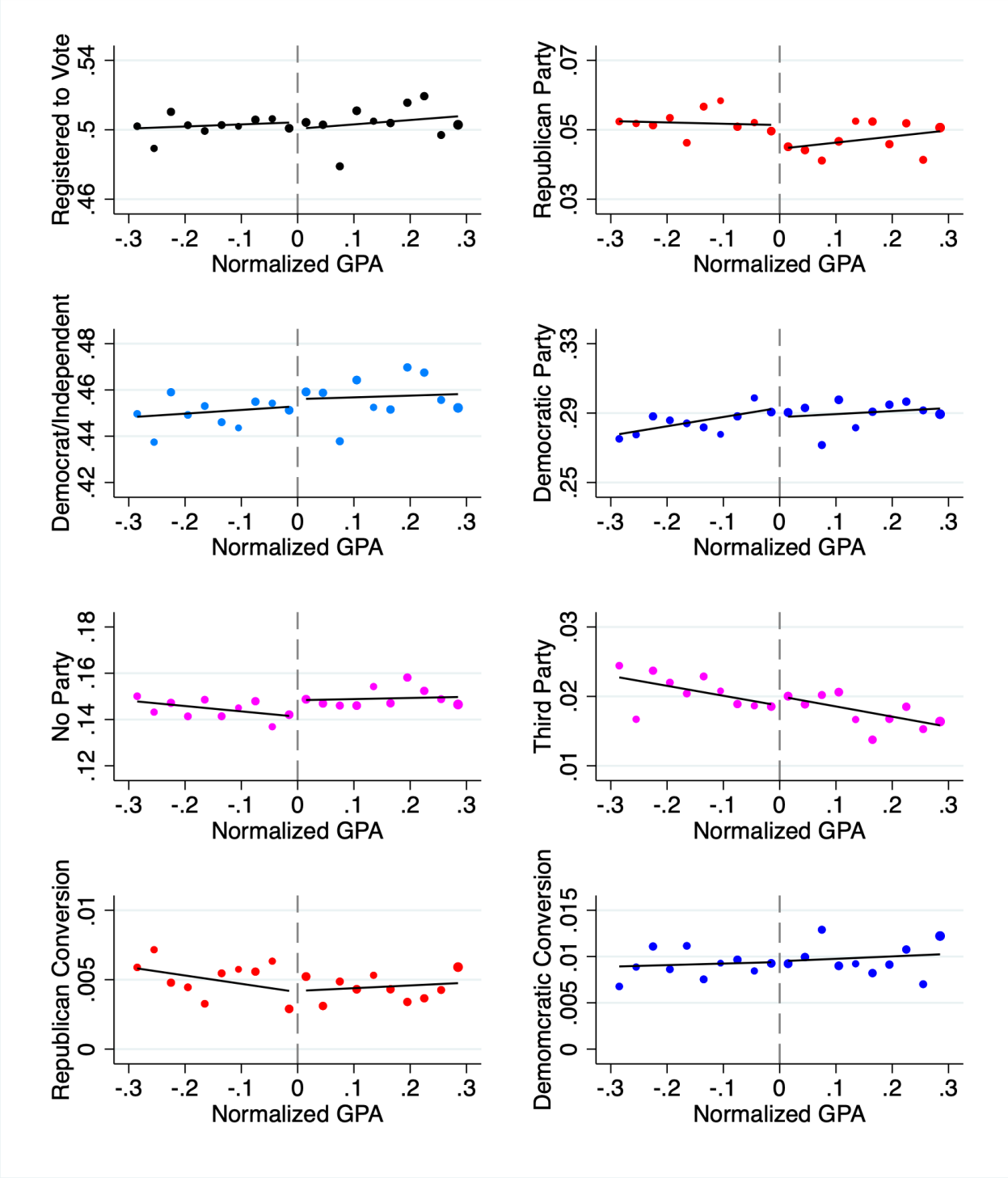

Figure E.11: RD Graphs of Voter Registration Outcomes for UC San Andreas Policy

Note: Reweighted GPA values are normalized to the admission cutoff for UC San Andreas. Outcomes correspond directly to those in Table E.3.

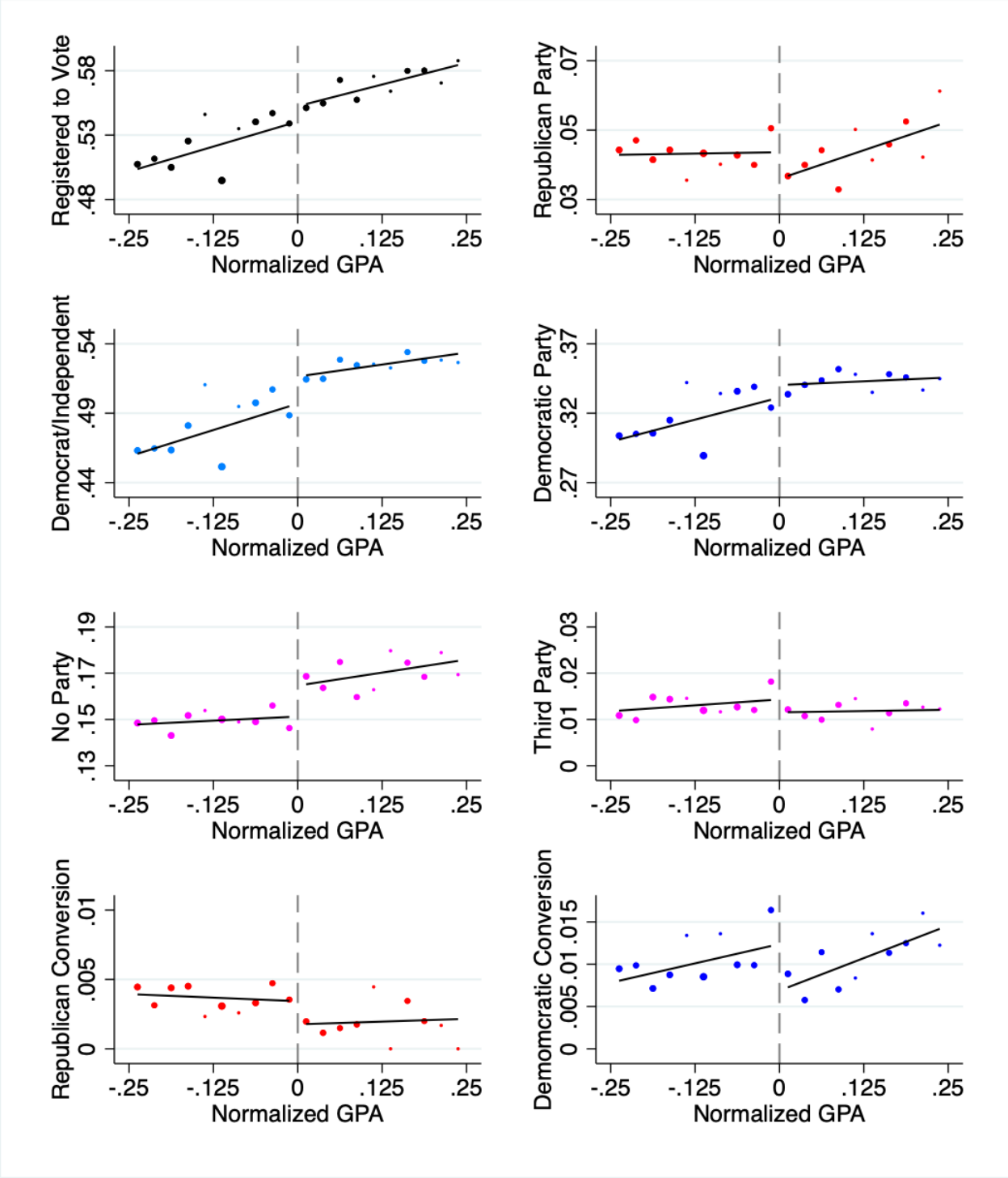

Figure E.12: RD Graphs of Voter Registration Outcomes for Elite UC Policy

Note: Reweighted GPA values are normalized to the admission cutoff for Elite UC. Outcomes correspond directly to those in Table E.4.

Table E.1: Balance Checks for Predicted Partisanship for UC San Andreas Policy

Outcome	(1)	(2)	(3)
Predicted Republican	-0.0003	-0.0004	0.0002
	(0.0004)	(0.0004)	(0.0006)
Predicted Dem/Ind	0.0003	0.0004	-0.0002
	(0.0004)	(0.0004)	(0.0006)
Predicted Democrat	0.0001	0.0001	-0.0011
	(0.0012)	(0.0011)	(0.0017)
Predicted No Party	0.0004	0.0004	0.0010
	(0.0008)	(0.0008)	(0.0012)
Predicted Third Party	-0.0001	-0.0001	-0.0001
	(0.0001)	(0.0001)	(0.0002)
Bandwidth	Optimal	0.3	0.3
Polynomial	1	1	2
Sample Size	Varies	85,482	85,482

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Results correspond to those in Table D.1, but with outcomes that are predicted using partisanship conditional on registration rather than unconditional partisanship. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

Table E.2: Balance Checks for Predicted Partisanship for Elite UC Policy

Outcome	(1)	(2)	(3)
Predicted Republican	-0.0002	-0.0005	-0.0010
	(0.0006)	(0.0005)	(0.0007)
Predicted Dem/Ind	0.0002	0.0005	0.0010
	(0.0006)	(0.0005)	(0.0007)
Predicted Democrat	-0.0025	0.0006	-0.0002
	(0.0019)	(0.0014)	(0.0021)
Predicted No Party	0.0022^{+}	-0.0000	0.0014
	(0.0013)	(0.0009)	(0.0014)
Predicted Third Party	0.0001	-0.0001	-0.0003
	(0.0002)	(0.0001)	(0.0002)
Bandwidth	Optimal	0.3	0.3
Polynomial	1	1	2
Sample Size	Varies	42,108	42,108

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Results correspond to those in Table D.1, but with outcomes that are predicted using partisanship conditional on registration rather than unconditional partisanship. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). Predicted outcomes are generated using the pre-treatment covariates listed in Section 2.

Table E.3: Effects of UC San Andreas Admission Policy on Registration and Partisanship

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Total Voter Registration Rate						
Registered to Vote	-0.0073	-0.0093	-0.0040	-0.0063	-0.0068	-0.0084
	(0.0078)	(0.0077)	(0.0069)	(0.0068)	(0.0106)	(0.0105)
B. Political Party Membership						
Republican Party	-0.0073*	-0.0073*	-0.0069*	-0.0069*	-0.0074	-0.0077+
	(0.0029)	(0.0028)	(0.0030)	(0.0029)	(0.0045)	(0.0045)
Democrat/Independent	0.0010	-0.0010	0.0030	0.0007	0.0007	-0.0007
	(0.0078)	(0.0077)	(0.0068)	(0.0068)	(0.0105)	(0.0105)
Democratic Party	-0.0032	-0.0044	-0.0058	-0.0072	-0.0045	-0.0050
	(0.0075)	(0.0074)	(0.0062)	(0.0062)	(0.0096)	(0.0095)
No Party Preference	0.0106*	0.0096*	0.0072	0.0064	0.0022	0.0013
	(0.0044)	(0.0044)	(0.0048)	(0.0048)	(0.0075)	(0.0074)
Third Party	0.0010	0.0010	0.0015	0.0015	0.0030	0.0029
	(0.0020)	(0.0020)	(0.0019)	(0.0019)	(0.0029)	(0.0029)
C. Early Life Conversion between Major Parties						
Republican Convert	-0.0002	-0.0002	0.0001	0.0001	0.0009	0.0010
	(0.0009)	(0.0009)	(0.0009)	(0.0009)	(0.0014)	(0.0014)
Democratic Convert	0.0009	0.0009	0.0001	-0.0000	0.0025	0.0024
	(0.0014)	(0.0014)	(0.0013)	(0.0013)	(0.0020)	(0.0020)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	85,482	85,482	85,482	85,482

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Democrat/Independent" refers to the fraction of students who are registered as Democrat, as a no party preference voter, or as a member of a third party. Democratic and Republican converts are voters who are currently registered with the Democratic and Republican Party in California, but at any time in the past were a registered member of the other major party.

Table E.4: Effects of Elite UC Admission Policy on Registration and Partisanship

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Total Voter Registration Rate						
Registered to Vote	-0.0017	-0.0007	0.0116	0.0089	-0.0007	0.0026
	(0.0119)	(0.0119)	(0.0090)	(0.0090)	(0.0137)	(0.0136)
B. Political Party Membership						
Republican Party	-0.0082^{+}	-0.0076^{+}	-0.0077*	-0.0073*	-0.0112*	-0.0101^{+}
	(0.0043)	(0.0043)	(0.0036)	(0.0036)	(0.0055)	(0.0055)
Democrat/Independent	0.0073	0.0072	0.0193*	0.0162^{+}	0.0104	0.0127
	(0.0119)	(0.0119)	(0.0091)	(0.0091)	(0.0137)	(0.0137)
Democratic Party	-0.0065	-0.0065	0.0088	0.0063	-0.0060	-0.0041
	(0.0113)	(0.0112)	(0.0085)	(0.0085)	(0.0129)	(0.0129)
No Party Preference	0.0174*	0.0170*	0.0133*	0.0125^{+}	0.0191^{+}	0.0194^{+}
	(0.0087)	(0.0087)	(0.0066)	(0.0066)	(0.0100)	(0.0100)
Third Party	-0.0036	-0.0034	-0.0028	-0.0027	-0.0027	-0.0025
	(0.0028)	(0.0028)	(0.0020)	(0.0020)	(0.0032)	(0.0032)
C. Early Life Conversion between Major Parties						
Republican Convert	-0.0026*	-0.0026*	-0.0017^{+}	-0.0017^{+}	-0.0031*	-0.0031*
	(0.0012)	(0.0012)	(0.0009)	(0.0009)	(0.0015)	(0.0015)
Democratic Convert	-0.0069**	-0.0063*	-0.0056**	-0.0058**	-0.0066*	-0.0061*
	(0.0026)	(0.0026)	(0.0018)	(0.0018)	(0.0029)	(0.0029)
Bandwidth	Optimal	Optimal	0.25	0.25	0.25	0.25
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	49,201	49,201	49,201	49,201

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Democrat/Independent" refers to the fraction of students who are registered as Democrat, as a no party preference voter, or as a member of a third party. Democratic and Republican converts are voters who are currently registered with the Democratic and Republican Party in California, but at any time in the past were a registered member of the other major party.

F Robustness Test and External Validity Appendix

Table F.1: Effects of UC Top Percent Policy on Partisanship (Conditional on Registration)

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Political Party Membership						
Republican Party	-0.0146*	-0.0142*	-0.0133**	-0.0130*	-0.0187*	-0.0181*
	(0.0059)	(0.0059)	(0.0052)	(0.0051)	(0.0078)	(0.0077)
Democrat/Independent	0.0146*	0.0142*	0.0133**	0.0130*	0.0187*	0.0181*
	(0.0059)	(0.0059)	(0.0052)	(0.0051)	(0.0078)	(0.0077)
Democratic Party	0.0055	0.0057	0.0042	0.0042	0.0032	0.0023
	(0.0089)	(0.0088)	(0.0089)	(0.0088)	(0.0135)	(0.0133)
No Party Preference	0.0125	0.0122	0.0143+	0.0138+	0.0186	0.0187
	(0.0089)	(0.0089)	(0.0083)	(0.0083)	(0.0126)	(0.0126)
Third Party	-0.0055+	-0.0053+	-0.0051+	-0.0051+	-0.0031	-0.0028
	(0.0029)	(0.0029)	(0.0029)	(0.0029)	(0.0044)	(0.0044)
B. Early Life Conversion between Major Parties						
Republican Convert	-0.0048**	-0.0048**	-0.0028+	-0.0028+	-0.0051*	-0.0050*
	(0.0018)	(0.0018)	(0.0015)	(0.0015)	(0.0023)	(0.0023)
Democratic Convert	-0.0026	-0.0026	-0.0028	-0.0028	-0.0029	-0.0030
	(0.0026)	(0.0026)	(0.0025)	(0.0025)	(0.0037)	(0.0037)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	42,108	42,108	42,108	42,108

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Democrat/Independent" refers to the fraction of students who are registered as Democrat, as a no party preference voter, or as a member of a third party. Democratic and Republican converts are voters who are currently registered with the Democratic and Republican Party in California, but at any time in the past were a registered member of the other major party.

Table F.2: Effects of the UC Top Percent Policy on Voter Registration Outcomes

Outcome	(1)	(2)	(3)	(4)
A. Total Voter Registration Rate				
Registered to Vote	0.0126^{+}	0.0124^{+}	0.0135	0.0138
	(0.0070)	(0.0070)	(0.0103)	(0.0102)
B. Political Party Membership				
Republican Party	-0.0063*	-0.0063*	-0.0084^{+}	-0.0083^{+}
	(0.0029)	(0.0029)	(0.0043)	(0.0043)
Democrat/Independent	0.0189**	0.0188**	0.0219*	0.0220*
	(0.0070)	(0.0070)	(0.0103)	(0.0103)
Democratic Party	0.0099	0.0098	0.0118	0.0116
	(0.0064)	(0.0064)	(0.0094)	(0.0094)
No Party Preference	0.0108*	0.0108*	0.0108	0.0112
	(0.0050)	(0.0050)	(0.0077)	(0.0076)
Third Party	-0.0018	-0.0018	-0.0007	-0.0008
	(0.0016)	(0.0016)	(0.0024)	(0.0024)
C. Early Life Conversion between Major Parties				
Republican Convert	-0.0010	-0.0011	-0.0023^{+}	-0.0024*
	(0.0008)	(0.0008)	(0.0012)	(0.0012)
Democratic Convert	-0.0015	-0.0015	-0.0016	-0.0015
	(0.0014)	(0.0014)	(0.0020)	(0.0020)
Bandwidth	0.3	0.3	0.3	0.3
Polynomial	1	1	2	2
Controls	No	Yes	No	Yes
HS-Year FEs	Yes	Yes	Yes	Yes
Sample Size	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). These outcomes correspond to those in Table 1.

Table F.3: Effects on Party Registration with Bias-Corrected CIs

Outcome	(1)	(2)
Voter Registration		
RD_Estimate	0.0140 (0.0085)	0.0131 (0.0084)
Robust 95% CI	[-.003 ; .036]	[-.004 ; .035]
Robust p-value	0.096	0.113
Republican Party		
RD_Estimate	-0.0082* (0.0034)	-0.0083* (0.0034)
Robust 95% CI	[-.017 ; -.001]	[-.017 ; -.002]
Robust p-value	0.021	0.018
Democrat/Independent		
RD_Estimate	0.0228** (0.0088)	0.0222* (0.0087)
Robust 95% CI	[.005 ; .046]	[.005 ; .045]
Robust p-value	0.015	0.017
Democratic Party		
RD_Estimate	0.0099 (0.0077)	0.0098 (0.0077)
Robust 95% CI	[-.006 ; .03]	[-.006 ; .03]
Robust p-value	0.180	0.181
No Party Preference		
RD_Estimate	0.0138* (0.0061)	0.0133* (0.0061)
Robust 95% CI	[.002 ; .03]	[.001 ; .029]
Robust p-value	0.027	0.032
Third Party		
RD_Estimate	-0.0019 (0.0017)	-0.0019 (0.0018)
Robust 95% CI	[-.006 ; .002]	[-.006 ; .002]
Robust p-value	0.369	0.385
Republican Convert		
RD_Estimate	-0.0026* (0.0011)	-0.0026* (0.0011)
Robust 95% CI	[-.005 ; -.001]	[-.005 ; -.001]
Robust p-value	0.015	0.017
Democratic Convert		
RD_Estimate	-0.0013 (0.0015)	-0.0014 (0.0015)
Robust 95% CI	[-.005 ; .002]	[-.005 ; .002]
Robust p-value	0.432	0.404
Bandwidth	MSE-Optimal	MSE-Optimal
Polynomial	1	1
Covariates	No	Yes

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Each row titled "RD Estimate" shows the conventional point estimate and standard errors in parentheses for a given outcome variable. These are calculated using a triangular kernel at the MSE-optimal bandwidth derived from Calonico et al. (2020). The rows "Robust 95% CI" and "Robust p-value" show the bias-corrected confidence interval and the bias-corrected p-value for the same outcome variable (Calonico et al., 2014). These outcomes correspond to those in Table 1.

Table F.4: Effects on Party Registration with Honest CIs

Outcome	(1)	(2)
Voter Registration		
RD Estimate	0.0140 (0.0075)	0.0139 (0.0074)
Robust 95% CI	[-.0022; .0284]	[-.0025; .0303]
Robust 90% CI	[.0001; .0278]	[.0000; .0277]
Republican Party		
RD Estimate	-0.0076 (0.0032)	-0.0072 (0.0030)
Robust 95% CI	[-.0147; -.0005]	[-.0143; -.0001]
Robust 90% CI	[-.0136; -.0017]	[-.0132; -.0011]
Democrat/Independent		
RD Estimate	0.0213 (0.0076)	0.0211 (0.0074)
Robust 95% CI	[.0049; .0378]	[.0046; .0376]
Robust 90% CI	[.0074; .0352]	[.0072; .0350]
Democratic Party		
RD Estimate	0.0101 (0.0068)	0.0105 (0.0069)
Robust 95% CI	[-.0044; .0248]	[-.0042; .0252]
Robust 90% CI	[-.0021; .0225]	[-.0019; .0229]
No Party Preference		
RD Estimate	0.0128 (0.0055)	0.0126 (0.0054)
Robust 95% CI	[.0009; .0247]	[.0008; .0245]
Robust 90% CI	[.0028; .0228]	[.0026; .0226]
Third Party		
RD Estimate	-0.0013 (0.0020)	-0.0020 (0.0017)
Robust 95% CI	[-.0058; .0032]	[-.0069; .0029]
Robust 90% CI	[-.0051; .0025]	[-.0062; .0023]
Republican Convert		
RD Estimate	-0.0024 (0.0010)	-0.0019 (0.0009)
Robust 95% CI	[-.0046; -.0002]	[-.0042; .0004]
Robust 90% CI	[-.0043; -.0006]	[-.0038; .0000]
Democratic Convert		
RD Estimate	-0.0012 (0.0015)	-0.0013 (0.0015)
Robust 95% CI	[-.0046; .0020]	[-.0046; .0020]
Robust 90% CI	[-.0040; .0015]	[-.0040; .0015]
Bandwidth	MSE-Optimal	0.3
Polynomial	1	1

Note: Each row titled "RD Estimate" shows the point estimate and standard errors in parentheses for a given outcome variable using a triangular kernel and the bounded seconded derivative method (Kolesar and Rothe, 2018). The rows "Robust 95% CI" and "Robust 90% CI" show the honest confidence intervals for the same outcome variable. These outcomes correspond to those in Table 1.

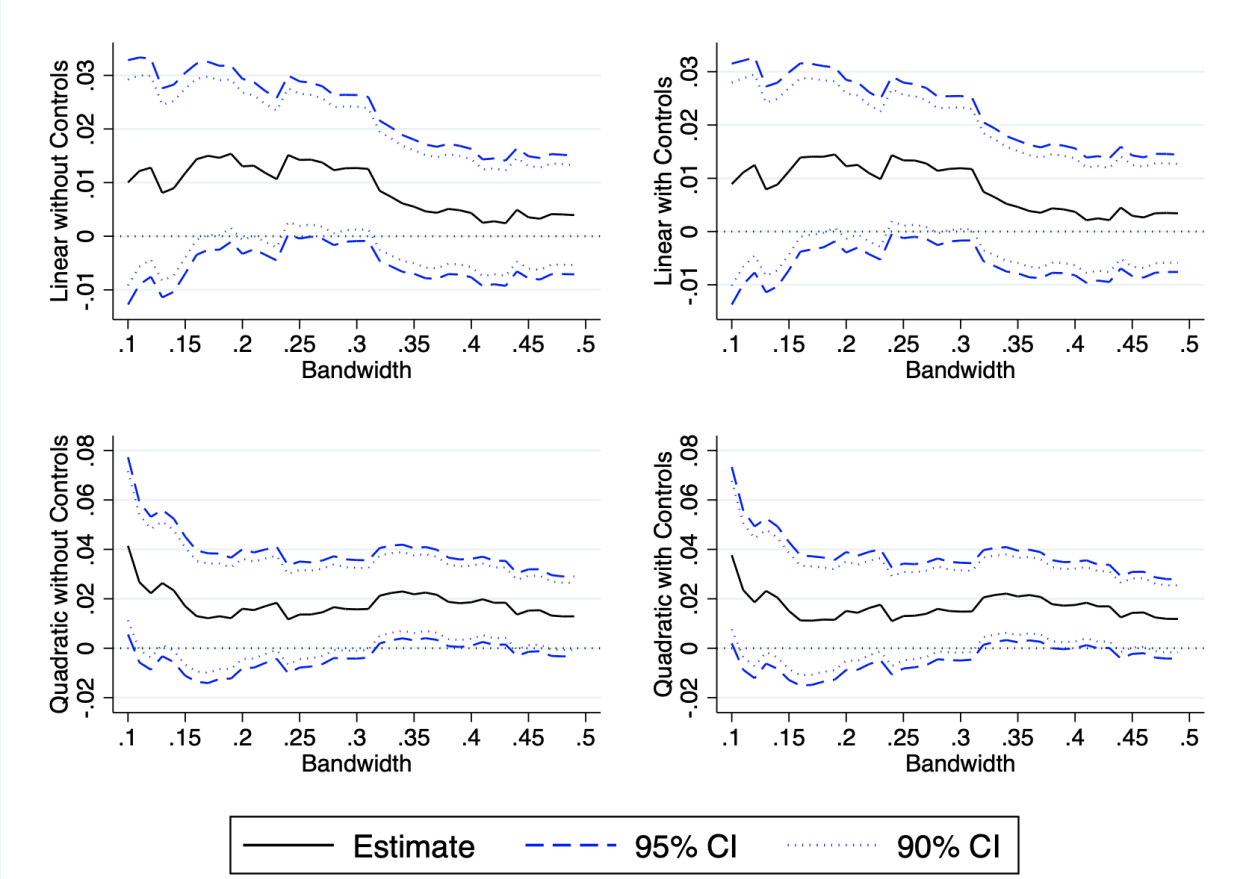

Figure F.1: Registered

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

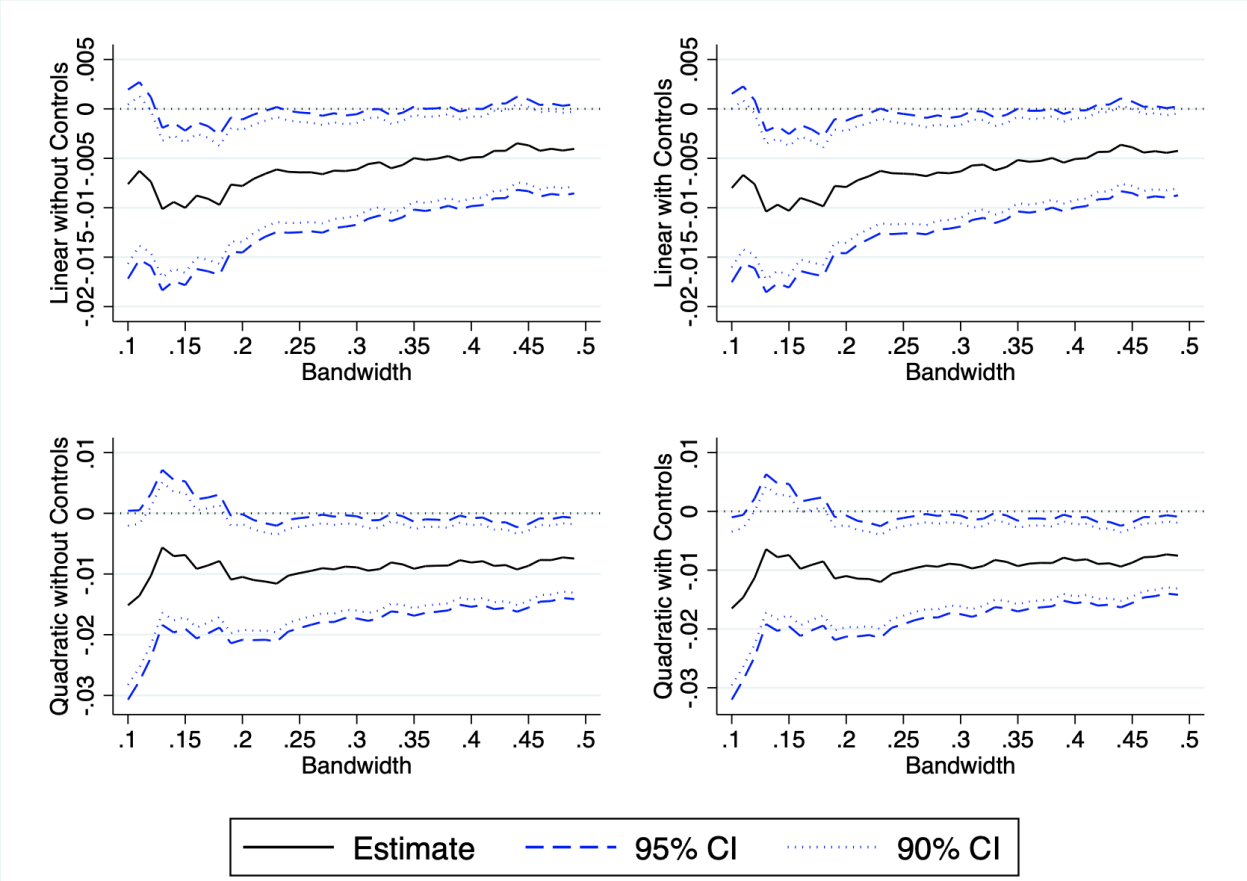

Figure F.2: Republican

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

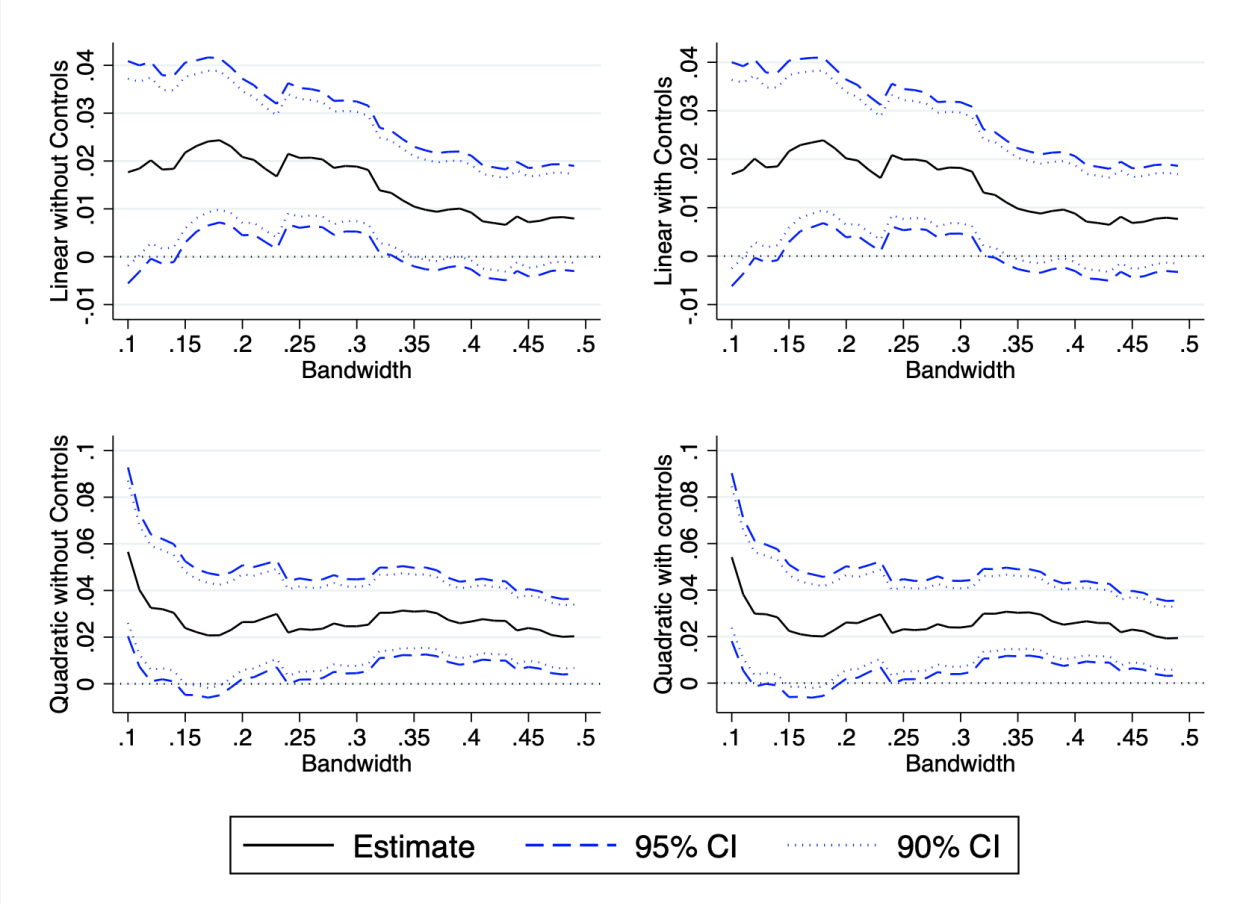

Figure F.3: Democrat or Independent

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

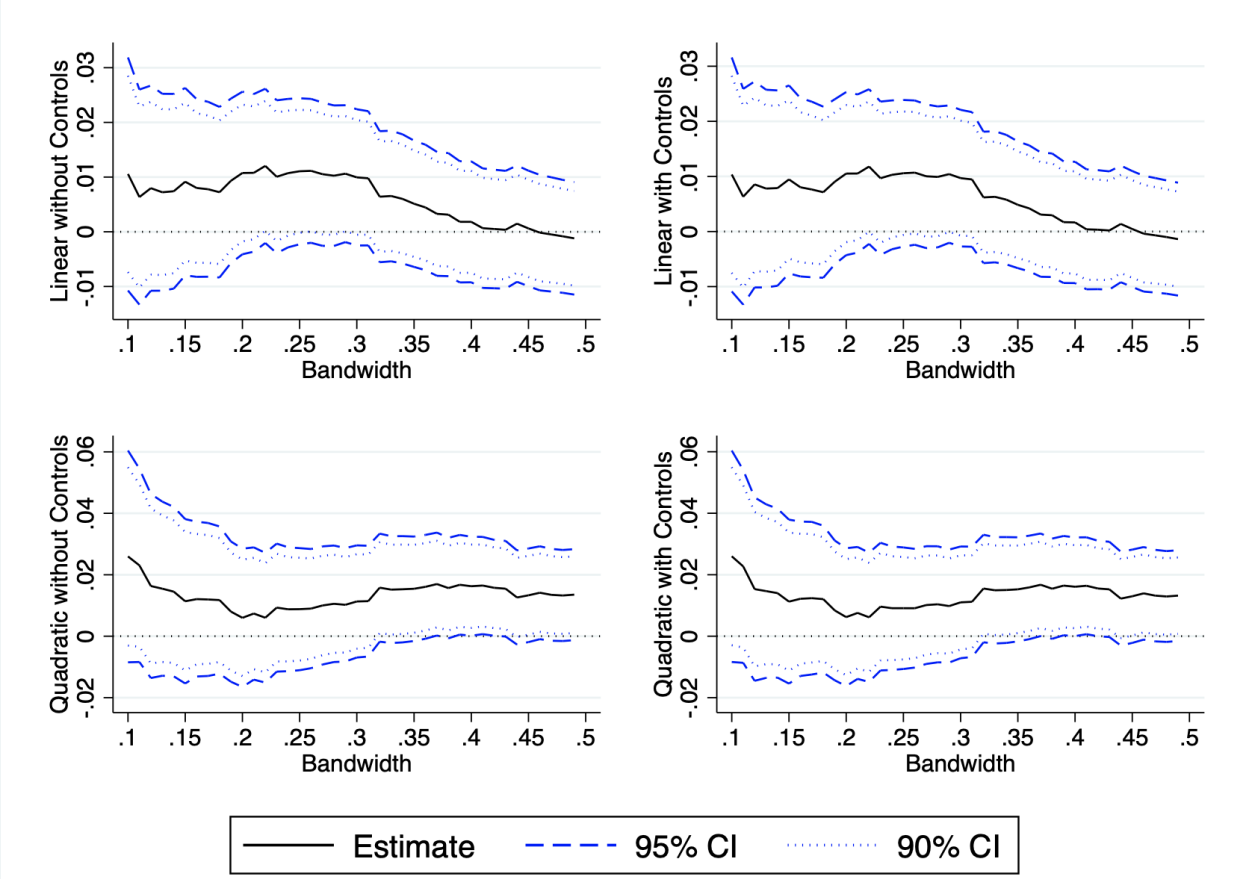

Figure F.4: Democrat

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

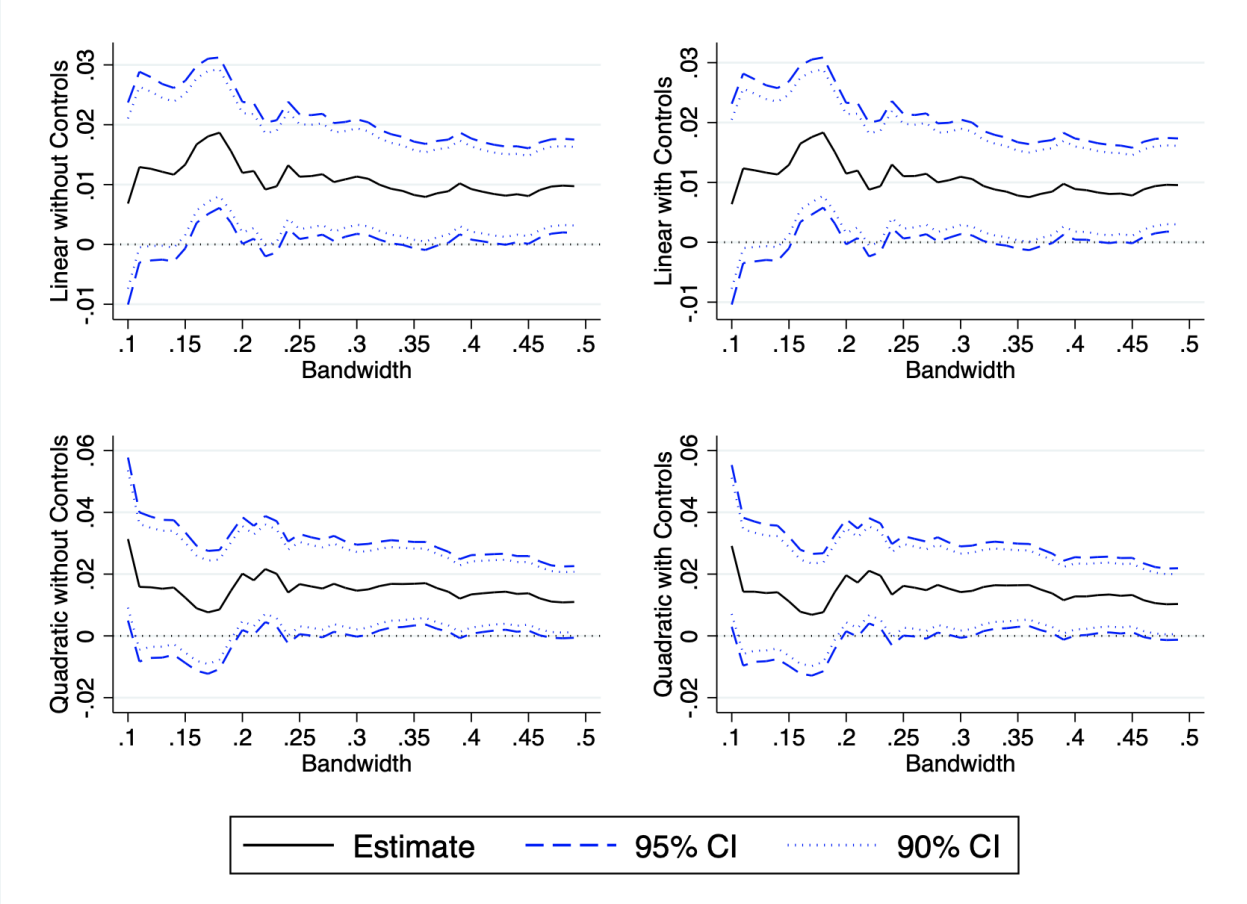

Figure F.5: No Party Preference

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

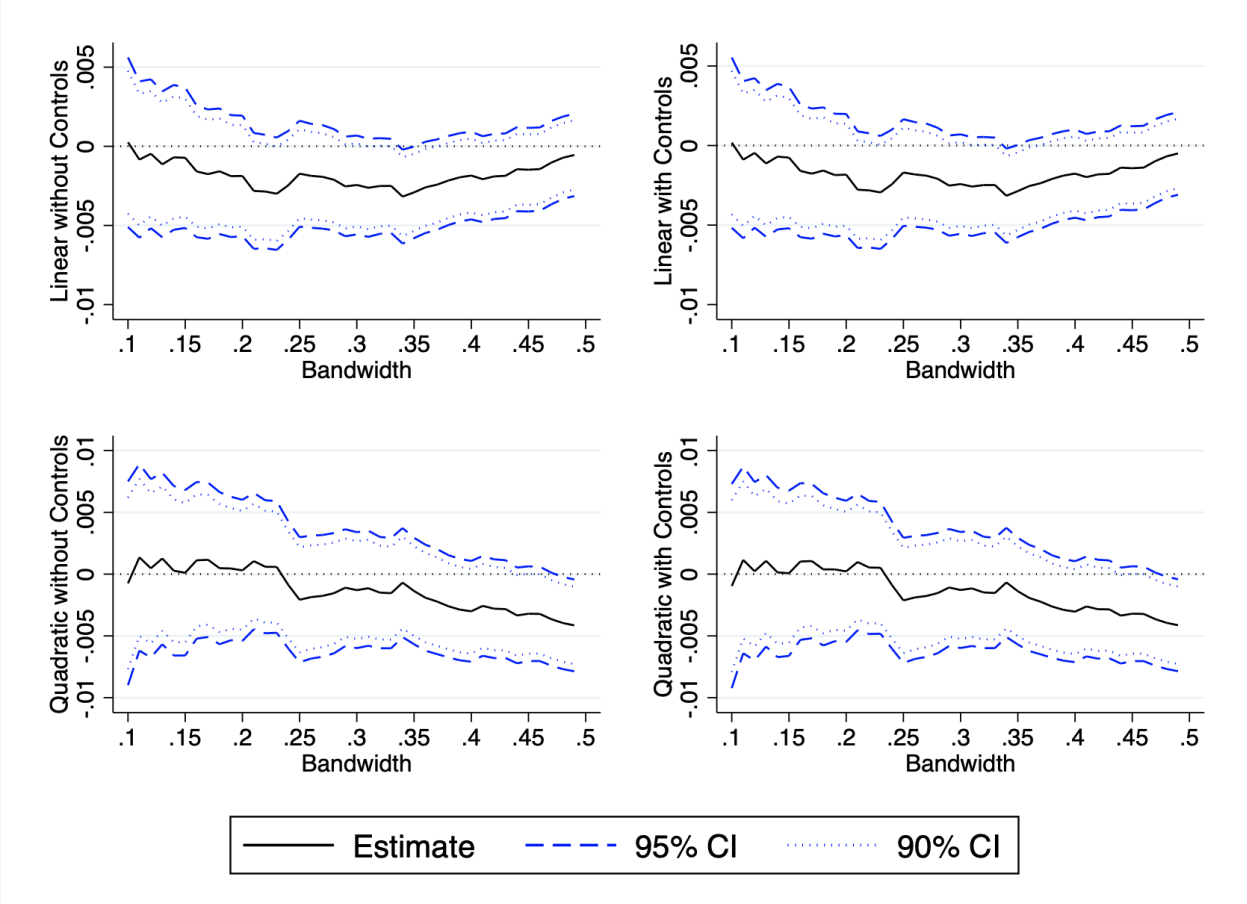

Figure F.6: Third Party

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

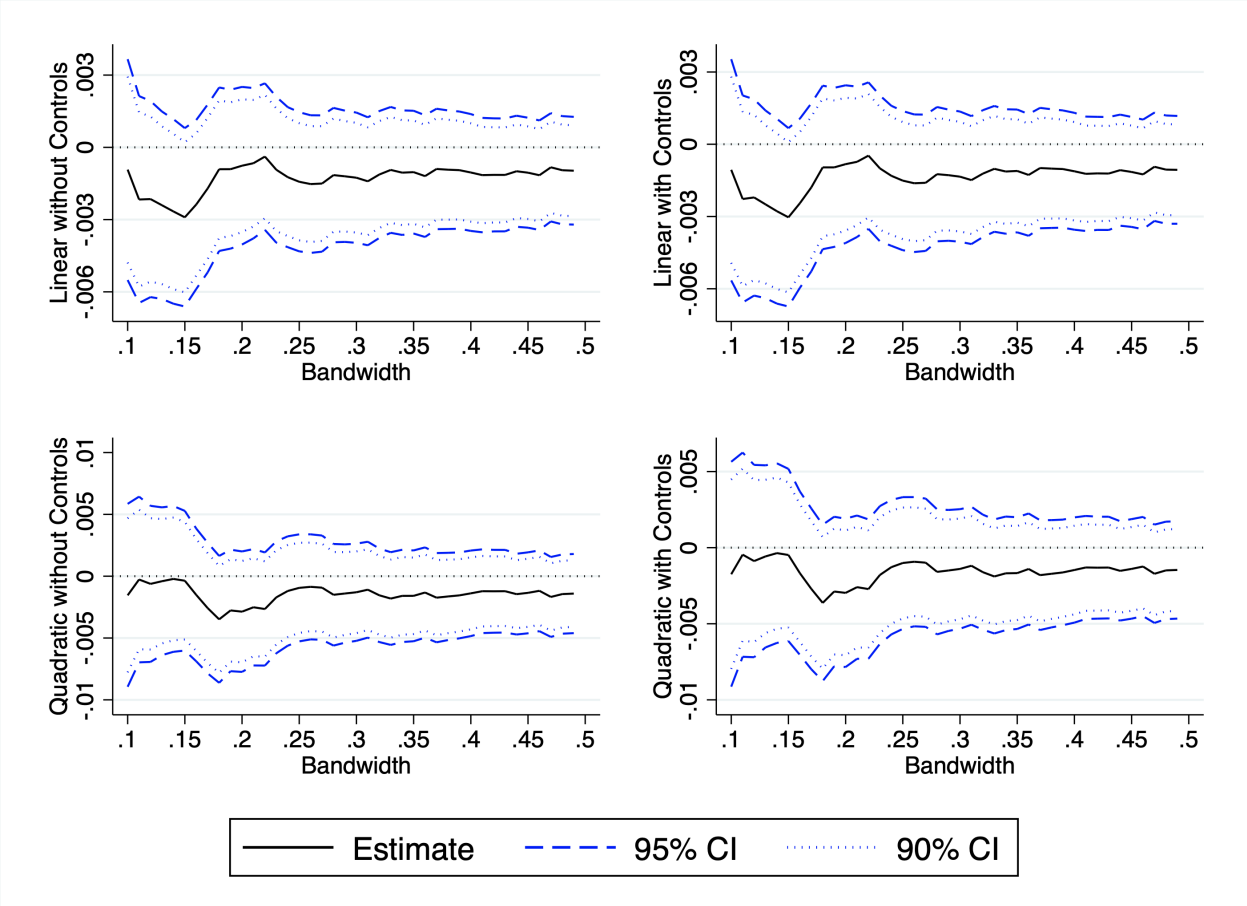

Figure F.7: Democratic Conversion

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

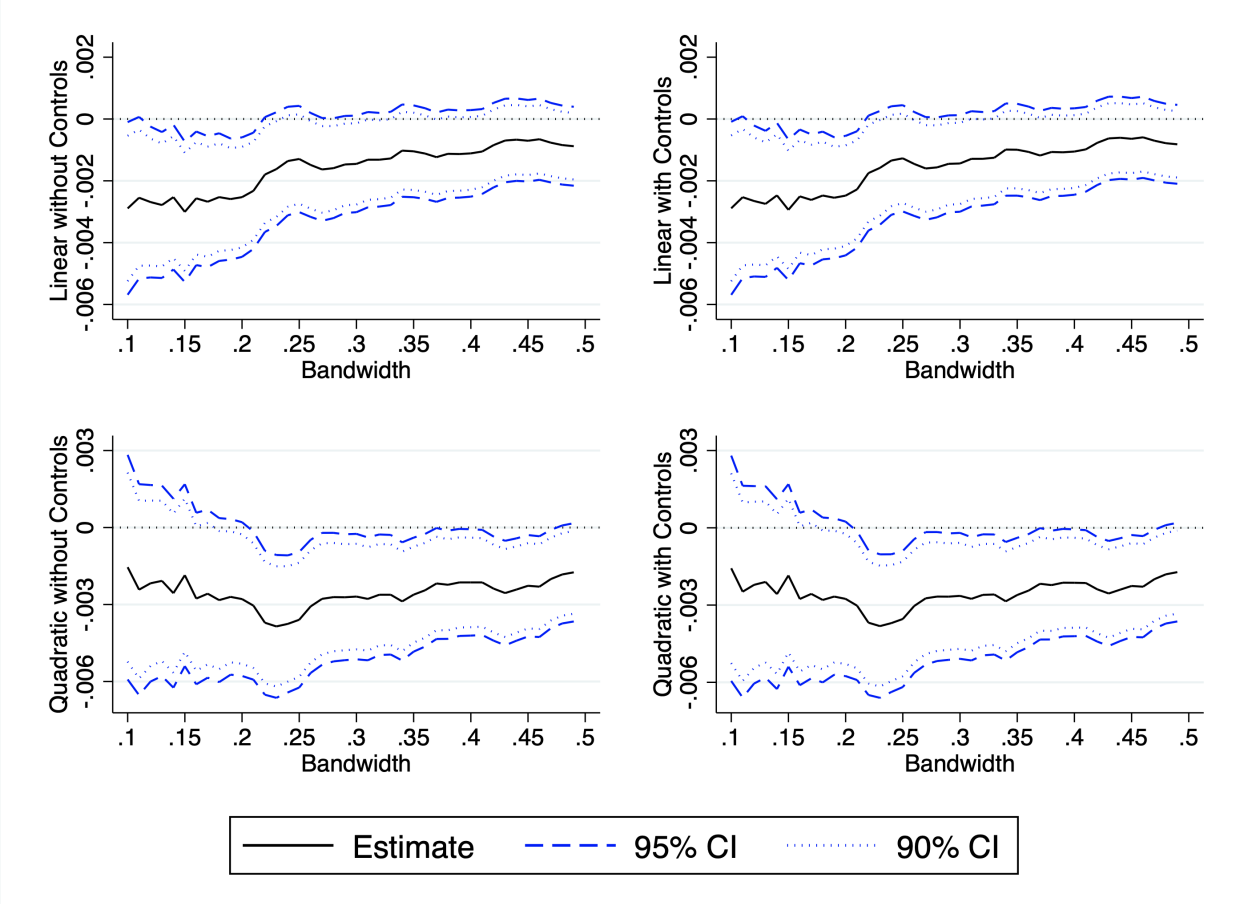

Figure F.8: Republican Conversion

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

Table F.5: Effects of Compulsory Schooling on Partisanship

Location	(1) All	(2) All	(3) FL	(4) FL	(5) CA	(6) CA
A. Political Party Membership						
Republican	-0.0044**	-0.0048**	-0.0027	-0.0038*	-0.0058**	-0.0058**
	(0.0012)	(0.0011)	(0.0018)	(0.0017)	(0.0015)	(0.0015)
Democrat/Independent	0.0044**	0.0048**	0.0027	0.0038*	0.0058**	0.0058**
	(0.0012)	(0.0011)	(0.0018)	(0.0017)	(0.0015)	(0.0015)
Democrat	0.0028*	0.0032*	0.0017	0.0025	0.0036*	0.0038*
	(0.0013)	(0.0013)	(0.0018)	(0.0018)	(0.0018)	(0.0018)
No Party	0.0011	0.0012	0.0014	0.0017	0.0011	0.0008
	(0.0012)	(0.0012)	(0.0019)	(0.0018)	(0.0016)	(0.0016)
Third Party	0.0004	0.0004	-0.0004	-0.0004	0.0011	0.0012
	(0.0006)	(0.0006)	(0.0007)	(0.0007)	(0.0010)	(0.0010)
Bandwidth	70	70	70	70	70	70
Polynomial	2	2	2	2	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	5,110,316	5,110,316	2,380,489	2,380,489	2,729,827	2,729,827

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors in parentheses. Bandwidths are measured in days relative to the compulsory schooling law cutoff birthdate. Democratic and Republican converts are voters who are currently registered with the Democratic and Republican Party in California, but at any time in the past had another party registration status. Controls include sex, state of voter registration, and year of birth fixed effects.

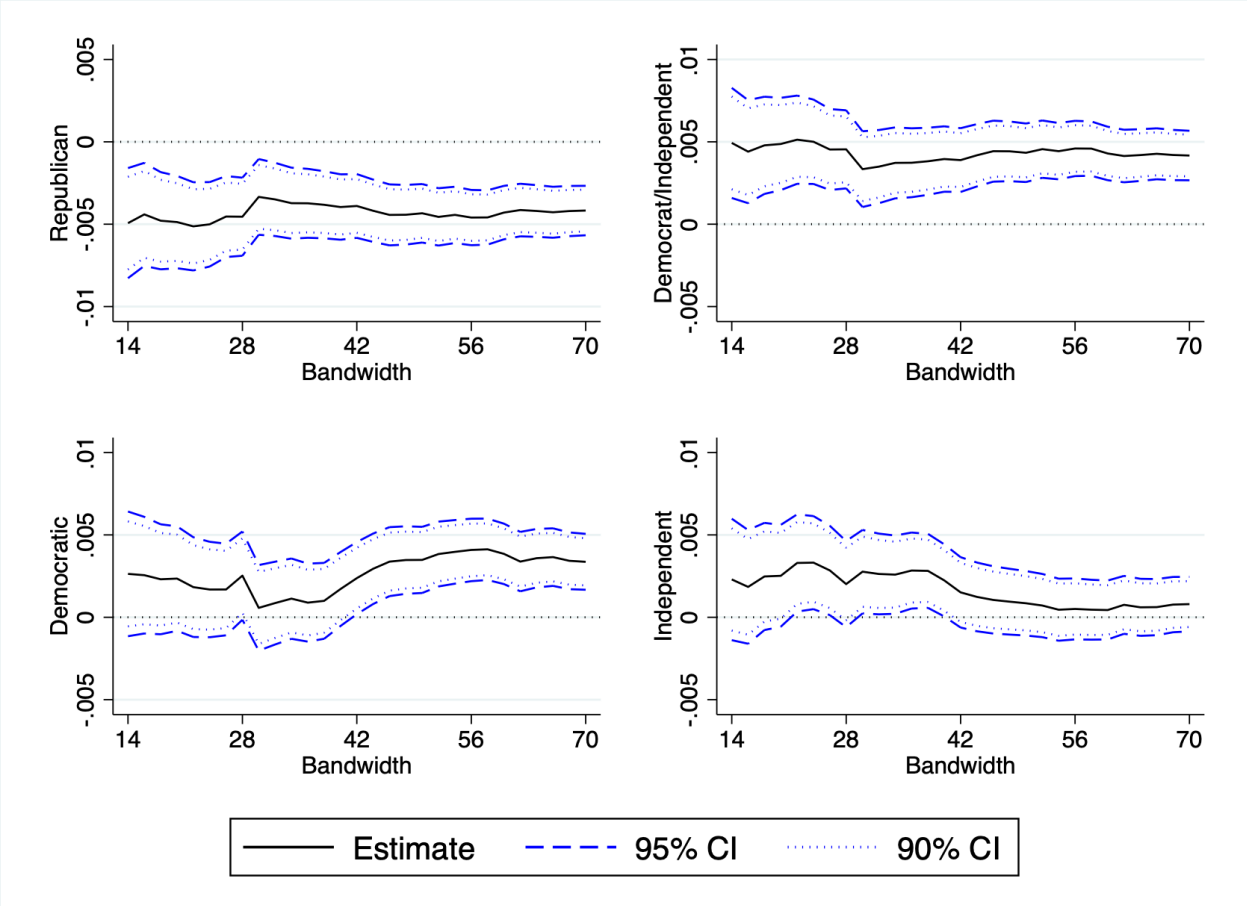

Figure F.9: Robustness of CSL Results by Bandwidth

Note: Treatment effects for each respective outcome variable are shown at each bandwidth, measured in days, on the horizontal axis.

Table F.6: Placebo Test of Compulsory Schooling among Naturalized Immigrants

	(1)	(2)
Location	CA	CA
Republican	-0.0001	0.0005
	(0.0015)	(0.0015)
Democrat/Independent	0.0001	-0.0005
	(0.0015)	(0.0015)
Democrat	-0.0009	-0.0008
	(0.0018)	(0.0018)
No Party	0.0013	0.0008
	(0.0016)	(0.0016)
Third Party	-0.0003	-0.0005
	(0.0007)	(0.0007)
Bandwidth	70	70
Polynomial	1	1
Controls	No	Yes
Sample Size	1,281,463	1,281,463

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors in parentheses. Bandwidths are measured in days relative to the compulsory schooling law cutoff birthdate. Democratic and Republican converts are voters who are currently registered with the Democratic and Republican Party in California, but at any time in the past had another party registration status. Controls include sex, state of voter registration, and year of birth fixed effects. The immigrant indicator includes individuals whose place of birth includes the following countries or regions: Mexico, Philippines, Vietnam, China, India, Korea, El Salvador, Iran, Taiwan, Guatemala, Hong Kong, Germany, Canada, Japan, United Kingdom, Peru, Thailand, Pakistan, Russia, Nicaragua, Armenia, and Ukraine, as well as any entries containing the term "Foreign-Born". These national origins collectively represent all foreign born entries in voter's place of birth field that appear for more than 0.25 percent of the sample of registered Californian voters who were not born in California but were born within 70 days of the CSL cutoff.

G Falsification Test Appendix

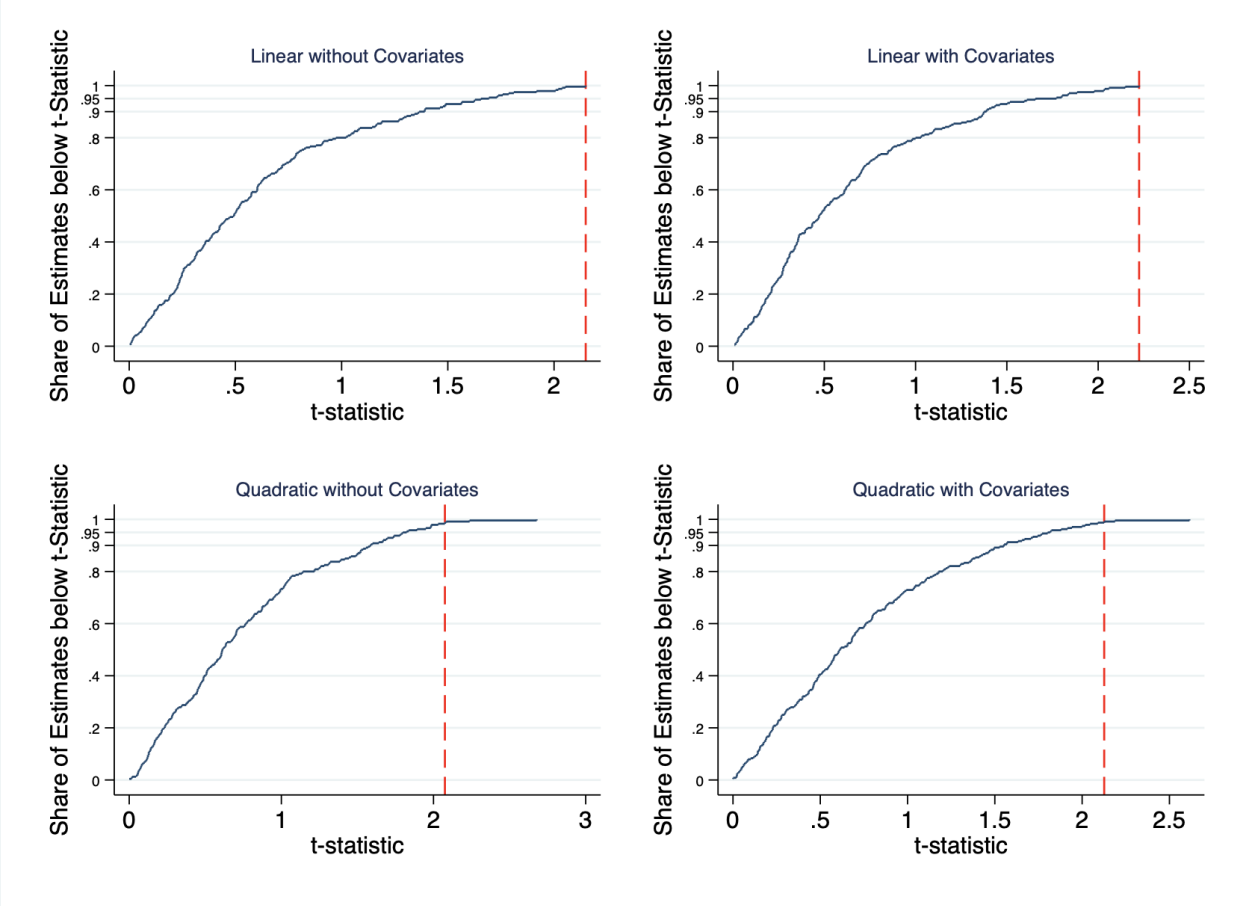

Figure G.1: Republican

Note: Each graph reflects the cumulative distribution of estimated t-statistics using the falsifcation tests described in Section 3.1.2. The red dashed line denotes the t-statistic estimated at the true 96th percentile policy threshold.

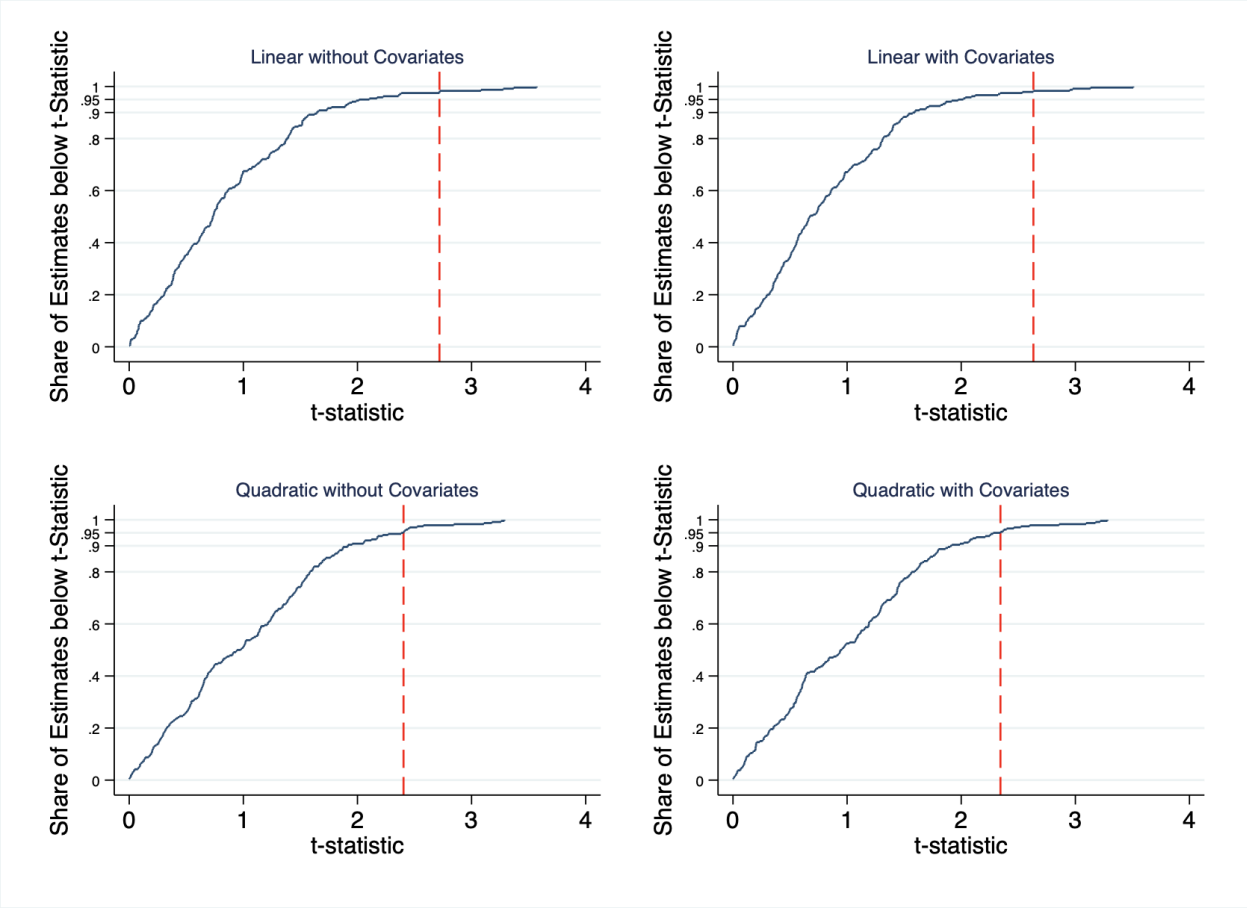

Figure G.2: Democrat or Independent

Note: Each graph reflects the cumulative distribution of estimated t-statistics using the falsifcation tests described in Section 3.1.2. The red dashed line denotes the t-statistic estimated at the true 96th percentile policy threshold.

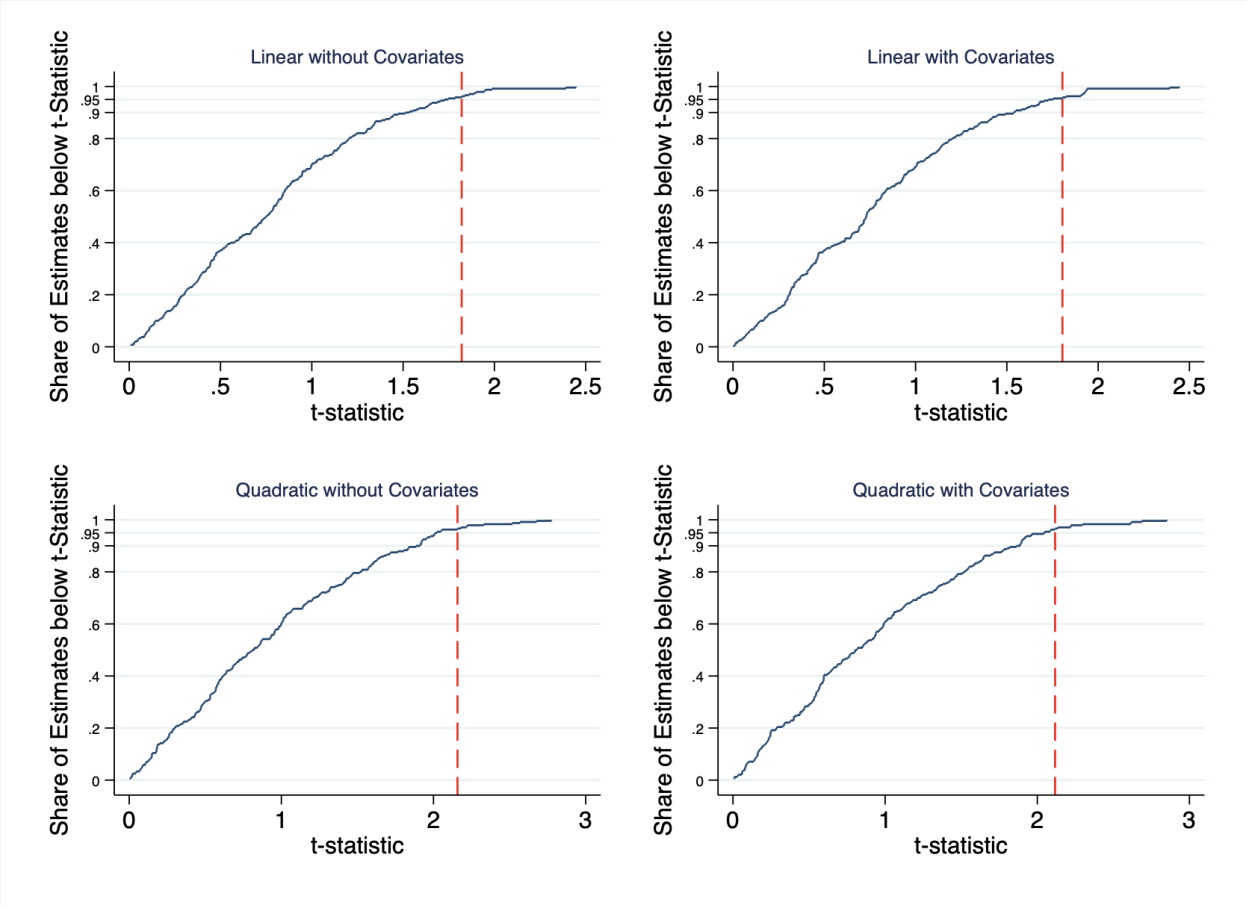

Figure G.3: Republican Conversion

Note: Each graph reflects the cumulative distribution of estimated t-statistics using the falsifcation tests described in Section 3.1.2. The red dashed line denotes the t-statistic estimated at the true 96th percentile policy threshold.

H Mechanisms Appendix

Table H.1: Effects of the UC Top Percent Policy on Enrollment by Student Characteristics

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Race, Ethnicity, and Nationality						
White	-0.0090**	-0.0093**	-0.0075**	-0.0077**	-0.0092**	-0.0100**
	(0.0024)	(0.0024)	(0.0022)	(0.0021)	(0.0031)	(0.0031)
Asian	0.0150**	0.0152**	0.0110**	0.0110**	0.0173**	0.0179**
	(0.0027)	(0.0026)	(0.0022)	(0.0022)	(0.0032)	(0.0031)
Black	0.0007	0.0007	0.0008+	0.0008+	0.0002	0.0002
	(0.0005)	(0.0004)	(0.0004)	(0.0004)	(0.0006)	(0.0006)
Hispanic	-0.0102**	-0.0098**	-0.0062**	-0.0060**	-0.0111**	-0.0106**
	(0.0012)	(0.0012)	(0.0009)	(0.0008)	(0.0012)	(0.0012)
International	0.0024**	0.0023**	0.0020**	0.0020**	0.0027**	0.0026**
	(0.0004)	(0.0003)	(0.0003)	(0.0003)	(0.0004)	(0.0004)
B. Peer Family Income						
Median Income	2958.64**	2830.96**	2708.18**	2653.35**	3516.16**	3324.75**
	(379.12)	(356.97)	(329.26)	(307.64)	(468.53)	(444.65)
Bottom 80 Percent	-0.0133**	-0.0127**	-0.0121**	-0.0118**	-0.0161**	-0.0152**
	(0.0017)	(0.0016)	(0.0014)	(0.0014)	(0.0021)	(0.0020)
Top 5 Percent	0.0134**	0.0128**	0.0120**	0.0118**	0.0151**	0.0143**
	(0.0014)	(0.0013)	(0.0012)	(0.0011)	(0.0017)	(0.0016)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Median Income" in this context refers to the median family income of peers at a given campus. "Bottom 80 Percent" and "Top 5 Percent" refer to the fraction of students at a given campus who hail from families within a given range of the household income distribution within the United States. Data are from Opportunity Insights.

Table H.2: Effects of the UC Top Percent Policy on Enrollment by Imputed Peer Ideology

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Self-Reported Freshman Ideology						
Far-Right Peers	-0.0003**	-0.0003**	-0.0003**	-0.0003**	-0.0004**	-0.0004**
	(0.0001)	(0.0001)	(0.0001)	(0.0001)	(0.0001)	(0.0001)
Conservative Peers	-0.0007	-0.0007	-0.0008+	-0.0008+	-0.0007	-0.0008
	(0.0005)	(0.0005)	(0.0005)	(0.0005)	(0.0007)	(0.0007)
Moderate Peers	-0.0035**	-0.0034**	-0.0030**	-0.0029**	-0.0041**	-0.0040**
	(0.0005)	(0.0005)	(0.0004)	(0.0004)	(0.0006)	(0.0005)
Liberal Peers	0.0045**	0.0045**	0.0040**	0.0039**	0.0050**	0.0050**
	(0.0007)	(0.0007)	(0.0006)	(0.0006)	(0.0009)	(0.0009)
Far-Left Peers	0.0001+	0.0001+	0.0001	0.0001	0.0001*	0.0001*
	(0.0000)	(0.0000)	(0.0000)	(0.0000)	(0.0001)	(0.0001)
B. GOP Share of Institution's Graduates						
GOP Graduate Share	-0.0053**	-0.0053**	-0.0051**	-0.0051**	-0.0058**	-0.0057**
	(0.0007)	(0.0007)	(0.0006)	(0.0006)	(0.0009)	(0.0009)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "GOP Share" refers to the share of registered voters from my sample who attended a given institution that are a member of the Republican Party in 2021. Data on other outcomes are imputed from UCLA's HERI surveys using the method described in Section 4.

Table H.3: Effects of the UC Top Percent Policy on Enrollment by Imputed Peer Religion

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Self-Reported Freshman Religion						
Protestant Peers	-0.0030**	-0.0030**	-0.0028**	-0.0028**	-0.0031**	-0.0032**
	(0.0008)	(0.0008)	(0.0007)	(0.0007)	(0.0010)	(0.0010)
Catholic Peers	-0.0037**	-0.0036**	-0.0035**	-0.0033**	-0.0043**	-0.0043**
	(0.0007)	(0.0007)	(0.0006)	(0.0006)	(0.0009)	(0.0009)
Jewish Peers	0.0022**	0.0021**	0.0017**	0.0016**	0.0023**	0.0023**
	(0.0003)	(0.0002)	(0.0002)	(0.0002)	(0.0003)	(0.0003)
Other Peers	0.0013**	0.0013**	0.0013**	0.0013**	0.0014**	0.0015**
	(0.0004)	(0.0004)	(0.0004)	(0.0003)	(0.0005)	(0.0005)
No Religion Peers	0.0034**	0.0034**	0.0033**	0.0032**	0.0037**	0.0038**
	(0.0008)	(0.0007)	(0.0007)	(0.0007)	(0.0010)	(0.0010)
B. Aggregate Self-Reported Christians						
Christian Peers	-0.0067**	-0.0065**	-0.0063**	-0.0061**	-0.0074**	-0.0075**
	(0.0013)	(0.0012)	(0.0011)	(0.0011)	(0.0016)	(0.0016)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Christian Peers" is a simple aggregation of the share of peers who self identify as Catholic or Protestant. Data are imputed from UCLA's HERI surveys using the method described in Section 4.

Table H.4: Effects of the UC Top Percent Policy on Enrollment by Imputed Faculty Ideology

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Self-Reported Faculty Ideology						
Far-Right Faculty	-0.0000	-0.0000	-0.0000	-0.0000	-0.0000	-0.0000
	(0.0000)	(0.0000)	(0.0000)	(0.0000)	(0.0000)	(0.0000)
Conservative Faculty	-0.0041**	-0.0041**	-0.0037**	-0.0037**	-0.0047**	-0.0047**
	(0.0006)	(0.0006)	(0.0006)	(0.0006)	(0.0008)	(0.0008)
Moderate Faculty	-0.0023**	-0.0023**	-0.0019**	-0.0019**	-0.0026**	-0.0026**
	(0.0004)	(0.0004)	(0.0003)	(0.0003)	(0.0005)	(0.0005)
Liberal Faculty	0.0048**	0.0048**	0.0043**	0.0043**	0.0057**	0.0057**
	(0.0007)	(0.0007)	(0.0007)	(0.0006)	(0.0009)	(0.0009)
Far-Left Faculty	0.0014**	0.0014**	0.0013**	0.0013**	0.0017**	0.0017**
	(0.0003)	(0.0002)	(0.0002)	(0.0002)	(0.0003)	(0.0003)
B. Aggregate Left-Liberal Faculty						
Left-Liberal Faculty	0.0063**	0.0062**	0.0056**	0.0056**	0.0073**	0.0074**
	(0.0010)	(0.0010)	(0.0009)	(0.0009)	(0.0013)	(0.0012)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Left-Liberal Faculty" is a simple aggregation of the share of faculty who self-identify as liberal or far-left. Data are imputed from UCLA's HERI surveys using the method described in Section 4.

Table H.5: Effects of the UC Top Percent Policy on Neighborhood Choice

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Census Block Characteristics						
Median Education	0.0357	0.0383	-0.0089	0.0032	0.0954*	0.0849*
	(0.0338)	(0.0288)	(0.0297)	(0.0250)	(0.0421)	(0.0355)
Median Income	157.37	238.73	-1153.84	-707.23	2152.43	1716.09
	(1185.33)	(1036.12)	(1041.79)	(902.27)	(1510.39)	(1317.17)
B. Local Partisanship						
Republican Neighbors	0.0002	0.0006	0.0003	0.0006	-0.0019	-0.0015
	(0.0021)	(0.0021)	(0.0021)	(0.0021)	(0.0031)	(0.0030)
Democratic Neigbhors	0.0009	0.0007	0.0004	0.0001	0.0019	0.0018
	(0.0022)	(0.0021)	(0.0021)	(0.0021)	(0.0031)	(0.0031)
No Party Neighbors	0.0001	-0.0001	-0.0007	-0.0008	0.0008	0.0005
	(0.0009)	(0.0009)	(0.0008)	(0.0008)	(0.0011)	(0.0011)
Third Party Neighbors	-0.0005	-0.0005	0.0001	0.0000	-0.0009*	-0.0008*
	(0.0003)	(0.0003)	(0.0003)	(0.0003)	(0.0004)	(0.0004)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	Varies	Varies	Varies	Varies

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Median Education" refers to the median years of schooling within a Californian registrant's census block. "Median Income" refers to the estimated median household income within a Californian registrant's census block. "Neighbors" refer to the respective proportion of registered voters with a given party registration status within a Californian registrant's local area. Data are from L2's VM2 California voter file.

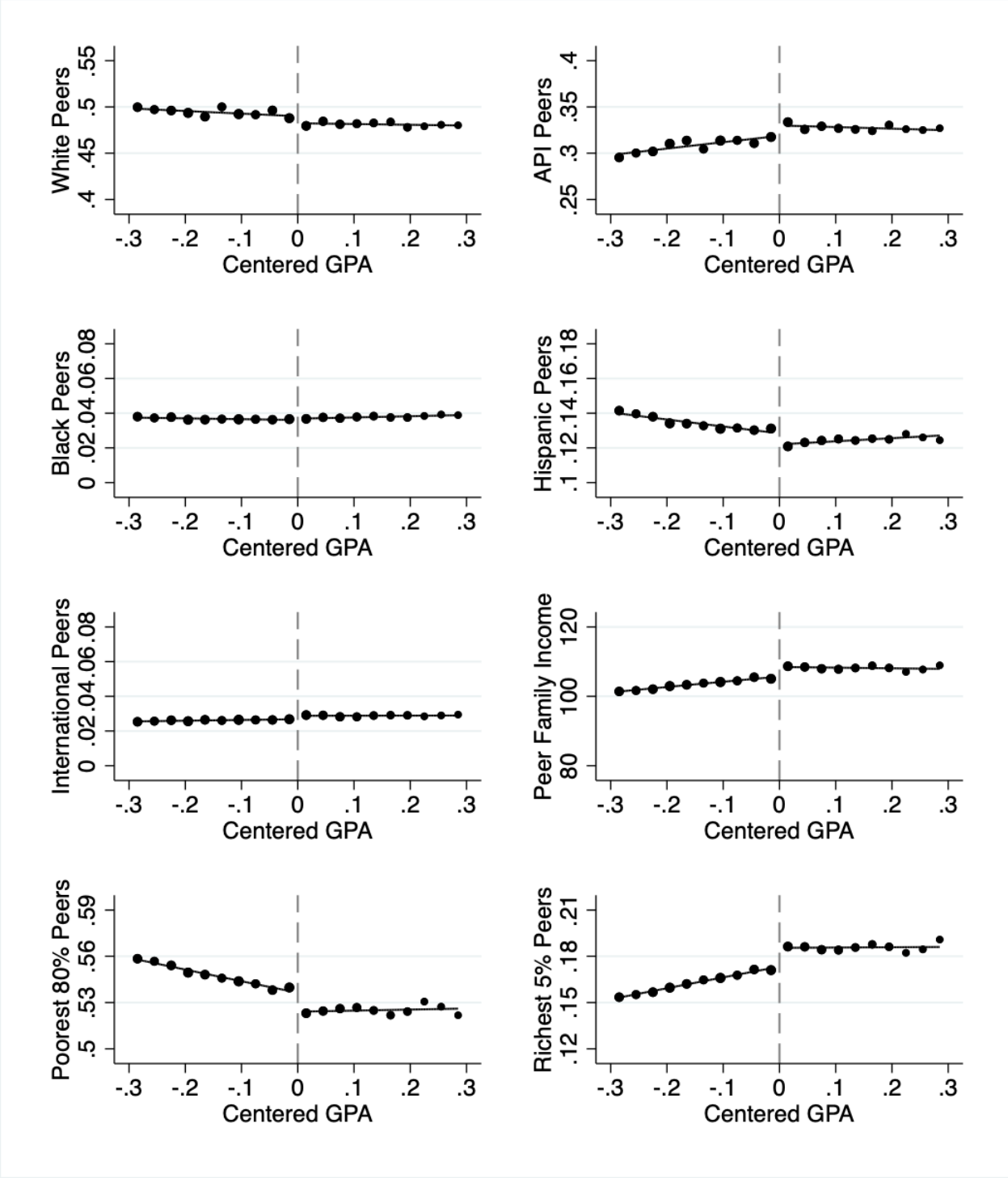

Figure H.1: RD Graphs of Peer Characteristics

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table H.1.

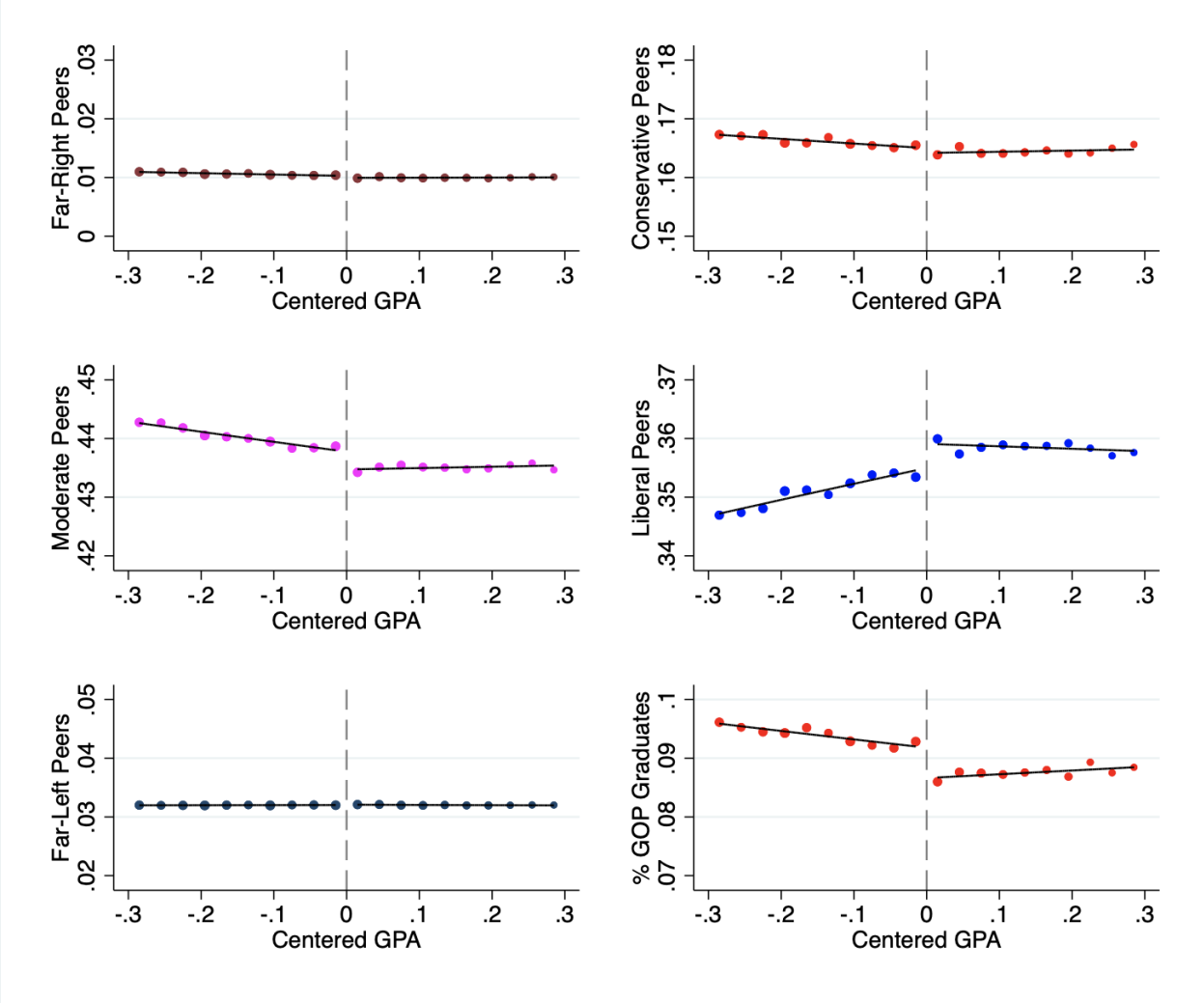

Figure H.2: RD Graphs of Imputed Peer Ideology

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table H.2.

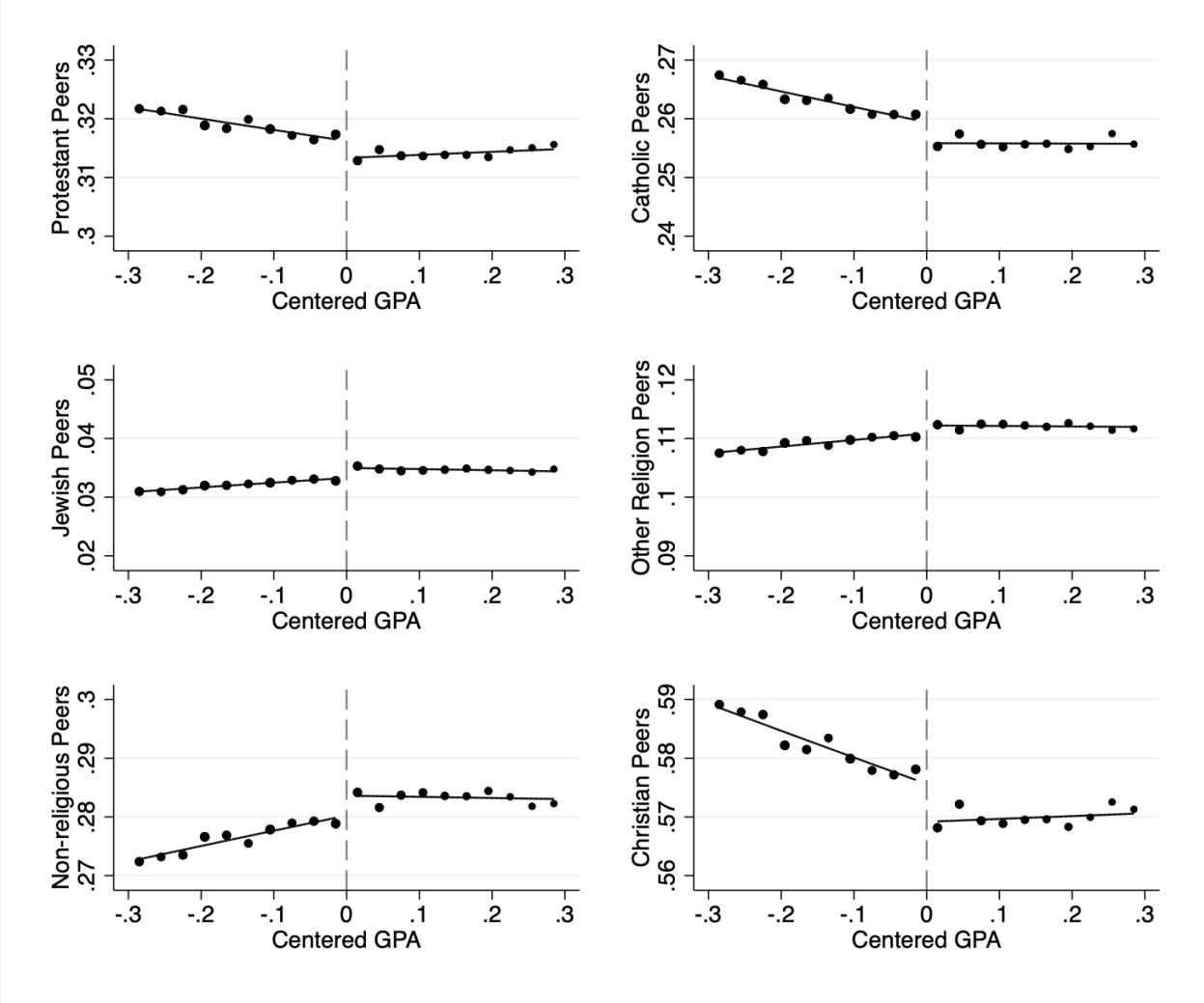

Figure H.3: RD Graphs of Imputed Peer Religion

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table H.3.

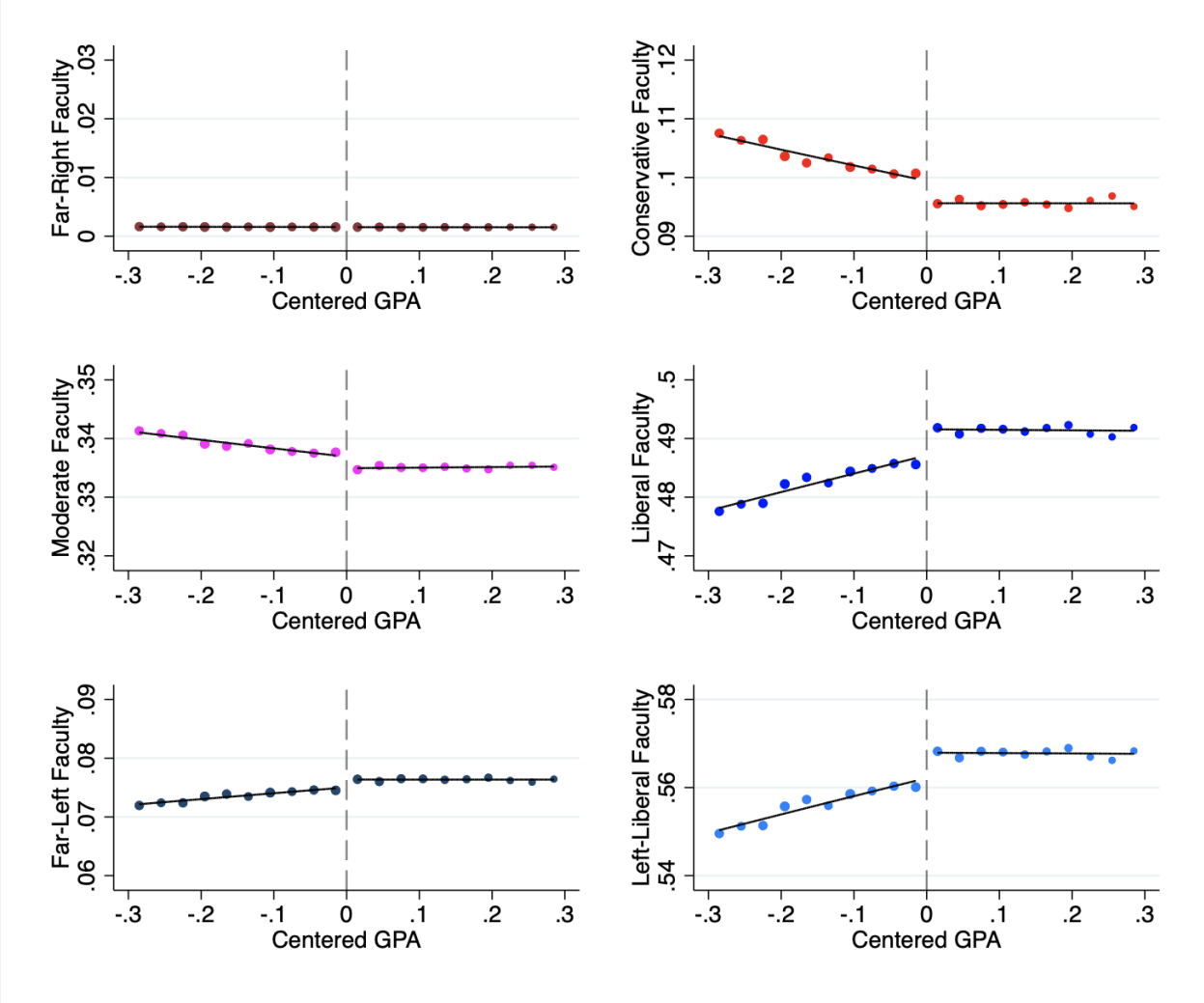

Figure H.4: RD Graphs of Imputed Faculty Ideology

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table H.4.

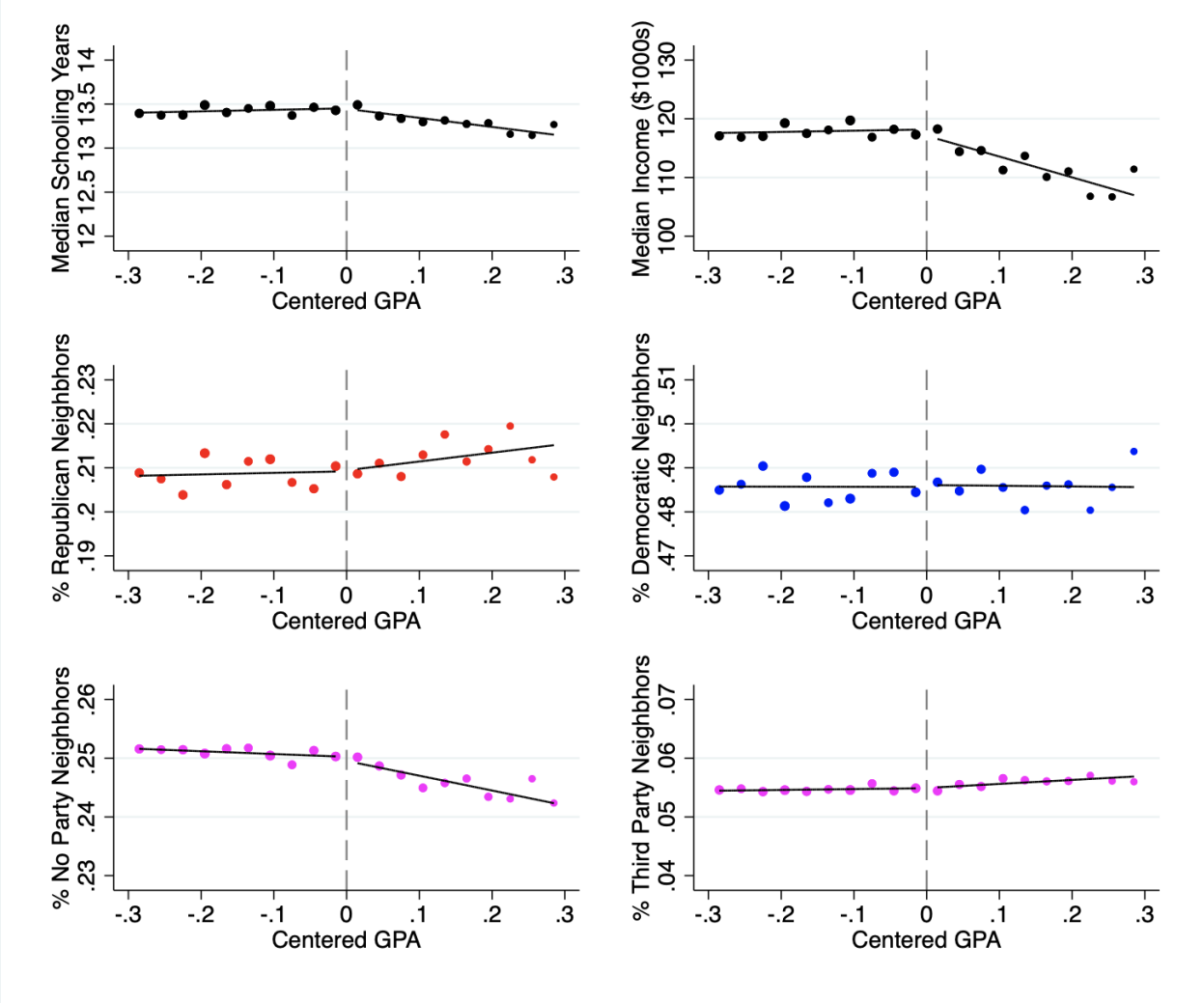

Figure H.5: RD Graphs of Neighborhood Characteristics

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table H.5.

I Voter Turnout Appendix

Given the UC's observed impact on partisanship and extant research on the civic externalities of education, it is important to test the university system's effects on voter turnout (Firoozi and Geyn, 2024). Starting with Figure I.1, I plot eight different measures of voter turnout. First, I show the extensive margin of ever having participated in a regular election and a measure of the total ballots a student cast in regular elections.[21] Next, I decompose the total number of ballots a student cast between 2012 and 2020 by whether they were cast in a presidential or midterm election cycle as well as by whether they were cast in a primary or general election. Finally, the bottom two panels illustrate the number of ballots cast in Republican and Democratic presidential primaries between 2012 and 2020. I find clear visual evidence of an increase in primary ballots cast, particularly in Democratic presidential primaries, and note noisy, positive increases in all other margins of voter turnout beside Republican primary participation.

Table I.1 reflects the results for each of the eight voter turnout outcomes in the preceding figure. Following the same order, Panel A highlights total election participation, Panel B decomposes the number of ballots cast between 2012 and 2020 by the type of election cycle, Panel C decomposes the number of ballots cast by whether they were a primary or general election, and Panel D closes out the table with the number of ballots cast in Republican and Democratic presidential primaries. Each column represents a different specification, varying the RDD bandwidth, inclusion of covariate controls, and the order of a polynomial control for the running variable. Consistent with the visual evidence of discontinuities in Figure I.1, I find that the UC's top percentile policy increases the number of ballots students eventually cast in primary elections by roughly 0.07 to 0.11 votes for each additional UC admission, with most of the effect accruing to Democratic presidential primaries (see Table C.2). Estimates of turnout effects in other elections are positive, with the exception of Republican presidential primaries, but too imprecisely identified to distinguish from zero.

For robustness checks and falsification tests I repeat the procedures used in Section 3.1.2. Tables I.2 through I.4 reflect my main estimates for these outcomes using bias-aware confidence intervals and with high dimensional high school-year fixed effects (Calonico et al., 2014; Kolesar and Rothe, 2018). I also demonstrate the robustness of my point estimates across the full range of potential bandwidths, varying both the order of a polynomial control for the running variable and the inclusion of covariate controls in Appendix Figures I.2 through I.9. For most outcomes, point estimates are fairly stable across bandwidth and

[21] Regular elections in this context refers to all elections coinciding with primary or general elections for federal offices, excluding special elections.

specification but are less consistent than those of registration outcomes. I also reproduce the "synthetic threshold" falsification test for primary election ballots and Democratic presidential primaries in Appendix Figures I.10 and I.11, finding that 4 out of 8 specifications exceed the 95th percentile of synthetic t-statistics at other thresholds and all specifications exceed the 90th percentile of synthetic t-statistics.

Table I.1: Effects of the UC Top Percent Policy on Voter Turnout Outcomes

Outcome	(1)	(2)	(3)	(4)	(5)	(6)
A. Total Voter Turnout Rates						
Ever Voted	0.0088	0.0079	0.0076	0.0069	0.0140	0.0130
	(0.0077)	(0.0076)	(0.0069)	(0.0069)	(0.0101)	(0.0100)
Total Votes Cast	0.0409	0.0374	0.0532	0.0507	0.0687	0.0640
	(0.0326)	(0.0326)	(0.0339)	(0.0338)	(0.0500)	(0.0498)
B. Presidential and Midterm Election Votes						
Presidential Votes	0.0257	0.0232	0.0373	0.0355	0.0468	0.0440
	(0.0220)	(0.0220)	(0.0237)	(0.0237)	(0.0348)	(0.0346)
Midterm Votes	0.0159	0.0152	0.0159	0.0152	0.0218	0.0200
	(0.0121)	(0.0121)	(0.0121)	(0.0121)	(0.0182)	(0.0181)
C. General and Primary Election Votes						
General Votes	0.0073	0.0046	0.0213	0.0197	0.0216	0.0189
	(0.0207)	(0.0206)	(0.0221)	(0.0221)	(0.0325)	(0.0323)
Primary Votes	0.0339*	0.0330*	0.0319*	0.0311*	0.0471*	0.0451*
	(0.0142)	(0.0141)	(0.0143)	(0.0143)	(0.0210)	(0.0210)
D. Partisan Primary Turnout Rates						
Republican Primaries	-0.0032	-0.0034	-0.0021	-0.0024	-0.0022	-0.0025
	(0.0034)	(0.0034)	(0.0033)	(0.0033)	(0.0049)	(0.0049)
Democratic Primaries	0.0170*	0.0167*	0.0185*	0.0183*	0.0263*	0.0258*
	(0.0084)	(0.0083)	(0.0087)	(0.0087)	(0.0128)	(0.0128)
Bandwidth	Optimal	Optimal	0.3	0.3	0.3	0.3
Polynomial	1	1	1	1	2	2
Controls	No	Yes	No	Yes	No	Yes
Sample Size	Varies	Varies	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). "Voted" refers to the extensive margin of ever having cast a ballot in a regularly scheduled federal election and "votes" refers to the aggregate number of ballots cast by an individual in a regularly scheduled federal election. Republican and Democratic primaries refer to the total ballots cast in partisan presidential primary elections.

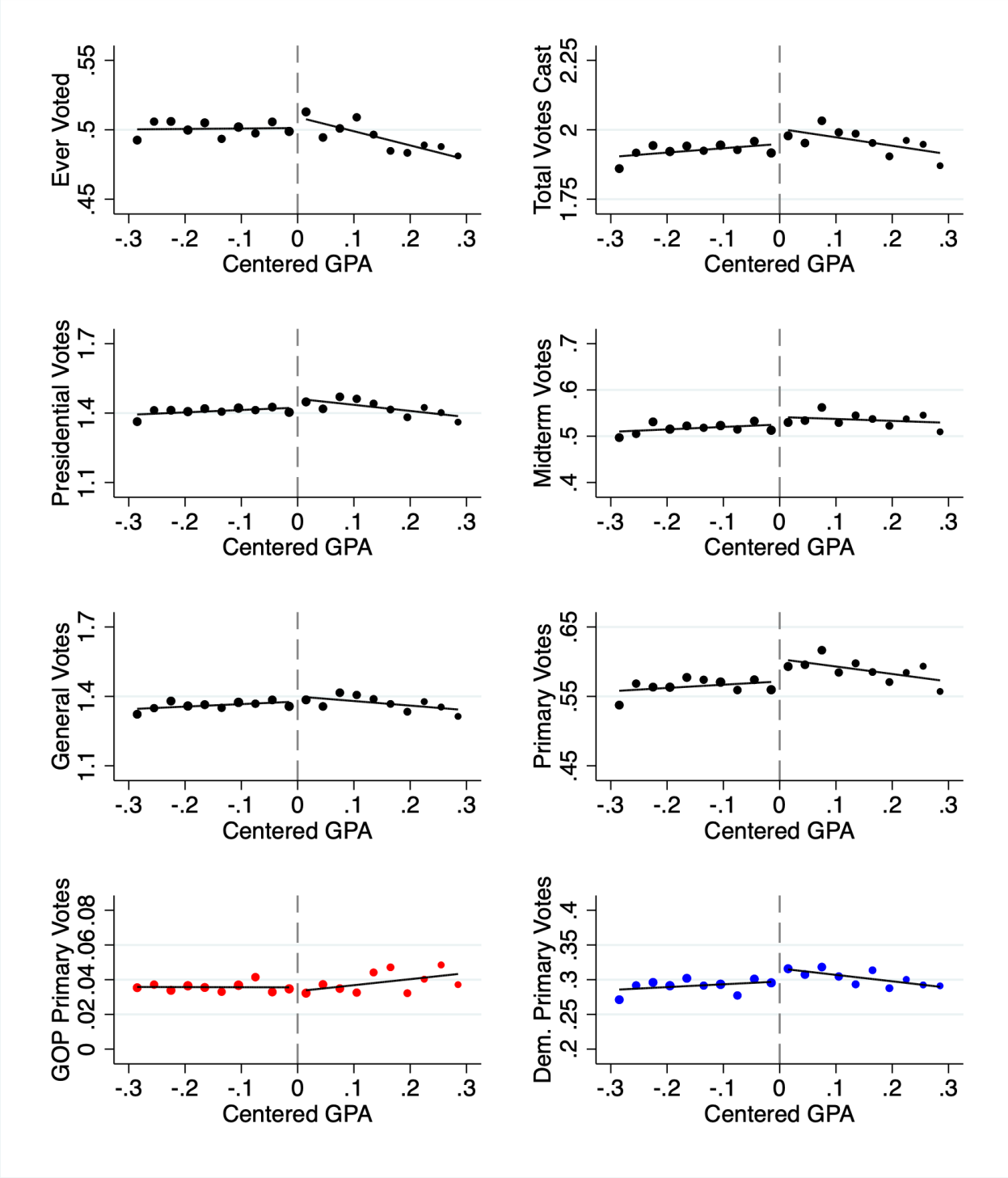

Figure I.1: RD Graphs of Voter Participation Outcomes

Note: Reweighted GPA values are normalized to the 96th percentile cutoff within an individual's high school cohort. Outcomes correspond directly to those in Table I.1.

Table I.2: Effects on Voter Turnout with Bias-Corrected CIs

Outcome	(1)	(2)
Ever Voted		
RD_Estimate	0.0106 (0.0084)	0.0098 (0.0084)
Robust 95% CI	[-.005 ; .033]	[-.006 ; .032]
Robust p-value	0.160	0.183
Total Votes Cast		
RD_Estimate	0.0584 (0.0368)	0.0544 (0.0362)
Robust 95% CI	[-.023 ; .149]	[-.026 ; .143]
Robust p-value	0.151	0.177
Presidential Votes		
RD_Estimate	0.0372 (0.0246)	0.0336 (0.0240)
Robust 95% CI	[-.019 ; .096]	[-.021 ; .091]
Robust p-value	0.185	0.219
Midterm Votes		
RD_Estimate	0.0184 (0.0136)	0.0170 (0.0134)
Robust 95% CI	[-.012 ; .052]	[-.013 ; .05]
Robust p-value	0.224	0.254
General Votes		
RD_Estimate	0.0197 (0.0232)	0.0163 (0.0227)
Robust 95% CI	[-.036 ; .073]	[-.038 ; .067]
Robust p-value	0.503	0.593
Primary Votes		
RD_Estimate	0.0382* (0.0158)	0.0365* (0.0156)
Robust 95% CI	[.005 ; .078]	[.004 ; .076]
Robust p-value	0.027	0.031
Republican Primary Votes		
RD_Estimate	-0.0020 (0.0037)	-0.0023 (0.0037)
Robust 95% CI	[-.011 ; .006]	[-.012 ; .005]
Robust p-value	0.544	0.486
Democratic Primary Votes		
RD_Estimate	0.0206* (0.0092)	0.0202* (0.0091)
Robust 95% CI	[.002 ; .044]	[.002 ; .043]
Robust p-value	0.033	0.034
Bandwidth	MSE-Optimal	MSE-Optimal
Polynomial	1	1
Covariates	No	Yes

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Each row titled "RD Estimate" shows the conventional point estimate and standard errors in parentheses for a given outcome variable. These are calculated using a triangular kernel at the MSE-optimal bandwidth derived from Calonico et al. (2020). The rows "Robust 95% CI" and "Robust p-value" show the bias-corrected confidence interval and the bias-corrected p-value for the same outcome variable (Calonico et al., 2014). These outcomes correspond to those in Table I.1.

Table I.3: Effects on Voter Turnout with Honest CIs

Outcome	(1)	(2)
Ever Voted		
RD Estimate	0.0104 (0.0076)	0.0100 (0.0074)
Robust 95% CI	[-.0061; .0269]	[-.0064; .0265]
Robust 90% CI	[-.0035; .0243]	[-.0039; .0240]
Total Votes Cast		
RD Estimate	0.0582 (0.0367)	0.0588 (0.0373)
Robust 95% CI	[-.0213; .1378]	[-.0210; .1386]
Robust 90% CI	[-.0088; .1253]	[-.0084; .1259]
Presidential Votes		
RD Estimate	0.0407 (0.0259)	0.0139 (0.0074)
Robust 95% CI	[-.0156; .0970]	[-.0156; .0971]
Robust 90% CI	[-.0067; .0882]	[-.0067; .0882]
Midterm Votes		
RD Estimate	0.0154 (0.0123)	0.0181 (0.0134)
Robust 95% CI	[-.0113; .0421]	[-.0094; .0455]
Robust 90% CI	[-.0071; .0379]	[-.0050; .0411]
General Votes		
RD Estimate	0.0213 (0.0245)	0.0212 (0.0243)
Robust 95% CI	[-.0319; .0745]	[-.0320; .0744]
Robust 90% CI	[-.0236; .0661]	[-.0236; .0661]
Primary Votes		
RD Estimate	0.0329 (0.0140)	0.0376 (0.0156)
Robust 95% CI	[.0028; .0631]	[.0060; .0691]
Robust 90% CI	[.0075; .0583]	[.0110; .0641]
Republican Primary Votes		
RD Estimate	-0.0020 (0.0037)	-0.0022 (0.0036)
Robust 95% CI	[-.0100; .0060]	[-.0102; .0058]
Robust 90% CI	[-.0088; .0047]	[-.0089; .0046]
Democratic Primary Votes		
RD Estimate	0.0198 (0.0088)	0.0214 (0.0096)
Robust 95% CI	[.0008; .0389]	[.0018; .0411]
Robust 90% CI	[.0038; .0359]	[.0049; .0380]
Bandwidth	MSE-Optimal	0.3
Polynomial	1	1

Note: Each row titled "RD Estimate" shows the point estimate and standard errors in parentheses for a given outcome variable using a triangular kernel and the bounded seconded derivative method (Kolesar and Rothe, 2018). The rows "Robust 95% CI" and "Robust 90% CI" show the honest confidence intervals for the same outcome variable. These outcomes correspond to those in Table 1.

Table I.4: Effects of the UC Top Percent Policy on Voter Turnout Outcomes

Outcome	(1)	(2)	(3)	(4)
A. Total Voter Turnout Rates				
Ever Voted	0.0073	0.0073	0.0113	0.0115
	(0.0069)	(0.0069)	(0.0102)	(0.0101)
Total Votes Cast	0.0548	0.0543	0.0637	0.0617
	(0.0342)	(0.0342)	(0.0508)	(0.0505)
B. Presidential and Midterm Election Votes				
Presidential Votes	0.0373	0.0365	0.0426	0.0407
	(0.0239)	(0.0239)	(0.0353)	(0.0351)
Midterm Votes	0.0175	0.0178	0.0211	0.0210
	(0.0122)	(0.0122)	(0.0184)	(0.0184)
C. General and Primary Election Votes				
General Votes	0.0226	0.0222	0.0171	0.0161
	(0.0223)	(0.0223)	(0.0329)	(0.0327)
Primary Votes	0.0322*	0.0320*	0.0466*	0.0456*
	(0.0144)	(0.0144)	(0.0214)	(0.0213)
D. Partisan Primary Turnout Rates				
Republican Primaries	-0.0022	-0.0021	-0.0020	-0.0019
	(0.0034)	(0.0034)	(0.0049)	(0.0049)
Democratic Primaries	0.0189*	0.0187*	0.0293*	0.0284*
	(0.0088)	(0.0088)	(0.0130)	(0.0130)
Bandwidth	0.3	0.3	0.3	0.3
Polynomial	1	1	2	2
Controls	No	Yes	No	Yes
HS-Year FEs	Yes	Yes	Yes	Yes
Sample Size	78,195	78,195	78,195	78,195

Note: $^{+}$ $p < 0.1$, * $p < 0.05$, ** $p < 0.01$. Heteroskedasticity robust standard errors clustered on high school cohort in parentheses. Optimal bandwidth refers to the MSE-optimal bandwidth derived from Calonico et al. (2020). These outcomes correspond to those in Table I.1.

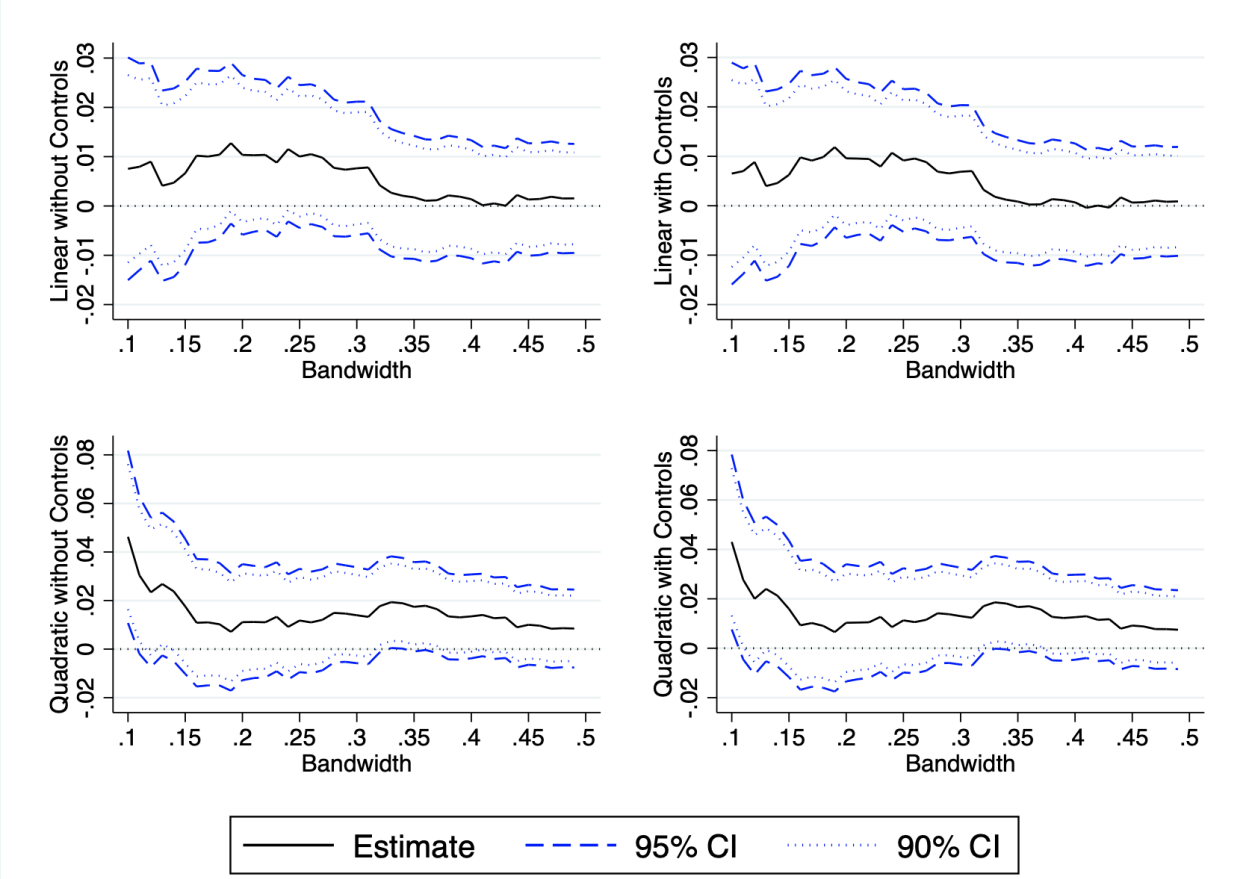

Figure I.2: Ever Voted

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

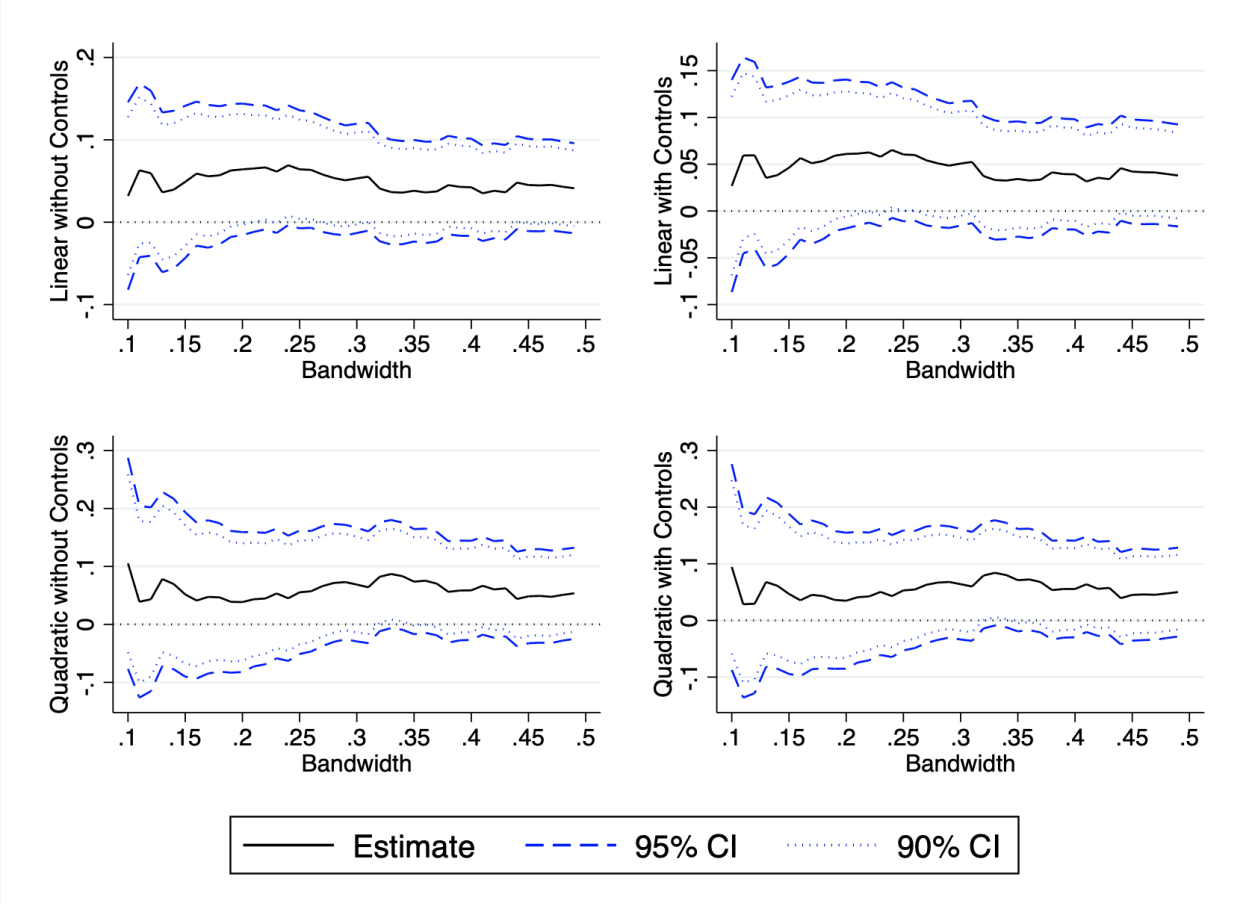

Figure I.3: Total Votes Cast

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

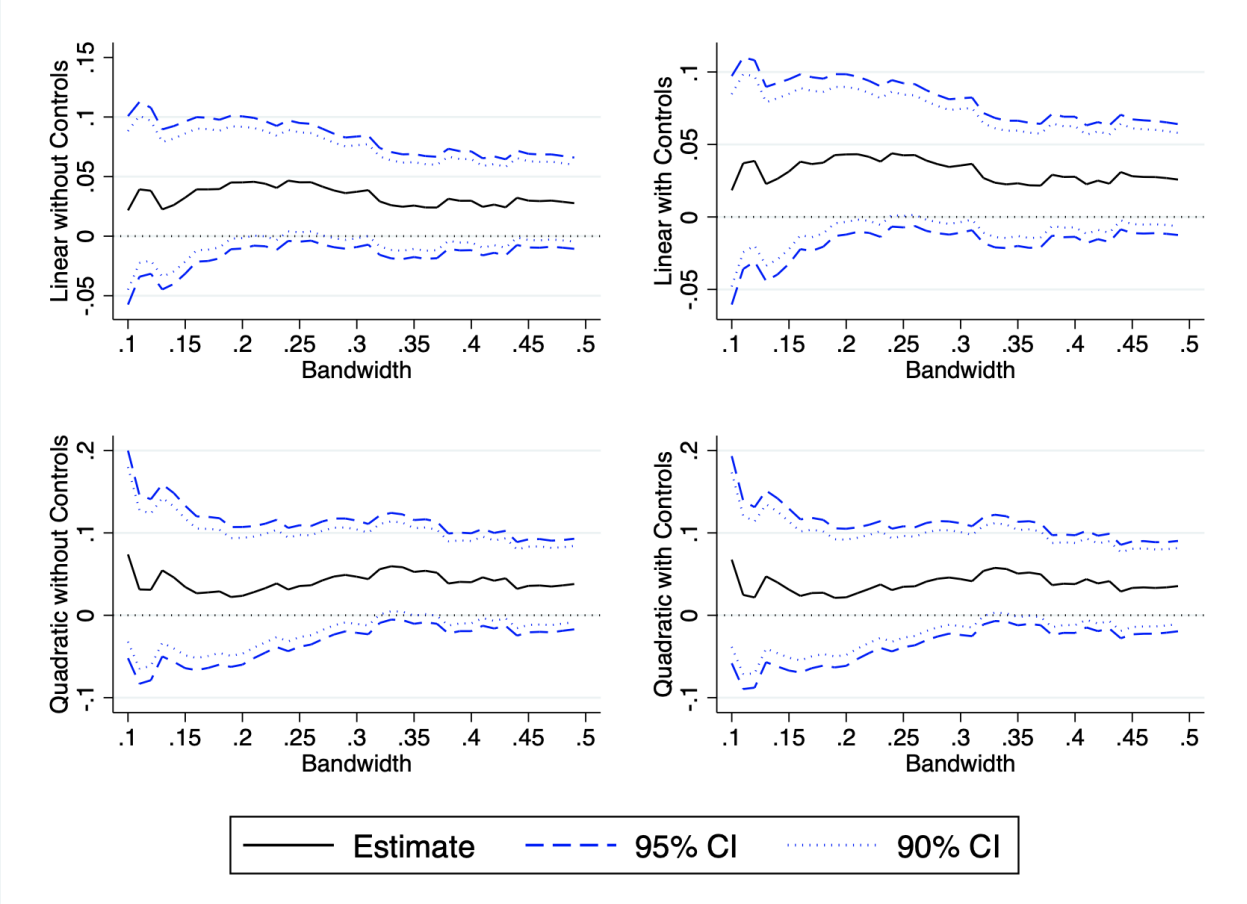

Figure I.4: Presidential Votes

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

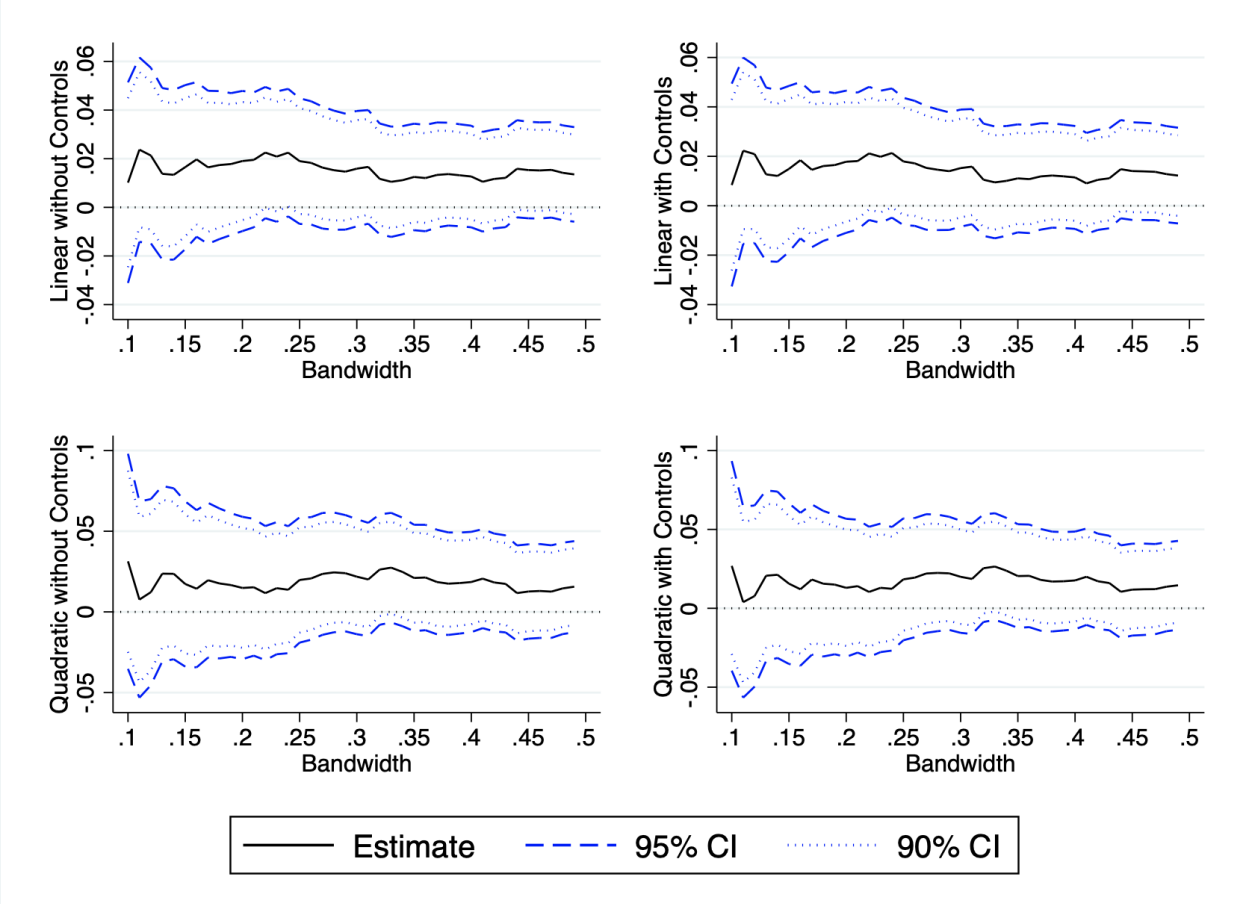

Figure I.5: Midterm Votes

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

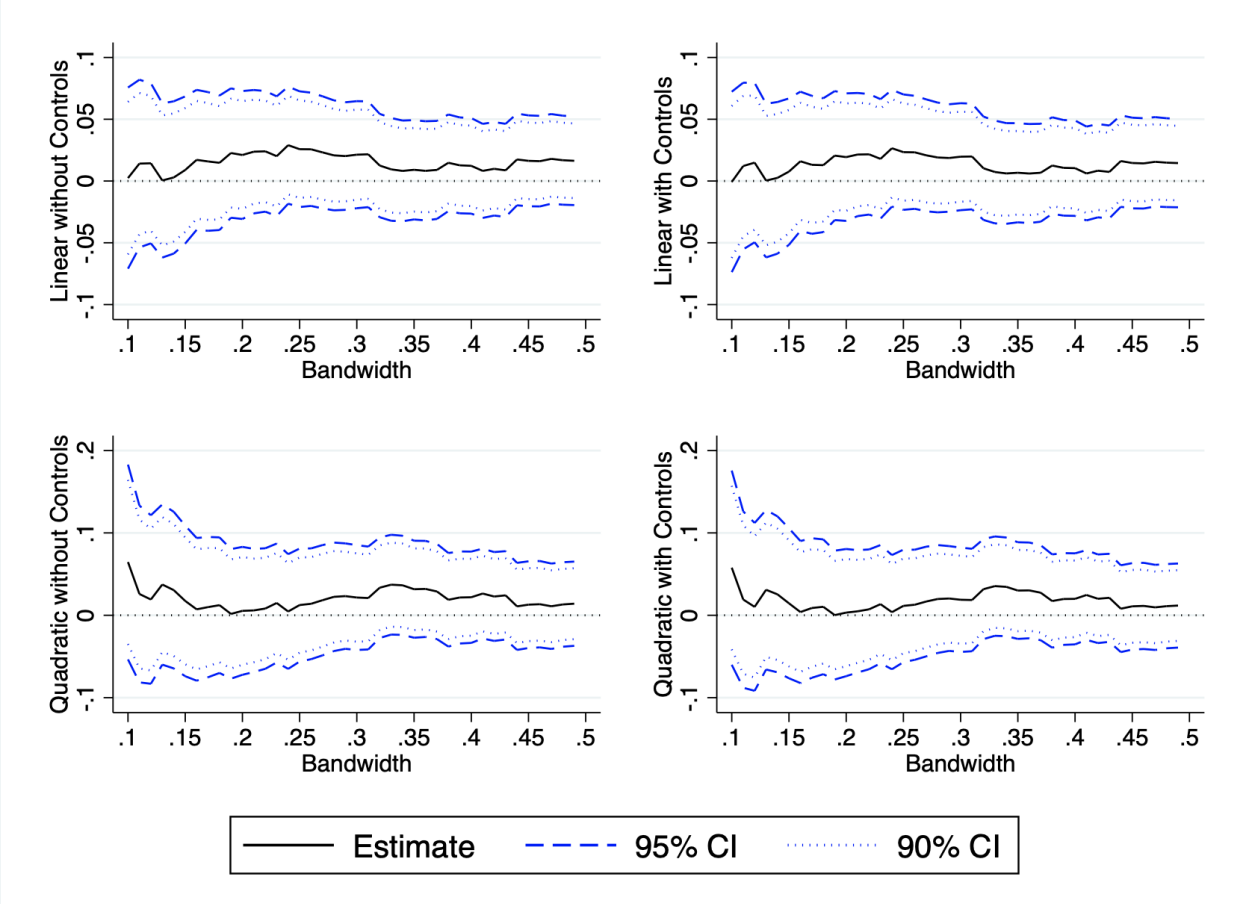

Figure I.6: General Votes

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

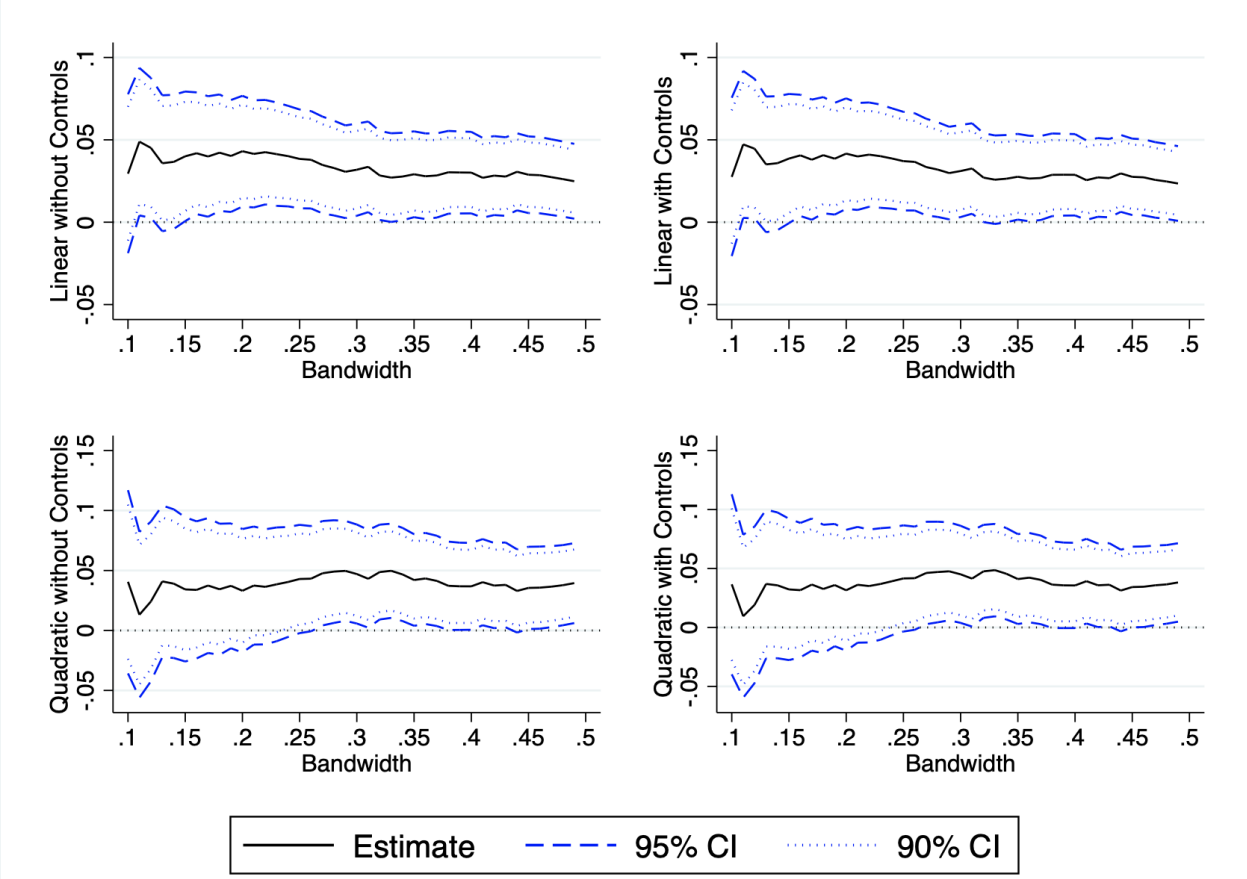

Figure I.7: Primary Votes

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

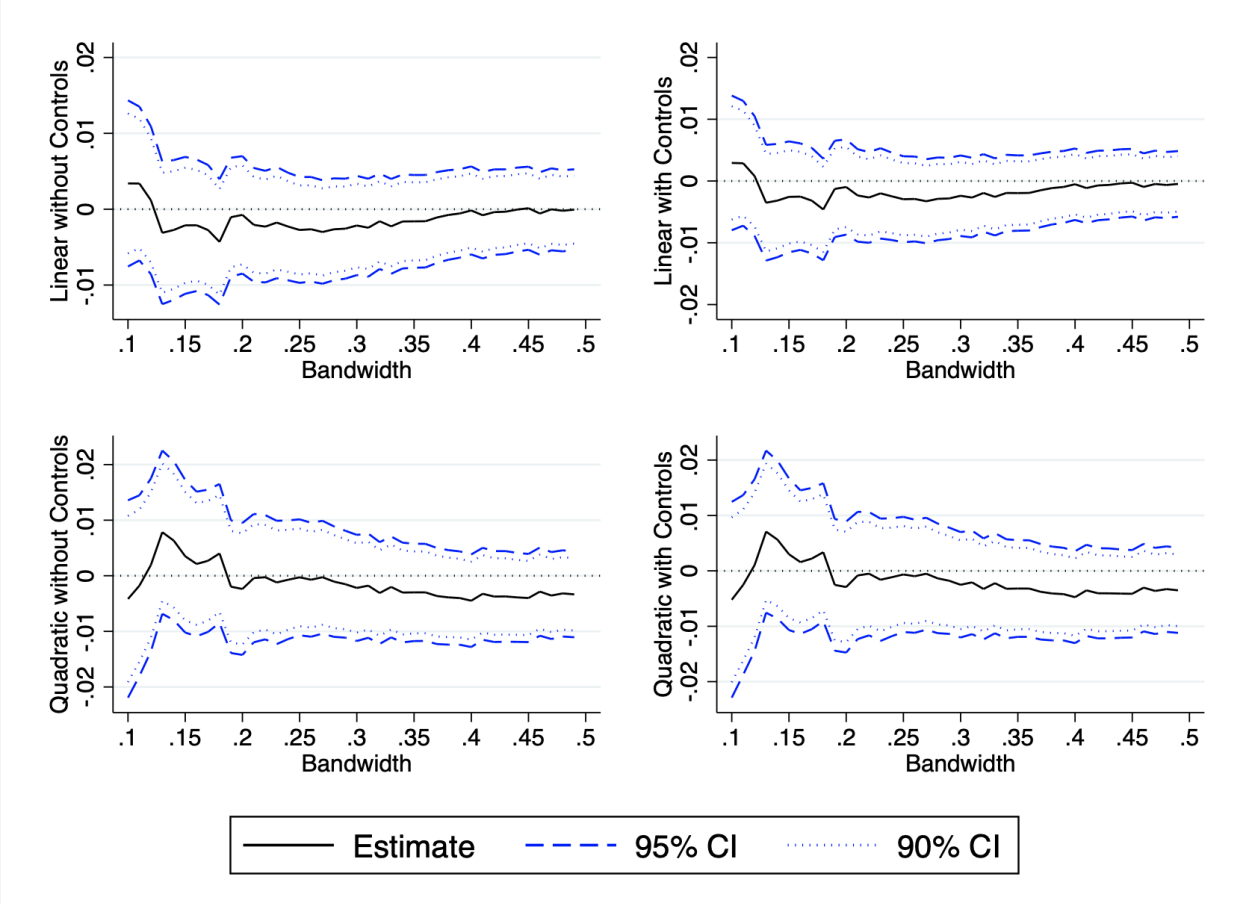

Figure I.8: Republican Primary Votes

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

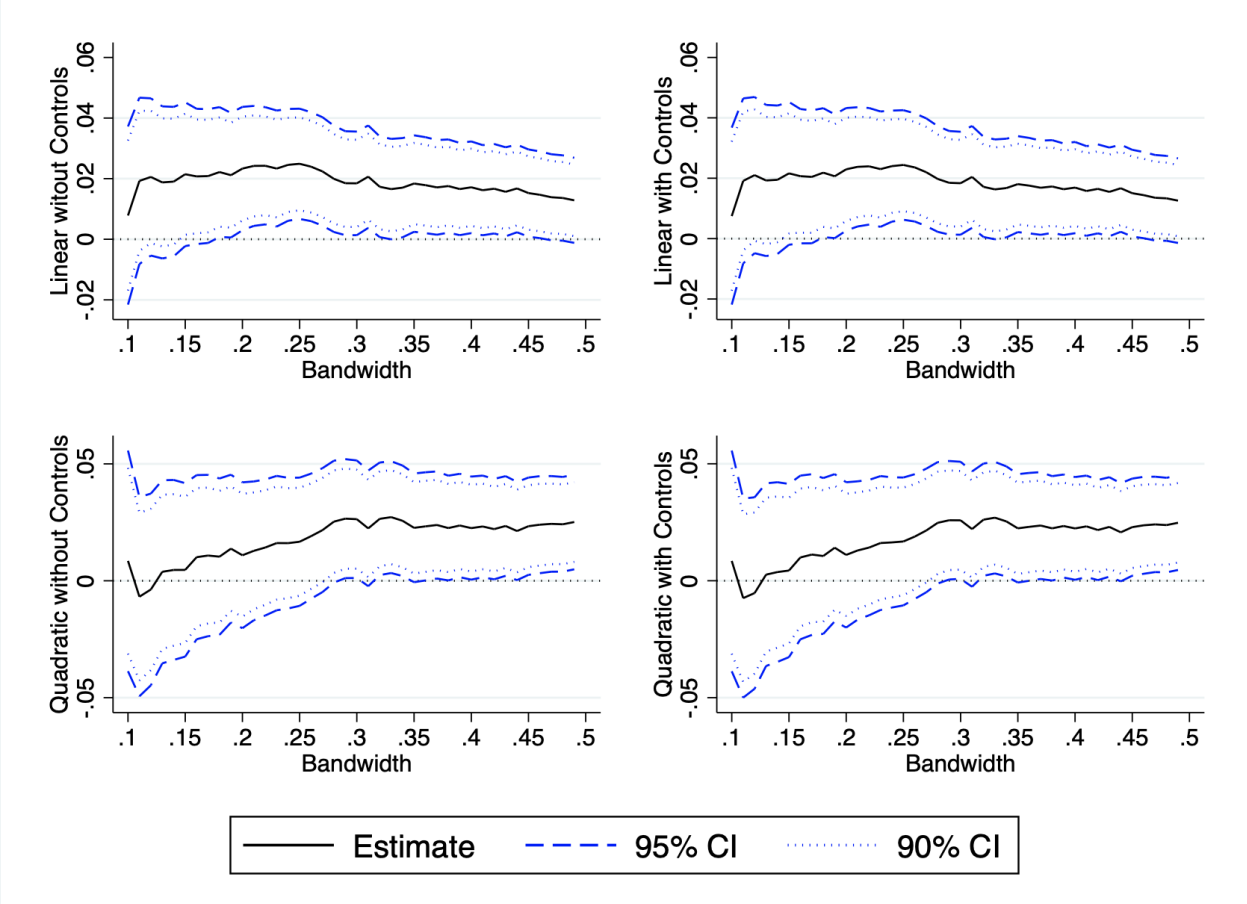

Figure I.9: Democratic Primary Votes

Note: The graph reflects the point estimate, 95 percent confidence interval, and 90 percent confidence interval of the effect of the top percent policy on the outcome of interest. Each panel represents a different specification.

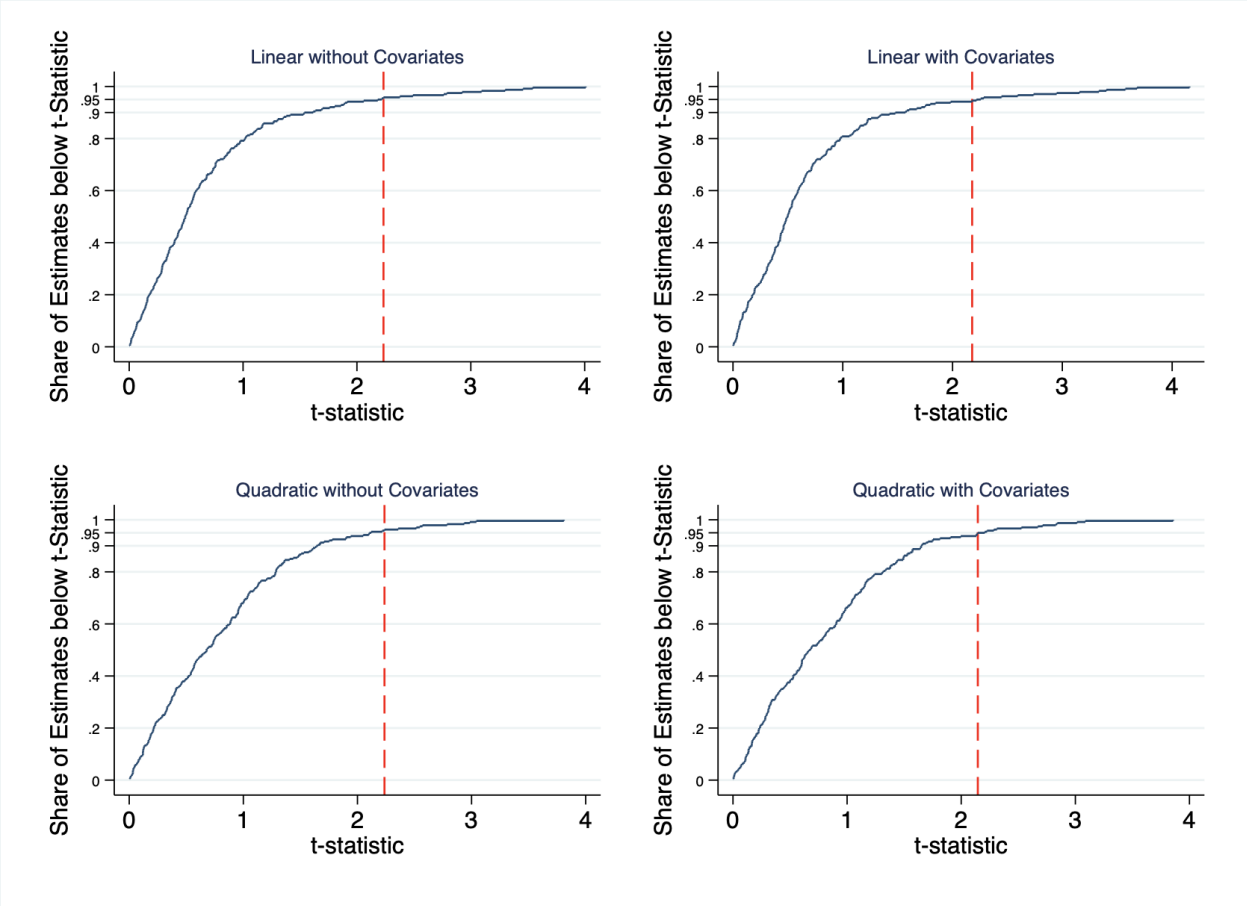

Figure I.10: Primary Votes Cast

Note: Each graph reflects the cumulative distribution of estimated t-statistics using the falsifcation tests described in Section 3.1.2. The red dashed line denotes the t-statistic estimated at the true 96th percentile policy threshold.

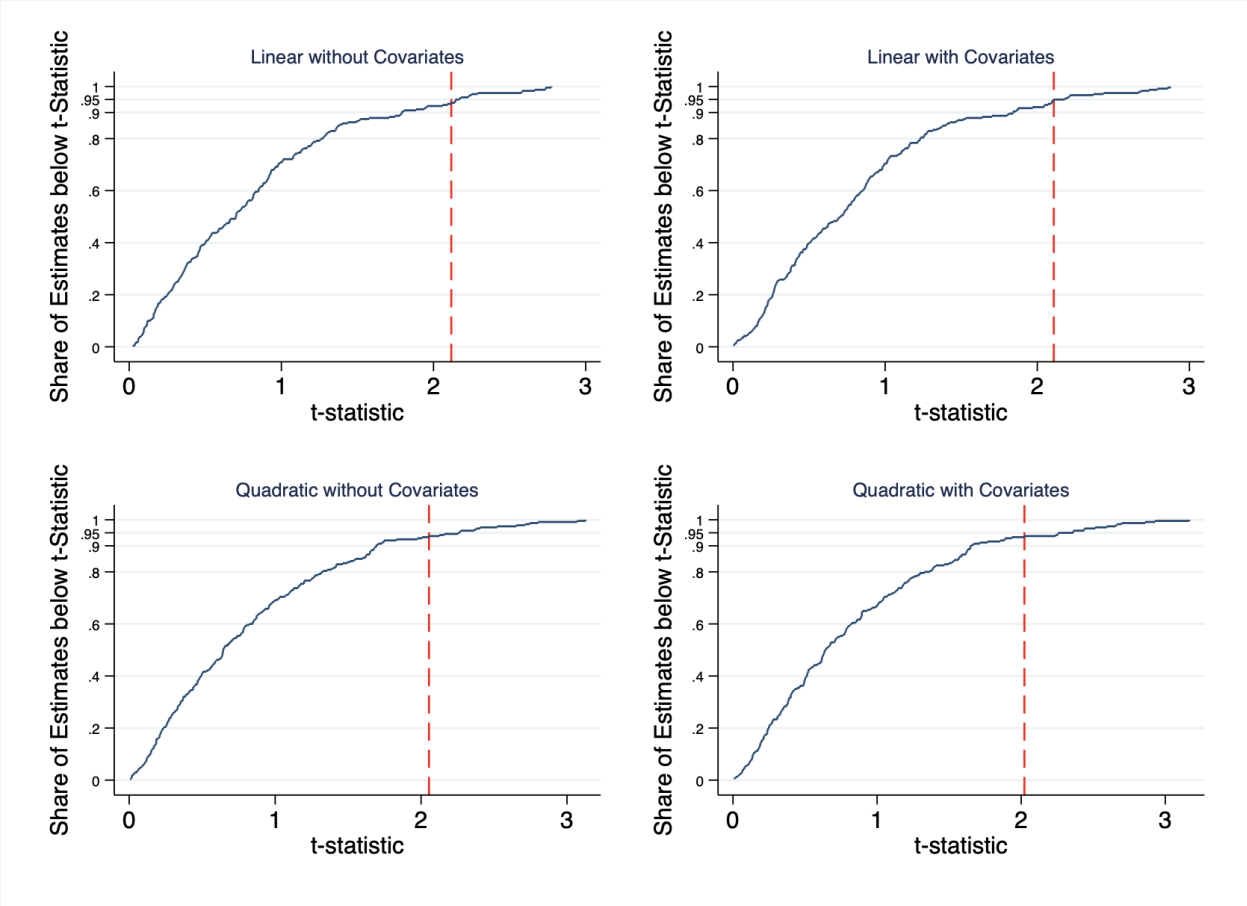

Figure I.11: Democratic Presidential Primary Votes Cast

Note: Each graph reflects the cumulative distribution of estimated t-statistics using the falsifcation tests described in Section 3.1.2. The red dashed line denotes the t-statistic estimated at the true 96th percentile policy threshold.

J CIRP Entering Freshman Survey Appendix

Table J.1: Political Ideology of Californian Students by Type of College

How would you characterize your political views?	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Far right	0.7	1.5	1.3	2.1	1.2
Conservative	14.6	24.8	17.5	18.6	19.2
Middle of the road	43.1	38.9	47.5	50.0	43.2
Liberal	38.4	31.8	30.8	25.5	33.3
Far left	3.2	3.1	2.9	3.7	3.1
N	120,552	139,172	125,714	9,993	395,431

Note: The data are from HERI's Entering Freshmen Survey among Californian institutions from 2000 to 2010. "UC" refers to UC freshmen, "Priv" refers to private college freshmen, "CSU" refers to CSU freshmen, and "2-year" refers to two-year college freshmen.

Table J.2: Economic Views of Californian Students by Type of College

View: A national health care plan is needed to cover everybody's medical costs	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	7.2	11.8	7.0	5.5	8.6
Somewhat Disagree	19.4	21.2	18.4	17.3	19.6
Somewhat Agree	43.9	39.8	42.4	42.7	42.1
Strongly Agree	29.5	27.2	32.2	34.5	29.7
N	59,400	55,756	54,829	2,707	172,692
View: Addressing global warming should be a federal priority	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	4.6	10.9	6.8	9.1	7.3
Somewhat Disagree	15.6	19.4	20.1	21.6	18.3
Somewhat Agree	41.9	36.7	42.0	44.6	40.3
Strongly Agree	37.8	33.0	31.1	24.7	34.1
N	25,510	22,801	22,122	287	70,720
View: Federal military spending should be increased	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	29.6	25.1	22.9	20.5	25.8
Somewhat Disagree	49.9	47.3	48.3	46.0	48.4
Somewhat Agree	17.5	23.5	24.2	26.8	21.8
Strongly Agree	3.0	4.1	4.6	6.7	4.0
N	87,181	88,685	79,958	5,558	261,382
View: The federal government is not doing enough to control pollution	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	2.0	3.4	2.8	2.8	2.7
Somewhat Disagree	12.3	16.6	16.0	16.9	14.9
Somewhat Agree	44.4	42.1	44.9	44.7	43.8
Strongly Agree	41.3	38.0	36.4	35.6	38.6
N	59,548	55,938	54,924	2,708	173,118

Note: The data are from HERI's Entering Freshmen Survey among Californian institutions from 2000 to 2010. "UC" refers to UC freshmen, "Priv" refers to private college freshmen, "CSU" refers to CSU freshmen, and "2-year" refers to two-year college freshmen.

Table J.3: Economic Views of Californian Students by Type of College

View: The federal government should raise taxes to reduce the deficit	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	15.4	19.6	20.8	24.5	18.5
Somewhat Disagree	49.0	48.8	51.2	51.5	49.7
Somewhat Agree	29.7	26.7	23.8	20.7	26.8
Strongly Agree	5.9	4.8	4.2	3.3	5.0
N	50,706	42,123	43,521	1,621	137,971
View: Through hard work, everybody can succeed in American society	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	4.7	5.3	3.8	3.5	4.6
Somewhat Disagree	19.8	21.1	15.5	13.9	18.8
Somewhat Agree	40.2	39.2	37.2	32.6	38.9
Strongly Agree	35.3	34.3	43.4	50.0	37.8
N	51,555	44,693	44,613	2,442	143,303
View: Wealthy people should pay a larger share of taxes than they do now	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	10.8	18.0	13.4	16.6	14.3
Somewhat Disagree	28.7	31.8	30.9	32.6	30.6
Somewhat Agree	40.9	34.9	38.5	34.2	37.9
Strongly Agree	19.5	15.3	17.1	16.6	17.2
N	110,887	132,583	120,352	8,420	372,242

Note: The data are from HERI's Entering Freshmen Survey among Californian institutions from 2000 to 2010. "UC" refers to UC freshmen, "Priv" refers to private college freshmen, "CSU" refers to CSU freshmen, and "2-year" refers to two-year college freshmen.

Table J.4: Sociocultural Views of Californian Students by Type of College

View: Abortion should be legal	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	15.5	27.6	21.3	30.1	22.0
Somewhat Disagree	15.8	14.5	17.6	19.7	16.0
Somewhat Agree	32.4	25.4	31.3	29.2	29.6
Strongly Agree	36.3	32.5	29.7	20.9	32.5
N	114,301	130,571	119,204	10,176	374,252
View: It is important to have laws prohibiting homosexual relationships	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	56.9	52.5	49.2	39.1	52.4
Somewhat Disagree	25.1	23.0	28.1	30.0	25.5
Somewhat Agree	11.6	12.4	13.7	16.8	12.7
Strongly Agree	6.4	12.1	8.9	14.0	9.4
N	103,144	121,853	110,066	8,179	343,242
View: Marijuana should be legalized	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	27.8	32.9	31.2	35.2	30.9
Somewhat Disagree	32.2	28.3	29.1	27.1	29.8
Somewhat Agree	27.7	26.4	26.7	24.1	26.8
Strongly Agree	12.3	12.3	13.0	13.6	12.5
N	113,751	130,050	118,662	10,133	372,596
View: Racial discrimination is no longer a major problem in America	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	38.0	37.0	35.8	35.5	36.9
Somewhat Disagree	45.5	45.4	43.9	40.5	44.8
Somewhat Agree	14.3	15.4	17.3	20.0	15.8
Strongly Agree	2.3	2.2	3.1	4.1	2.6
N	113,962	130,302	118,683	10,106	373,053

Note: The data are from HERI's Entering Freshmen Survey among Californian institutions from 2000 to 2010. "UC" refers to UC freshmen, "Priv" refers to private college freshmen, "CSU" refers to CSU freshmen, and "2-year" refers to two-year college freshmen.

Table J.5: Sociocultural Views of Californian Students by Type of College

View: Same-sex couples should have the right to legal marital status	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	12.5	23.3	16.7	23.2	17.9
Somewhat Disagree	16.4	16.4	18.7	20.9	17.2
Somewhat Agree	29.5	23.7	30.0	29.8	27.6
Strongly Agree	41.7	36.6	34.7	26.1	37.3
N	113,369	129,623	118,132	10,045	371,169
View: The activities of married women are best confined to the home and family	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	61.0	61.6	52.4	41.3	57.9
Somewhat Disagree	21.7	21.4	24.5	28.1	22.7
Somewhat Agree	12.8	12.5	17.1	22.5	14.4
Strongly Agree	4.6	4.4	6.0	8.1	5.1
N	59,179	89,497	75,205	6,681	230,562
View: The death penalty should be abolished	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	20.2	24.1	27.6	30.6	24.2
Somewhat Disagree	41.1	38.1	41.0	37.7	39.9
Somewhat Agree	24.8	22.3	20.6	20.9	22.5
Strongly Agree	13.8	15.4	10.7	10.8	13.3
N	102,917	121,422	109,820	8,147	342,306
View: The federal government should do more to control the sale of handguns	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	4.8	6.8	6.1	6.8	6.0
Somewhat Disagree	13.7	13.9	14.3	13.9	14.0
Somewhat Agree	42.5	38.2	39.2	33.9	39.7
Strongly Agree	39.0	41.1	40.4	45.5	40.3
N	110,705	132,563	119,936	8,438	371,642

Note: The data are from HERI's Entering Freshmen Survey among Californian institutions from 2000 to 2010. "UC" refers to UC freshmen, "Priv" refers to private college freshmen, "CSU" refers to CSU freshmen, and "2-year" refers to two-year college freshmen.

Table J.6: Sociocultural Views of Californian Students by Type of College

View: There is too much concern in the courts for the rights of criminals	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	7.9	8.5	6.8	7.7	7.8
Somewhat Disagree	37.5	35.7	30.7	25.8	34.4
Somewhat Agree	47.6	46.9	51.9	52.0	48.9
Strongly Agree	6.9	8.8	10.7	14.6	9.0
N	112,581	128,426	117,293	10,035	368,335
View: Undocumented immigrants should be denied access to public education	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Strongly Disagree	29.8	24.9	32.0	38.4	29.1
Somewhat Disagree	36.3	34.9	30.4	28.3	33.9
Somewhat Agree	22.9	24.8	22.5	19.6	23.3
Strongly Agree	11.0	15.4	15.0	13.6	13.7
N	51,242	44,268	44,369	2,428	142,307

Note: The data are from HERI's Entering Freshmen Survey among Californian institutions from 2000 to 2010. "UC" refers to UC freshmen, "Priv" refers to private college freshmen, "CSU" refers to CSU freshmen, and "2-year" refers to two-year college freshmen.

Table J.7: Descriptive Statistics on Californian Students by Type of College

Race/Ethnicity Group	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
American Indian	0.1	0.2	0.2	0.3	0.1
Asian	39.4	15.3	17.8	13.3	23.3
Black	2.5	3.3	4.2	5.0	3.4
Hispanic	14.6	10.4	24.8	45.5	17.2
White	30.1	55.5	38.3	21.8	41.4
Other	3.7	3.0	3.9	4.7	3.5
Two or more race/ethnicity	9.6	12.3	10.9	9.4	11.0
N	124,121	144,094	132,593	11,043	411,851
Citizenship status:	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Neither/None of the above	2.0	3.2	2.1	5.1	2.5
Permanent resident	7.1	2.9	5.4	8.4	5.1
U.S. citizen	91.0	93.9	92.5	86.5	92.4
N	127,474	145,738	136,435	11,304	420,951
Your religious preference	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Protestant	28.8	42.7	33.5	35.9	35.3
Roman Catholic	23.3	25.2	32.2	36.6	27.1
Jewish	3.7	3.3	1.7	0.6	2.8
Other	12.8	6.3	8.8	8.8	9.1
None	31.5	22.6	23.9	18.0	25.6
N	123,909	142,197	130,339	10,591	407,036

Note: The data are from HERI's Entering Freshmen Survey among Californian institutions from 2000 to 2010. "UC" refers to UC freshmen, "Priv" refers to private college freshmen, "CSU" refers to CSU freshmen, and "2-year" refers to two-year college freshmen.

Table J.8: Political Ideology of American Students by Type of College

How would you characterize your political views?	Institution Type			
	University	4-year	2-year	Total
	%	%	%	%
Far right	1.5	1.8	2.5	1.7
Conservative	21.2	22.3	20.3	21.8
Middle of the road	43.8	45.8	51.9	45.0
Liberal	30.4	26.9	21.3	28.4
Far left	3.0	3.2	4.0	3.1
N	1,655,052	2,060,615	42,014	3,757,681

Note: The data are from HERI's Entering Freshmen Survey among American institutions from 2000 to 2010. "University" refers to research university freshmen, "4-year" refers to teaching college freshmen, and "2-year" refers to community college freshmen.

Table J.9: Descriptive Statistics on American Students by Type of College

Race/Ethnicity Group	Institution Type			
	University	4-year	2-year	Total
	%	%	%	%
American Indian	0.3	0.3	1.6	0.3
Asian	11.4	4.7	4.2	7.6
Black	6.5	8.2	16.8	7.6
Hispanic	5.7	5.1	21.7	5.6
White	68.0	74.1	47.0	71.1
Other	2.2	1.9	3.0	2.0
Two or more race/ethnicity	5.9	5.8	5.7	5.8
N	1,722,161	2,164,260	46,650	3,933,071
Citizenship status:	Institution Type			
	University	4-year	2-year	Total
	%	%	%	%
Neither/None of the above	2.1	2.0	2.8	2.0
Permanent resident	3.2	1.8	5.7	2.5
U.S. citizen	94.6	96.2	91.5	95.5
N	1,765,970	2,199,465	47,923	4,013,358
Your religious preference	Institution Type			
	University	4-year	2-year	Total
	%	%	%	%
Protestant	40.2	46.8	52.8	44.0
Roman Catholic	27.7	28.4	24.8	28.0
Jewish	4.6	2.3	0.5	3.3
Other	6.9	5.0	7.0	5.9
None	20.7	17.5	14.9	18.9
N	1,708,947	2,133,420	45,219	3,887,586

Note: The data are from HERI's Entering Freshmen Survey among American institutions from 2000 to 2010. "University" refers to research university freshmen, "4-year" refers to teaching college freshmen, and "2-year" refers to community college freshmen.

K HERI Faculty Survey Appendix

Table K.1: Political Ideology of Californian Faculty by Type of College

How would you characterize your political views?	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Far right	0.1	0.2	0.1	0.4	0.2
Conservative	8.4	14.8	12.3	20.8	13.2
Middle of the road	34.2	35.5	33.7	42.0	35.2
Liberal	50.4	43.8	46.3	34.0	45.1
Far left	6.9	5.6	7.6	2.8	6.3
N	1,632	2,768	2,640	712	7,752

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.2: Campus Views of Californian Faculty by Type of College

View: Racist/sexist speech should be prohibited on campus	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Disagree strongly	27.2	22.4	20.7	19.9	22.6
Disagree somewhat	24.0	23.6	25.1	19.5	24.0
Agree somewhat	24.2	23.9	24.7	25.2	24.3
Agree strongly	24.6	30.1	29.5	35.4	29.1
N	687	1,261	1,268	226	3,442

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.3: Political Ideology of Californian Faculty by STEM and Type of College

Panel A. STEM Faculty					
	Institution Type				
How would you characterize your political views?	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Far right	0.3	0.2	0.1	0.0	0.2
Conservative	10.1	14.7	13.7	23.8	13.4
Middle of the road	42.2	38.7	38.7	45.4	40.2
Liberal	45.0	42.7	42.9	29.2	42.7
Far left	2.5	3.7	4.6	1.5	3.5
N	733	653	786	130	2,302
Panel B. Non-STEM Faculty					
	Institution Type				
How would you characterize your political views?	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Far right	0.0	0.2	0.1	0.5	0.2
Conservative	7.0	14.9	11.7	20.1	13.1
Middle of the road	27.7	34.5	31.6	41.2	33.1
Liberal	54.8	44.2	47.7	35.1	46.2
Far left	10.5	6.2	8.8	3.1	7.5
N	899	2,115	1,854	582	5,450

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.4: Career Objectives of Californian Faculty by Type of College

Objective: Becoming an authority in my field	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	1.5	9.8	9.8	16.4	8.7
Somewhat important	11.5	23.2	25.1	26.5	21.6
Very important	33.2	34.0	34.3	32.0	33.7
Essential	53.8	33.0	30.9	25.1	36.0
N	1,680	2,818	2,685	737	7,920
Objective: Influencing the political structure	**Institution Type**				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	44.6	39.4	36.5	39.1	39.5
Somewhat important	37.6	40.4	38.5	37.1	38.8
Very important	13.6	15.9	18.1	17.3	16.3
Essential	4.2	4.2	6.9	6.5	5.4
N	1,672	2,809	2,673	734	7,888
Objective: Influencing social values	**Institution Type**				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	27.4	17.8	17.2	13.6	19.2
Somewhat important	41.1	33.9	38.7	33.5	37.0
Very important	24.0	35.3	31.8	36.5	31.8
Essential	7.5	13.0	12.4	16.5	12.0
N	1,670	2,807	2,676	735	7,888
Objective: Helping to promote racial understanding	**Institution Type**				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	8.4	6.1	5.8	5.3	6.4
Somewhat important	40.0	31.0	29.1	25.8	31.7
Very important	33.5	36.9	37.8	37.4	36.5
Essential	18.1	26.0	27.4	31.5	25.3
N	1,664	2,804	2,667	737	7,872
Objective: Obtaining recognition from my colleagues for contribution to my field	**Institution Type**				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	3.0	10.9	9.0	21.6	9.6
Somewhat important	21.9	34.3	34.3	43.7	32.6
Very important	42.1	35.9	37.6	23.6	36.6
Essential	33.0	18.9	19.1	11.1	21.2
N	1,669	2,803	2,675	737	7,884

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.5: Instructional Goals of Californian Faculty by Type of College

UG Goal: Develop moral character	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	14.4	8.0	10.6	5.0	9.9
Somewhat important	41.0	28.7	36.5	26.9	33.8
Very important	29.6	34.5	33.5	37.0	33.4
Essential	14.9	28.9	19.4	31.2	22.9
N	1,569	2,671	2,627	722	7,589
UG Goal: Help students develop personal values	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	12.9	6.4	8.2	4.6	8.2
Somewhat important	40.1	26.9	35.5	23.8	32.3
Very important	33.9	39.5	39.0	44.0	38.6
Essential	13.1	27.2	17.4	27.6	20.9
N	1,565	2,665	2,626	720	7,576
UG Goal: Enhance students' knowledge of and appreciation for other races	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	14.6	10.9	4.6	8.7	9.3
Somewhat important	35.2	24.9	25.8	25.5	27.0
Very important	32.1	33.8	33.2	28.1	32.6
Essential	18.1	30.3	36.4	37.7	31.1
N	321	758	624	231	1,934
UG Goal: Prepare students for responsible citizenship	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	10.5	9.4	6.9	4.8	8.3
Somewhat important	34.8	28.8	27.2	26.7	29.3
Very important	38.2	40.4	38.3	37.4	39.0
Essential	16.5	21.4	27.5	31.2	23.4
N	978	1,739	1,648	439	4,804
UG Goal: Develop ability to think clearly	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not important	0.0	0.2	0.0	0.0	0.1
Somewhat important	0.5	0.4	0.7	1.0	0.6
Very important	8.5	11.2	9.4	12.6	10.2
Essential	91.0	88.2	89.9	86.4	89.2
N	1,582	2,683	2,651	723	7,639

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.6: Institutional Goals of Californian Faculty by Type of College

Inst Priority: To promote the intellectual development of students	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Low priority	1.6	1.3	3.9	2.4	2.3
Medium priority	10.7	9.8	18.2	13.1	13.2
High priority	33.2	34.7	32.6	37.3	33.9
Highest priority	54.5	54.2	45.3	47.2	50.6
N	1,650	2,784	2,656	718	7,808
Inst Priority: To help students examine and understand their personal values	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Low priority	21.0	7.3	17.2	7.9	13.6
Medium priority	45.2	25.8	42.7	34.5	36.4
High priority	27.6	40.1	29.6	37.2	33.6
Highest priority	6.3	26.8	10.5	20.3	16.3
N	1,637	2,775	2,644	718	7,774
Inst Priority: To help students learn how to bring about change in society	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Low priority	38.9	23.4	32.1	22.5	29.5
Medium priority	41.8	40.6	40.6	42.7	41.0
High priority	15.8	26.7	20.1	24.7	22.0
Highest priority	3.4	9.3	7.3	10.1	7.4
N	1,616	2,755	2,630	712	7,713
Inst Priority: To maintain a climate where different opinions can be aired	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Low priority	5.9	9.6	8.5	8.1	8.1
Medium priority	25.5	27.1	25.4	27.8	26.2
High priority	41.7	40.5	40.4	42.6	41.0
Highest priority	26.9	22.8	25.8	21.5	24.6
N	573	698	733	270	2,274
Inst Priority: To develop among students and faculty multicultural appreciation	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Low priority	9.8	8.8	6.8	6.7	8.1
Medium priority	38.5	33.6	29.3	31.1	33.1
High priority	36.5	39.5	38.9	39.6	38.6
Highest priority	15.2	18.2	25.0	22.6	20.2
N	572	697	737	270	2,276

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.7: Teaching Methods of Californian Faculty by Type of College

Instructional Method: Class discussions	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	8.5	6.7	4.9	6.9	6.5
Some	34.2	21.7	25.8	22.4	25.7
Most	22.8	19.5	21.3	17.9	20.7
All	34.4	52.0	48.0	52.8	47.1
N	1,310	2,244	2,352	665	6,571
Instructional Method: Cooperative learning (small groups)	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	48.4	27.9	26.9	22.6	31.1
Some	35.3	33.8	39.4	34.8	36.2
Most	9.1	15.9	15.6	16.5	14.5
All	7.1	22.5	18.1	26.1	18.2
N	1,305	2,232	2,348	660	6,545
Instructional Method: Experiential learning/Field studies	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	60.9	46.5	48.7	52.4	50.7
Some	25.7	27.9	29.4	24.4	27.7
Most	7.0	11.7	11.7	9.6	10.6
All	6.4	13.9	10.2	13.6	11.0
N	1,294	2,223	2,341	655	6,513
Instructional Method: Teaching assistants	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	25.6	56.4	70.3	79.4	58.5
Some	39.7	22.3	21.6	13.1	24.2
Most	19.0	10.2	4.5	3.3	8.9
All	15.8	11.1	3.6	4.2	8.5
N	774	1,639	1,690	427	4,530

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.8: Teaching Methods of Californian Faculty by Type of College

Instructional Method: Group projects	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	56.2	35.8	37.1	48.0	41.5
Some	34.8	37.9	41.2	31.0	37.8
Most	5.5	13.5	12.2	10.7	11.2
All	3.5	12.8	9.4	10.3	9.5
N	1,297	2,229	2,344	662	6,532
Instructional Method: Extensive lecturing	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	8.1	20.6	15.3	23.8	16.5
Some	21.1	31.6	29.5	30.2	28.6
Most	38.0	27.8	32.5	26.4	31.4
All	32.8	20.1	22.7	19.6	23.5
N	1,304	2,228	2,347	663	6,542
Instructional Method: Readings on racial and ethnic issues	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	73.7	56.2	58.6	62.7	61.2
Some	15.5	25.4	23.5	21.4	22.3
Most	4.8	9.2	9.0	7.4	8.1
All	6.0	9.3	8.9	8.5	8.4
N	1,294	2,227	2,340	660	6,521
Instructional Method: Readings on women and gender issues	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	72.7	56.8	59.0	62.7	61.4
Some	16.6	25.4	24.7	22.4	23.1
Most	5.7	9.2	8.6	8.2	8.2
All	5.0	8.5	7.8	6.7	7.4
N	1,299	2,230	2,344	660	6,533

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.9: Research and Teaching Activity among Californian Faculty by Type of College

Do your interests lie primarily in teaching or research?	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Very heavily in teaching	2.5	26.4	24.4	65.7	24.2
In both, but leaning toward teaching	16.0	33.3	38.2	25.4	30.6
In both, but leaning toward research	62.3	32.9	31.1	7.3	36.2
Very heavily in research	19.2	7.4	6.3	1.6	9.0
N	1,690	2,813	2,697	728	7,928
What is your principal activity in your current position at this institution?	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Administration	10.7	19.2	16.1	20.5	16.5
Teaching	35.5	62.3	70.4	68.1	60.1
Research	47.7	15.3	10.8	0.4	19.1
Services to clients and patients	4.8	1.8	1.2	7.0	2.7
Other	1.3	1.3	1.5	4.0	1.6
N	1,633	2,817	2,683	753	7,886
Publish: Articles in academic or professional journals	Institution Type UC	Priv	CSU	2-year	Total
	%	%	%	%	%
None	2.4	18.7	12.7	66.1	17.6
1-2	2.8	15.1	14.7	19.6	12.8
3-4	4.0	13.4	14.8	8.2	11.4
5-10	10.4	15.9	18.1	3.7	14.4
11-20	16.4	11.6	15.1	1.2	12.8
21-50	27.9	11.5	13.5	0.8	14.7
51+	36.2	13.8	11.0	0.3	16.4
N	1,666	2,760	2,663	723	7,812

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.10: Descriptive Statistics on Californian Faculty by Type of College

What is your present academic rank?	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Professor	59.7	41.1	55.4	18.5	47.8
Associate Professor	18.3	21.5	15.1	4.6	17.1
Assistant Professor	18.5	17.1	11.9	1.8	14.2
Lecturer	2.3	8.2	13.5	3.4	8.3
Instructor	0.1	6.1	1.5	54.6	7.8
Other	1.1	5.9	2.6	17.3	4.8
N	1,703	2,846	2,720	742	8,011
Race/Ethnicity Group	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
American Indian	0.2	0.2	0.4	0.5	0.3
Asian	8.1	3.8	7.7	4.3	6.0
Black	1.4	1.5	2.6	3.3	2.0
Hispanic	2.7	1.7	5.2	7.2	3.6
White	83.8	88.6	78.7	77.0	83.1
Other	1.6	1.2	2.6	1.5	1.8
Two or more race/ethnicity	2.3	3.0	2.9	6.1	3.1
N	1,675	2,838	2,690	749	7,952
STEM	Institution Type				
	UC	Priv	CSU	2-year	Total
	%	%	%	%	%
Not STEM	55.1	76.4	70.5	82.0	70.4
STEM	44.9	23.6	29.5	18.0	29.6
N	1,717	2,895	2,740	760	8,112

Note: The data are from HERI's Faculty Survey among Californian institutions from 1989 to 1998. "UC" refers to UC faculty, "Priv" refers to private college faculty, "CSU" refers to CSU faculty, and "2-year" refers to community college faculty.

Table K.11: Political Ideology of American Faculty by Type of College

How would you characterize your political views?	Institution Type			
	University	4-year	2-year	Total
	%	%	%	%
Far right	0.3	0.4	0.4	0.3
Conservative	13.5	19.2	24.3	16.8
Middle of the road	36.8	39.2	44.8	38.4
Liberal	44.5	37.4	28.2	40.2
Far left	5.0	3.8	2.3	4.3
N	39,220	33,702	6,533	79,455

Note: The data are from HERI's Faculty Survey among American institutions from 1989 to 1998. "University" refers to research university faculty, "4-year" refers to teaching college faculty, and "2-year" refers to community college faculty.